Predicting and Estimation from Time Series

Class 25. 16 Nov 2010

An automotive example

- $\mathcal{L}_{\mathcal{A}}$ Determine automatically, by only *listening* to a running automobile, if it is:
	- \Box Idling; or
	- \Box Travelling at constant velocity; or
	- \Box Accelerating; or
	- \Box **Decelerating**
- \mathbb{R}^3 Assume (for illustration) that we only record energy level (SPL) in the sound
	- \Box The SPL is measured once per second

What we know

- **An automobile that is at rest can accelerate,** or continue to stay at rest
- **An accelerating automobile can hit a steady**state velocity, continue to accelerate, or decelerate
- \blacksquare A decelerating automobile can continue to decelerate, come to rest, cruise, or accelerate
- A automobile at a steady-state velocity can stay in steady state, accelerate or decelerate

- The probability distribution of the SPL of the sound is different in the various conditions
	- $\textcolor{orange}\blacksquare$ As shown in figure
		- π In reality, depends on the car
- The distributions for the different conditions overlap
	- □ Simply knowing the current sound level is not enough to know the state of the car

- $\mathcal{C}^{\mathcal{A}}$ The state-space model
	- \Box Assuming all transitions from a state are equally probable

- \blacksquare A T=0, before the first observation, we know nothin g of the state
	- □ Assume all states are equally likely

- **At T=0 we observe the sound level** $x_0 = 67$ dB SPL \Box The observation modifies our belief in the state of the system
- $P(x_0 | \text{idle}) = 0$
- $\text{P}(\text{x}_0 | \text{deceleration}) = 0.0001$
- $\text{P}(\text{x}_0|\text{acceleration}) = 0.7$
- $\text{P}(\text{x}_0|\text{cruising}) = 0.5$
	- □ Note, these don't have to sum to 1
	- \Box In fact, since these are densities, any of them can be > 1

Estimating state after at observing $\rm x_{0}$

- P(state $| x_0$) = C P(state)P(x_0) state)
	- \Box P(idle | x_0) = 0
	- \Box P(deceleration | x₀) = C 0.000025
	- \Box P(cruising | x_0) = C 0.125
	- \Box P(acceleration | x_0) = C 0.175

R Normalizing

- \Box P(idle | x_0) = 0
- \Box P(deceleration | x_0) = 0.000083
- \Box P(cruising | x_0) = 0.42
- \Box P(acceleration | x_0) = 0.57

- \blacksquare At T=0, after the first observation, we must u pdate our belief about the states
	- □ The first observation provided some evidence about the state of the system
	- □ It modifies our belief in the state of the system

Predicting the probability of idling at T=1

- \Box P(idling | idling) = 0.5;
- \Box P(idling | deceleration) = 0.25
- \Box P(idling at T=1| x_0) = $P(I_{T=0} | x_0) P(I | I) + P(D_{T=0} | x_0) P(I | D) = 2.1 \times 10^{-5}$
- \mathbb{R}^3 In general, for any state S

$$
\mathbf{P}(S_{T=1} \mid x_0) = \Sigma_{S_{T=0}} \mathbf{P}(S_{T=0} \mid x_0) \mathbf{P}(S_{T=1} \mid S_{T=0})
$$

- **At T=1 we observe** x_1 = 63dB SPL
- \blacksquare P(x₁ | idle) = 0
- \blacksquare P(x₁ deceleration) = 0.2
- \blacksquare P(x₁ | acceleration) = 0.001
- \blacksquare P(x₁ | cruising) = 0.5

Update after observing x_1

- P(state $|x_{0:1}\rangle$ = C P(state $|x_0\rangle P(x_1|$ state)
	- $P(idle | x_{0:1}) = 0$
	- ❏ \Box P(deceleration | $x_{0,1}$) = C 0.066
	- **p** P(cruising $| x_{0:1} \rangle = C 0.165$
	- \Box P(acceleration | $x_{0:1}$) = C 0.00033

R Normalizing

- $P(idle | x_{0:1}) = 0$
- \Box P(deceleration | $x_{0:1}$) = 0.285
- \Box P(cruising | $x_{0:1}$) = 0.713
- \Box P(acceleration | $x_{0:1}$) = 0. 0014

- The updated probability at T=1 incorporates information from both x_{0} and x_{1}
	- \Box $\textsf{u}\textsf{u}$ It is NOT a local decision based on $\textsf{x}_{\textsf{1}}$ alone
	- \Box Because of the Markov nature of the process, the state at T=0 affects the state at T=1 $\,$
		- F. x_{0} provides evidence for the state at T=1

Estimating a Unique state

- What we have estimated is a *distribution* over the states
- **If we had to guess a state, we would pick the** most likely state from the distributions

$$
\blacksquare
$$
 State(T=0) = Accelerating

$$
\blacksquare
$$
 State(T=1) = Cruising

- T \blacksquare At T-0 the predicted state distribution is the initial state probability
- T At each time T, the current estimate of the distribution over states considers *all* observations $\mathrm{x}_0 \, ... \, \mathrm{x}_\mathrm{T}$
	- □ A natural outcome of the Markov nature of the model
- $\mathcal{L}_{\mathcal{A}}$ The prediction +update is identical to the forward computation for HMMs to within a normalizing constant

