Machine Learning for Signal Processing Expectation Maximization Mixture Models

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Learning Distributions for Data

• Problem: Given a collection of examples from some data, estimate its distribution

- Solution: Assign a model to the distribution
 Learn parameters of model from data
- Models can be arbitrarily complex
 - Mixture densities, Hierarchical models.

A Thought Experiment



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- A person shoots a loaded dice repeatedly
- You observe the series of outcomes
- You can form a good idea of how the dice is loaded
 - Figure out what the probabilities of the various numbers are for dice
- P(number) = count(number)/count(rolls)
- This is a *maximum likelihood* estimate
 - Estimate that makes the observed sequence of numbers most probable

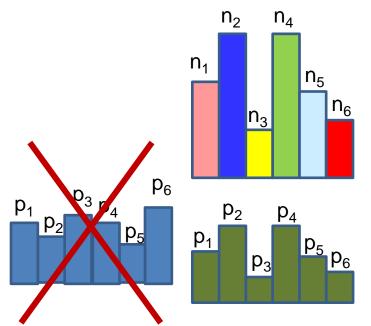
The Multinomial Distribution

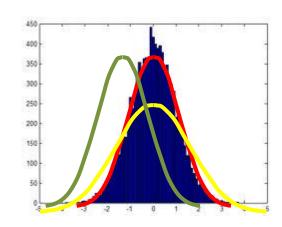
• A probability distribution over a discrete collection of items is a *Multinomial*

P(X : X belongs to a discrete set) = P(X)

- E.g. the roll of dice
 X : X in (1,2,3,4,5,6)
- Or the toss of a coin
 X : X in (head, tails)

Maximum Likelihood Estimation





- Basic principle: Assign a form to the distribution
 - E.g. a multinomial
 - Or a Gaussian
- Find the *distribution* that best fits the histogram of the data

Defining "Best Fit"

- The data are generated by draws from the distribution
 - I.e. the generating process draws from the distribution
- Assumption: The world is a boring place
 - The data you have observed are very typical of the process
- Consequent assumption: The distribution has a high probability of generating the observed data
 - Not necessarily true
- Select the distribution that has the *highest* probability of generating the data
 - Should assign lower probability to less frequent observations and vice versa

Maximum Likelihood Estimation: Multinomial

Probability of generating (n₁, n₂, n₃, n₄, n₅, n₆)

$$P(n_1, n_2, n_3, n_4, n_5, n_6) = Const \prod_i p_i^{n_i}$$

- Find $p_1, p_2, p_3, p_4, p_5, p_6$ so that the above is maximized
- Alternately maximize

$$\log(P(n_1, n_2, n_3, n_4, n_5, n_6)) = \log(Const) + \sum_i n_i \log(p_i)$$