■ Forward Algorithm:

Normalized:

 $P(S_T | x_{0:T}) = [\Sigma_{S_T} P(x_{0:T}, S_T')]^{-1} P(x_{0:T}, S_T) = C P(x_{0:T}, S_T)$

Decomposing the forward algorithm

- $P(X_{0:T}, S_T) = P(X_T | S_T) \Sigma_{S_{T-1}}$ $P(\mathrm{x}_{0:T-1},\:\mathrm{S}_{\mathrm{T}-1})\:\mathrm{P}(\mathrm{S}_{\mathrm{T}}|\:\mathrm{S}_{\mathrm{T}-1})$
- Predict:
- $P(X_{0:T-1}, S_T) = \sum_{S_{T-1}} P(X_{0:T-1}, S_{T-1}) P(S_T | S_{T-1})$ $P(\mathrm{x}_{0:T-1},\:\mathrm{S}_{\mathrm{T}-1})\:\mathrm{P}(\mathrm{S}_{\mathrm{T}}\vert\:\mathrm{S}_{\mathrm{T}-1})$
- **Update:**
- $P(X_{0:T}, S_T) = P(X_T | S_T) P(X_{0:T-1}, S_T)$

The state is estimated from the updated distribution

 \Box The updated distribution is propagated into time, not the state

- $\mathcal{L}_{\mathcal{A}}$ The probability distribution for the observations at the next time is a mixture:
	- ❏ $P(x_T|x_{0:T-1}) = \sum_{S_T} P(x_T|S_T) P(S_T|x_{0:T-1})$
- \blacksquare The actual observation can be predicted from $P(x_T | x_{0:T-1})$ 11-755/18797

Predicting the next observation

■ MAP estimate:

 \Box argmax $_{\mathrm{x_{T}}}$ P(x_T|x_{0:T-1}) $P(X_T | X_{0:T-1}$

NMSE estimate:

 \Box Expectation(x_T | x_{0:T-1})

Difference from Viterbi decodin g

- Estimating only the *current* state at any time
	- □ Not the state sequence
	- □ Although we are considering all past observations
- The most likely state at T and T+1 may be such that there is no valid transition between \mathbf{S}_T and $\mathbf{S}_{\mathsf{T+1}}$

A *known* state model

- HMM assumes a very coarsely quantized state space
	- □ Idling / accelerating / cruising / decelerating
- **Actual state can be finer**
	- □ Idling, accelerating at various rates, decelerating at various rates, cruising at various speeds
- Solution: Many more states (one for each acceleration /deceleration rate, crusing speed)?

■ Solution: A *continuous* valued state

The real-valued state model

A state equation describing the dynamics of the system

$$
s_t = f(s_{t-1}, \varepsilon_t)
$$

- \textsf{s}_t is the state of the system at time t
- \texttt{b} ϵ_{t} is a driving function, which is assumed to be random
- **The state of the system at any time depends only on the** state at the previous time instant and the driving term at the current time
- **An observation equation relating state to observation**

$$
o_t = g(s_t, \gamma_t)
$$

- *^o*^t is the observation at time t
- \Box $\gamma_{\rm t}$ is the noise affecting the observation (also random)
- $\mathcal{L}^{\text{max}}_{\text{max}}$ ■ The observation at any time depends only on the current state of the system and the noise

Continuous state system

$$
S_t = f(S_{t-1}, \mathcal{E}_t)
$$

$$
O_t = g(S_t, \gamma_t)
$$

- The state is a continuous valued parameter that is not directly seen
	- \Box The state is the position of navlab or the star
- The observations are dependent on the state and are the only way of knowing about the state
	- □ Sensor readings (for navlab) or recorded image (for the telescope)

Statistical Prediction and Estimation

- Given an *a priori* probability distribution for the state
	- □ P₀(s): Our belief in the state of the system before we observe any data
		- $\mathcal{C}^{\mathcal{A}}$ Probability of state of navlab
		- $\mathcal{L}_{\mathcal{A}}$ Probability of state of stars
- Given a sequence of observations O_0 .. O_t
- \mathcal{L}^{max} ■ Estimate state at time *t*

Prediction and update at $t = 0$

Prediction

□ Initial probability distribution for state

- $P(S_0) = P_0(S_0)$
- **Update:**
	- □ Then we observe *o*₀
	- □ We must update our belief in the state

 $(S_0)P(o_0 | s)$ $P_0(S_0)P(o_0 | s_0)$ $(s_0 | o_0) = \frac{1}{s_0} \frac{(b_0)^T (b_0 + b)}{(b_0 + b_0)} = \frac{1}{s_0} \frac{(b_0)^T (b_0 + b_0)}{(b_0 + b_0)}$ $P(s_0 | o_0) = \frac{P(s_0)P(o_0 | s)}{P(o_0)} = \frac{P_0(s_0)P(o_0 | s)}{P(o_0)}$

 $P(s_0|o_0) = C.P_0(s_0)P(o_0|s_0)$

The observation probability: $P(o|s)$

- $o_t = g(s_t, \gamma_t)$
	- This is a (possibly many-to-one) stochastic function of state $\bm{{\mathsf{s}}}_\text{t}$ and noise $\bm{{\mathsf{\gamma}}}_\text{t}$
	- \Box Noise γ_t is random. Assume it is the same dimensionality as o_t
- \blacksquare Let $\mathsf{P}_\gamma(\gamma_\mathsf{t})$ be the probability distribution of γ_t
- Let $\{ \gamma : g(\mathcal{S}_t, \gamma) = o_t \}$ be the set of γ that result in o_t

$$
P(o_t | s_t) = \sum_{\gamma:g(s_t, \gamma) = o_t} \frac{P_{\gamma}(\gamma)}{|J_{g(s_t, \gamma)}(o_t)|}
$$

The observation probability

- $P(o|s) = ?$ $o_t = g(s_t, \gamma_t)$ \mathcal{L} $\qquad \qquad =\quad \sum$ $\frac{f^{1-\nu}f^{j}}{g^{j}g^{j}(s-y)-g}$ $\int_{g^{j}(s-y)}(0,0)dy$ $P(o_1 | s_2) = \sum_{i=1}^{n} a_i$ $\sum_{g(s_t, \gamma) = o_t} | J_{g(s_t, \gamma)}(o_t) |$ $(Q_t | S_t) = \sum_{\mathcal{S} \in \mathcal{S}} \frac{P_{\gamma}(\gamma)}{1 - \gamma}$ $\frac{1}{\gamma}$ (γ
- **The J is a jacobian**

 γ : $g(s_t, \gamma)$

t

 $g(s_t, \gamma) = o_t$ \mathbf{v} $g(s_t, \gamma)$

 $g(s_t, \gamma) = o_t$ **i g** (s_t, γ) **i f** *f*

For scalar functions of scalar variables, it is simply a derivative: $J_{g(s_t,y)}(o_t) = \left|\frac{\partial}{\partial s_t}\right|$ ∂^{γ} =| $g(s_t, \gamma)$ (o_t) $\models \left| \frac{\partial o_t}{\partial o_t} \right|$ $|J_{g(s_t,y)}(o_t)| = \left|\frac{\partial o}{\partial s}\right|$

Predicting the next state

Given $P(s_0|o_0)$, what is the probability of the state at t=1

$$
P(s_1 | o_0) = \int_{\{s_0\}} P(s_1, s_0 | o_0) ds_0 = \int_{\{s_0\}} P(s_1 | s_0) P(s_0 | o_0) ds_0
$$

 $\begin{bmatrix} 1 & 1 \\ 1 & 1 \end{bmatrix}$ State progression function:

$$
S_t = f(S_{t-1}, \mathcal{E}_t)
$$

 \Box $\varepsilon_{\rm t}$ is a driving term with probability distribution $\mathsf{P}_{\varepsilon}(\varepsilon_{\rm t})$

P(s_{t-1}) can be computed similarly to P(o|s) □ P(s₁|s₀) is an instance of this

And moving on

- \blacksquare P(s₁|o₀) is the predicted state distribution for $t = 1$
- \blacksquare Then we observe o $_1$
	- □ We must update the probability distribution for s1 $P(S_1|O_{0:1}) = CP(S_1|O_0)P(O_1|S_1)$
- We can continue on

Discrete vs. Continuous state systems

Prediction at time 0: $P(s_0) = \pi(s_0)$

 $P({\textnormal{s}}_{\textnormal{0}} \mid {\textnormal{O}}_{\textnormal{0}}) \ = C \ \pi \ ({\textnormal{s}}_{\textnormal{o}}) P({\textnormal{O}}_{\textnormal{0}} \mid {\textnormal{s}}_{\textnormal{o}})$ Update after O_0 :

Prediction at time 1:

 $P(s_1 | \mathbf{O}_0) = \sum P(s_0 | \mathbf{O}_0) P(s_1 | s_0)$ | $P(s_1$ s_0

Update after O_1 :

 $P(s_1 | O_0, O_1) = C P(s_1 | O_0) P(O_1 | s_1)$

$$
\sum_{s} \sum_{t=0}^{s} s_{t} = f(s_{t-1}, \mathcal{E}_{t})
$$

 $P(s_0) = P(s)$

$$
P(s_0| O_0) = C P(s_0) P(O_0| s_0)
$$

$$
P(s_1 | \mathbf{O}_0) = \int_{-\infty}^{\infty} P(s_0 | \mathbf{O}_0) P(s_1 | s_0) ds_0
$$

 $P(s_1 | O_0) P(O_1 | s_1)$ $P(s_1 | O_0, O_1) = C P(s_1 | O_0) P(O_1 | s_1)$

Discrete vs. Continuous State S ystems

$$
S_{t} = f(S_{t-1}, \mathcal{E}_{t})
$$
\n
$$
P(s_{t} | O_{0:t-1}) = \sum_{s_{t-1}}^{S} P(s_{t-1} | O_{0:t-1}) P(s_{t} | s_{t-1})
$$
\n
$$
P(s_{t} | O_{0:t-1}) = \int_{-\infty}^{\infty} P(s_{t-1} | O_{0:t-1}) P(s_{t} | s_{t-1})
$$
\n
$$
P(s_{t} | O_{0:t}) = CP(s_{t} | O_{0:t-1}) P(O_{t} | s_{t})
$$
\n
$$
P(s_{t} | O_{0:t}) = CP(s_{t} | O_{0:t-1}) P(O_{t} | s_{t})
$$

Discrete vs. Continuous State S ystems

1 1 \longleftarrow \rightarrow 2 03Initial state prob. Parameters π

$$
\text{transition prob } \{T_{ij}\} = P(s_t = j \mid s_{t-1} = i) \qquad \qquad P(S_t \mid S_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid S_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-1} = j) \qquad \qquad P(S_t \mid s_{t-1} = j) = P(s_t \mid s_{t-
$$

P (O | *^s*) Observation prob

$$
S_t = f(S_{t-1}, \mathcal{E}_t)
$$

$$
O_t = g(S_t, \gamma_t)
$$

$$
P(s)
$$

$$
P(s_t | s_{t-1})
$$

$$
P(o | s)
$$

Special case: Linear Gaussian model

$$
S_t = A_t S_{t-1} + \mathcal{E}_t
$$

\n
$$
P(\varepsilon) = \frac{1}{\sqrt{(2\pi)^d | \Theta_{\varepsilon}|}} \exp(-0.5(\varepsilon - \mu_{\varepsilon})^T \Theta_{\varepsilon}^{-1} (\varepsilon - \mu_{\varepsilon}))
$$

\n
$$
O_t = B_t S_t + \gamma_t
$$

\n
$$
P(\gamma) = \frac{1}{\sqrt{(2\pi)^d | \Theta_{\gamma}|}} \exp(-0.5(\gamma - \mu_{\gamma})^T \Theta_{\gamma}^{-1} (\gamma - \mu_{\gamma}))
$$

■ A *linear* state dynamics equation

- □ Probability of state driving term ε is Gaussian
- **a** Sometimes viewed as a driving term μ_{ε} and additive zero-mean noise
- A *linear* observation equation
	- $\textcolor{orange}\blacksquare$ Probability of observation noise γ is Gaussian
- A_t, B_t and Gaussian parameters assumed known \Box May vary with time

The initial state probability

$$
P_0(s) = \frac{1}{\sqrt{(2\pi)^d |R|}} \exp\left(-0.5(s-\overline{s})R^{-1}(s-\overline{s})^T\right)
$$

 $P_0(s)$ = Gaussian(s; \overline{s}, R)

- We also assume the *initial* state distribution to be Gaussian
	- □ Often assumed zero mean

The observation probability

$$
o_t = B_t s_t + \gamma_t \qquad P(\gamma) = Gaussian(\gamma; \mu_\gamma, \Theta_\gamma)
$$

$$
P(o_t | s_t) = Gaussian(o_t; \mu_{\gamma} + B_t s_t, \Theta_{\gamma})
$$

- \mathbb{R}^3 The probability of the observation, given the state, is simply the probability of the noise, with the mean shifted
	- □ Since the only uncertainty is from the noise
- $\mathcal{O}(\mathcal{O}_\mathcal{C})$ The new mean is the mean of the distribution of the noise + the value of the observation in the absence of noise

The updated state probability at $T=0$ $P(S_0 | 0_0) = C R(S_0) R(0_0 | s_0)$ $P(s_0) = Gaussian(s_0; \overline{s}, R)$ $P(o_0 | s_0) = Gaussian(o_0; \mu_\gamma + B_0 s_0, \Theta_\gamma)$

 $P(s_0 | o_0) = CGaussian(s_0; \overline{s}, R)Gaussian(o_0; \mu_\gamma + B_0 s_0, \Theta_\gamma)$

Note 1: product of two Gaussians

The product of two Gaussians is a Gaussian $Gaussian(s; \overline{s}, R)Gaussian(o; \mu + Bs, \Theta)$ $\exp ($ $(0.5(s-\bar{s})^T R^{-1}(s-\bar{s}))C_2 \exp(-\bar{s})$ $\exp(-0.5(o - \mu - Bs)^{T} \Theta^{-1}(o - \mu - Bs))$ $C_1 \exp(-0.5(s-\bar{s})^T R^{-1}(s-\bar{s}))C_2 \exp(-0.5(o-\mu - Bs)^T \Theta^{-1}(o-\mu - Bs)^T$ $\mathcal{L}^{\text{I}}(s-\bar{s})\mathcal{L}_{2} \exp(-0.5(o-\mu-Bs)^{2}\Theta^{-1}(o-\mu))$

$$
C.Gaussian(s; (R^{-1} + B^{T} \Theta^{-1} B)^{-1} (R^{-1} \overline{s} + B^{T} \Theta^{-1} (o - \mu)) (R^{-1} + B^{T} \Theta^{-1} B)^{-1})
$$

The updated state probability at $T=0$ $P(S_0 | 0_0) = C R(S_0) R(0_0 | s_0)$ $P(s_0) = Gaussian(s_0; \overline{s}, R)$

 $P(o_0 | s_0) = Gaussian(o_0; \mu_\gamma + B_0 s_0, \Theta_\gamma)$

$$
P(s_0 | o_0) =
$$

\n
$$
Gaussian(s_0; (R^{-1} + B_0^T \Theta_{\gamma}^{-1} B_0)^{-1} (R^{-1}\overline{s} + B_0^T \Theta^{-1} (o_0 - \mu_{\gamma})) (R^{-1} + B_0^T \Theta_{\gamma}^{-1} B_0)^{-1})
$$

\n
$$
P(s_0 | o_0) = Gaussian(s_0; \hat{s}_0, \hat{R}_0)
$$

The state transition probability $S_t = A_t S_{t-1} + \varepsilon_t$ $P(\varepsilon) = Gaussian(\varepsilon; \mu_{\varepsilon}, \Theta_{\varepsilon})$ $P(s_t | s_{t-1}) = Gaussian(s_t; \mu_{\varepsilon} + A_t s_{t-1}, \Theta_{\varepsilon})$

■ The probability of the state at time *t*, given the state at time *t*-1 is simply the probability of the driving term, with the mean shifted

Note 2: integral of product of two Gaussians

 $\mathcal{L}_{\mathcal{A}}$ **The integral of the product of two Gaussians** is a Gaussian

$$
\int_{-\infty}^{\infty} Gaussian(x;\mu_x,\Theta_x)Gaussian(y;Ax+b,\Theta_y)dx =
$$

$$
\int_{-\infty}^{\infty} C_1 \exp\left(-0.5(x-\mu_x)^T\Theta_x^{-1}(x-\mu_x)\right)C_2 \exp\left(-0.5(y-Ax-b)^T\Theta_y^{-1}(y-Ax-b)\right)dx
$$

$$
=Gaussian(y; A\mu_x + b, \Theta_y + A\Theta_x A^T)
$$

The predicted state probability at t=1
\n
$$
P(s_1 | o_0) = \int_{-\infty}^{\infty} P(s_0 | o_0) P(s_1 | s_0) ds_0
$$
\n
$$
P(s_1 | s_0) = Gaussian(s_1; \mu_{\varepsilon} + A_1 s_0, \Theta_{\varepsilon})
$$
\n
$$
P(s_0 | o_0) = Gaussian(s_0; \hat{s}_0, \hat{R}_0)
$$
\n
$$
P(s_1 | o_0) = \int_{-\infty}^{\infty} Gaussian(s_0; \hat{s}_0, \hat{R}_0) Gaussian(s_1; \mu_{\varepsilon} + A_1 s_0, \Theta_{\varepsilon}) ds_0
$$
\n
$$
P(s_1 | o_0) = Gaussian(s_1; A_1 \hat{s}_0 + \mu_{\varepsilon}, \Theta_{\varepsilon} + A_1 \hat{R}_0 A_1^T)
$$

Remains Gaussian

The updated state probability at T=1 $P(S_1 | o_{0:1}) = C R_{S_1} | o_0 R_{O_1} | s_1$ $P(s_1 | o_0) = Gaussian(s_1; A_1\hat{s}_0 + \mu_{\varepsilon}, \Theta_{\varepsilon} + A_1\hat{R}_0A_1^T)$ $P(o_1 | s_1) = Gaussian(o_1; \mu_\gamma + B_1 s_1, \Theta_\gamma)$ $P(s_1 | o_{0:1}) = Gaussian(s_1; \hat{s}_1, \hat{R}_1)$

The Kalman Filter!

Prediction at T

$$
P(s_t | o_{0:t-1}) = Gaussian(s_t; A_t \hat{s}_{t-1} + \mu_{\varepsilon}, \Theta_{\varepsilon} + A_t \hat{R}_{t-1} A_t^T)
$$

$$
P(s_t | o_{0:t}) = Gaussian(s_t; \overline{s}_t, R_t)
$$

 Update at T p

$$
P(s_t | o_{0:t}) =
$$

\n
$$
Gaussian(s_t; (R_t^{-1} + B_t^T \Theta_y^{-1} B_t)^{-1} (R_t^{-1} \overline{s}_t + B_t^T \Theta_y^{-1} (o_t - \mu_y)) (R_t^{-1} + B_t^T \Theta_y^{-1} B_t)^{-1})
$$

$$
P(s_t | o_{0:t}) = Gaussian(s_t; \hat{s}_t, \hat{R}_t)
$$

All distributions remain Gaussian

The Kalman filter

- The actual state estimate is the *mean* of the updated distribution
- **Predicted state at time to Predicted state at time formation**

$$
\overline{s}_{t} = mean[P(s_{t} | o_{0:t-1})] = A_{t} \hat{s}_{t-1} + \mu_{\varepsilon}
$$

■ Updated estimate of state at time *t*

$$
\hat{s}_t = \text{mean}[P(s_t \mid o_{0:t})] = (R_t^{-1} + B_t^T \Theta_y^{-1} B_t)^{-1} (R_t^{-1} \overline{s}_t + B_t^T \Theta_y^{-1} (o_t - \mu_y))
$$

Stable Estimation

ˆ $\hat{S}_t = mean[P(S_t | o_{0:t})] = (R_t^{-1} + B_t^T \Theta_y^{-1} B_t)^{-1} (R_t^{-1} \overline{S}_t + B_t^T \Theta_y^{-1} (O_t - \mu_y))$ $\overline{\gamma}_t^{-1} \overline{S}_t + B_t^T \Theta_{\gamma}^{-1} (O_t - \mu_{\gamma})$ $^{-1}$ $\mathbf{D}^T \mathbf{C}^{-1} \mathbf{D}$ *tT* $t \rightarrow t$ $\rightarrow t$ $\rightarrow t$ $\hat{S}_t = \text{mean}[P(s_t \mid o_{0:t})] = (R_t^{-1} + B_t^T \Theta_{\nu}^{-1} B_t)^{-1} (R_t^{-1} \overline{s}_t + B_t^T \Theta_{\nu}^{-1} (O_t^{-1} \overline{s}_t + B_t^{-1} \Theta_{\nu}^{-1} C_t^{-1})]$

- **The above equation fails if there is no** observation noisethe contract of the contract of
	- \Box Θ $_{\gamma} = 0$
	- □ Paradoxical?
	- □ Happens because we do not use the relationship between *o* and *s* effectively
- **Alternate derivation required** □ Conventional Kalman filter formulation

Estimating
$$
P(s | o)
$$

\n**Propping subscript to brevity**
\n $P(s | o_{0:t-1}) = Gaussian(s; \overline{s}, R)$
\n $o = Bs + \gamma$
\n $P(\gamma) = \frac{1}{\sqrt{(2\pi)^d |\Theta_{\gamma}|}} exp(-0.5\epsilon^T \Theta_{\gamma}^{-1} \epsilon)$

■ Define y as the noiseless version of o

<u>y = Bs</u>

■ Define the following extended vectors:

$$
y = Bs
$$
 $o = y + \gamma$

• Define the following extended vectors:

$$
Y = \begin{bmatrix} y \\ s \end{bmatrix} \quad O = \begin{bmatrix} o \\ s \end{bmatrix} \quad G = \begin{bmatrix} \gamma \\ 0 \end{bmatrix} \quad O = Y + G
$$

$$
P(G) = Gaussian \begin{bmatrix} \Theta_{\gamma} & 0 \\ 0 & 0 \end{bmatrix}
$$

The probability distribution of Y
\n
$$
y = Bs
$$

\n $Y = \begin{bmatrix} y \\ s \end{bmatrix}$
\n $P(s | o_{0:t-1}) = Gaussian(s; \overline{s}, R)$
\nSince s is Gaussian, Y is Gaussian

Expectation[y] = $E[Bs] = B\overline{s}$ $E[(y - E[y])(s - \overline{s})^T] = E[B(s - \overline{s})(s - \overline{s})^T] = BR$

 $P(Y | o_{0:t-1}) = Gaussian(Y; \mu_Y, \Theta_Y)$

$$
\mu_Y = \begin{bmatrix} B\overline{S} \\ \overline{S} \end{bmatrix}; \quad \Theta_Y = \begin{bmatrix} BRB^T & BR \\ RB^T & R \end{bmatrix}
$$

The probability distribution of O
\n
$$
O = Y + G
$$
\n
$$
P(G) = Gaussian \left(G; 0, \begin{bmatrix} \Theta_y & 0 \\ 0 & 0 \end{bmatrix} \right)
$$
\n
$$
P(Y \mid o_{0:t-1}) = Gaussian(Y; \mu_Y, \Theta_Y) = \begin{bmatrix} B\overline{S} \\ \overline{S} \end{bmatrix}; \quad \Theta_Y = \begin{bmatrix} BRB^T & BR \\ RB^T & R \end{bmatrix}
$$
\n
$$
P(O \mid o_{0:t-1}) = Gaussian(O; \mu_Y, \Theta_O) \quad \Theta_O = \begin{bmatrix} BRB^T + \Theta_y & BR \\ RB^T & R \end{bmatrix}
$$

- $\mathcal{L}_{\mathcal{A}}$ ■ The mean of the sum of independent Gaussian RVs is the sum of the means
- $\overline{\mathcal{A}}$ The covariance of the sum of independent Gaussian RVs is the sum of the covariances

The probability distribution of O

 $P(O | o_{0:t-1}) = P(o, s | o_{0:t-1}) = Gaussian(O; \mu_Y, \Theta_o)$

$$
C \exp\left(-0.5\left[(o-B\overline{s})\right)(s-\overline{s})\right]^T \left[\frac{BRB^T + \Theta_{\gamma}}{RB^T} + \frac{BR}{R}\right]^{-1} \left[\frac{o-B\overline{s}}{s-\overline{s}}\right]
$$

Net Writing it out in extended form

A matrix inverse identity
\n
$$
\begin{bmatrix} A & B \\ B^T & C \end{bmatrix}^{-1} = \begin{bmatrix} A^{-1} + A^{-1}B(C - B^T A^{-1}B)^{-1}B^T A^{-1} & -A^{-1}B(C - B^T A^{-1}B)^{-1} \\ -(C - B^T A^{-1}B)^{-1}B^T A^{-1} & (C - B^T A^{-1}B)^{-1} \end{bmatrix}
$$

Work it out..

Applying it to the inverse covariance of *O*:

$$
\begin{bmatrix} BRB^T + \Theta_{\gamma} & BR \\ RB^T & R \end{bmatrix}^{-1} = \begin{bmatrix} * & * \\ -(R - (BRB^T + \Theta_{\gamma}))^{-1}RB^T(BRB^T + \Theta_{\gamma})^{-1} & RB^T (R - BRB^T - \Theta_{\gamma})^{-1} BR \end{bmatrix}
$$

A matrix inverse identity
\n
$$
\begin{bmatrix} A & B \\ B^T & C \end{bmatrix}^{-1} = \begin{bmatrix} A^{-1} + A^{-1}B(C - B^T A^{-1}B)^{-1} B^T A^{-1} & -A^{-1}B(C - B^T A^{-1}B)^{-1} \\ -(C - B^T A^{-1}B)^{-1} B^T A^{-1} & (C - B^T A^{-1}B)^{-1} \end{bmatrix}
$$

Work it out..

Applying it to the inverse covariance of *O*:

$$
\begin{bmatrix} BRB^T + \Theta_{\gamma} & BR \\ RB^T & R \end{bmatrix}^{-1} = \begin{bmatrix} * & * \\ -(R - (BRB^T + \Theta_{\gamma}))^{-1} RB^T (BRB^T + \Theta_{\gamma})^{-1} & (RB^T (R - BRB^T - \Theta_{\gamma}))^{-1} BR \end{bmatrix}
$$

Conditional distribution from Gaussians

Given any jointly Gaussian variables x and y such that $P(x,y)$ is Gaussian

$$
P(x, y) = P\left[\begin{bmatrix} x \\ y \end{bmatrix}\right] = C \exp\left(-0.5\left[(x - \mu_x) \quad \left(y - \mu_y\right)\right]^T \begin{bmatrix} \Theta_{xx} & \Theta_{xy} \\ \Theta_{yx} & \Theta_{yy} \end{bmatrix}^{-1} \begin{bmatrix} x - \mu_x \\ y - \mu_y \end{bmatrix}\right)
$$

■ The *conditional* distribution P(y|x) is given by $\left(y;\mu_{\textnormal{\textit{y}}}+\Theta_{\textnormal{\textit{yx}}} \Theta_{\textnormal{\textit{x}x}}^{-1}(x-\mu_{\textnormal{\textit{x}}}), \;\;\Theta_{\textnormal{\textit{y}y}}-\Theta_{\textnormal{\textit{y}x}} \Theta_{\textnormal{\textit{x}x}}^{-1}\Theta_{\textnormal{\textit{x}x}}\right)$ $\bigg)$ $P(y|x) = Gaussian(y; \mu_{y} + \Theta_{yx} \Theta_{xx}^{-1}(x - \mu_{x}), \Theta_{yy} - \Theta_{yx} \Theta_{xx}^{-1} \Theta_{yy}$ $(y|x) = Gaussian(y; \mu_{y} + \Theta_{yx} \Theta_{xx}^{-1}(x - \mu_{x}), \Theta_{yy} - \Theta_{yx} \Theta_{xx}^{-1})$

Stable Estimation

$$
P(O | o0:t-1) = P(o, s | o0:t-1) = Gaussian(O; \muY, \Thetao)
$$

\n
$$
C \exp\left(-0.5[(o-B\overline{s}) (s-\overline{s})]^T \begin{bmatrix} BRB^T + \Theta_{\gamma} & BR\\ RB^T & R \end{bmatrix}^{-1} \begin{bmatrix} o-B\overline{s} \\ s-\overline{s} \end{bmatrix}\right)
$$

The conditional distribution of *s*

 $P(s\mid o_{0:t}) = Gaussian(s;(I-RB^T(BRB^T + \Theta_{\gamma})^{-1}B)\overline{s} + RB^T(BRB^T + \Theta_{\gamma})^{-1}o, (R-RB^T(BRB^T + \Theta_{\gamma})^{-1}BR))$ $1 \text{p} = \text{p}$ p (p p \text $S(S|o_{0:t}) = Gaussian(s;(I-RB'(BRB'+\Theta_{\gamma})^{-1}B)\overline{s} + RB'(BRB'+\Theta_{\gamma})^{-1}o,(R-RB'(BRB'+\Theta_{\gamma})^{-1}o)$

\blacksquare Note that we are not computing Θ_γ $^{\text{-1}}$ in this . formulation

The Kalman filter

- The actual state estimate is the *mean* of the updated distribution
- **Predicted state at time to Predicted state at time formation**

$$
\overline{s}_{t} = s_{t}^{pred} = mean[P(s_{t} | o_{0:t-1})] = A_{t} \hat{s}_{t-1} + \mu_{\varepsilon}
$$

■ Updated estimate of state at time *t*

 $P(s_t | o_{0:t}) = Gaussian(s; (I - RB^T (BRB^T + \Theta_y)^{-1}B)\overline{s} + RB^T (BRB^T + \Theta_y)^{-1}o, (R - RB^T (BRB^T + \Theta_y)^{-1}BR)$

$$
\hat{s}_{t} = \text{mean}[P(s_{t} \mid o_{0t})] = (I - R_{t} B_{t}^{T} (B_{t} R_{t} B_{t}^{T} + \Theta_{\gamma})^{-1} B_{t}) \bar{s}_{t} + R_{t} B_{t}^{T} (B_{t} R_{t} B_{t}^{T} + \Theta_{\gamma})^{-1} o_{t}
$$

The Kalman filter

Prediction

$$
\overline{s}_t = s_t^{pred} = mean[P(s_t | o_{0:t-1})] = A_t \hat{s}_{t-1} + \mu_{\varepsilon}
$$

$$
R_t = \Theta_{\varepsilon} + A_t \hat{R}_{t-1} A_t^T
$$

Update

$$
\hat{s}_t = \left(I - R_t B_t^T \left(B_t R_t B_t^T + \Theta_{\gamma}\right)^{-1} B_t \right) \overline{s}_t + R_t B_t^T \left(B_t R_t B_t^T + \Theta_{\gamma}\right)^{-1} o_t
$$
\n
$$
\hat{R}_t = R_t - R_t B_t^T \left(B_t R_t B_t^T + \Theta_{\gamma}\right)^{-1} B_t R_t
$$

The Kalman filter **Prediction**

$$
\overline{s}_t = A_t \hat{s}_{t-1} + \mu_{\varepsilon}
$$

$$
R_t = \Theta_{\varepsilon} + A_t \hat{R}_{t-1} A_t^T
$$

Update

$$
K_t = R_t B_t^T (B_t R_t B_t^T + \Theta_{\gamma})^{-1}
$$

$$
\hat{S}_t = \overline{S}_t + K_t (O_t - B_t \overline{S}_t)
$$

$$
\hat{R}_t = (I - K_t B_t) R_t
$$

The Kalman Filter

- Very popular for tracking the state of processes
	- □ Control systems
	- $\textcolor{red}{\mathsf{u}}$ Robotic tracking
		- **The Co** Simultaneous localization and mapping
	- □ Radars
	- □ Even the stock market..

■ What are the parameters of the process?

Kalman filter contd.

$$
s_{t} = A_{t} s_{t-1} + \varepsilon_{t}
$$

$$
o_{t} = B_{t} s_{t} + \gamma_{t}
$$

■ Model parameters A and B must be known

- Often the state equation includes an *additional* driving term: $\;{\bf s}_{\sf t} = {\sf A}_{\sf t}{\sf s}_{{\sf t}\text{-1}}+{\sf G}_{\sf t}{\sf u}_{\sf t}+{\varepsilon}_{\sf t}$
- \Box The parameters of the driving term must be known
- \blacksquare The initial state distribution must be known

Defining the parameters

- \mathbb{R}^3 State state must be carefully defined
	- $\texttt{w} \in \texttt{C}$. for a robotic vehicle, the state is an extended vector that includes the current velocity and acceleration
		- $S = [X, dX, d^2X]$
- State equation: Must incorporate appropriate constraints
	- □ If state includes acceleration and velocity, velocity at next time $=$ current velocity $+$ acc. $*$ time step

$$
\Box \text{St} = AS_{t-1} + e
$$

$$
A = [1 \t 0.5t^2; 0 1 t; 0 0 1]
$$

Parameters

- \mathbb{R}^2 Observation equation:
	- □ Critical to have accurate observation equation
	- □ Must provide a valid relationship between state and observations
- Observations typically high-dimensional
	- □ May have higher or lower dimensionality than state

Problems

$$
S_t = f(S_{t-1}, \mathcal{E}_t)
$$

$$
O_t = g(S_t, \gamma_t)
$$

- \blacksquare f() and/or g() may not be nice linear functions □ Conventional Kalman update rules for are no longer valid
- **E** and/or γ may not be Gaussian □ Gaussian based update rules no longer valid

Solutions (Next Tuesday)
\n
$$
s_{t} = f(s_{t-1}, \varepsilon_{t})
$$
\n
$$
o_{t} = g(s_{t}, \gamma_{t})
$$

 \blacksquare f() and/or g() may not be nice linear functions

- □ Conventional Kalman update rules for are no lon ger valid
- **Extended Kalman Filter**
- **E** and/or γ may not be Gaussian
	- □ Gaussian based update rules no longer valid
	- **Particle Filters**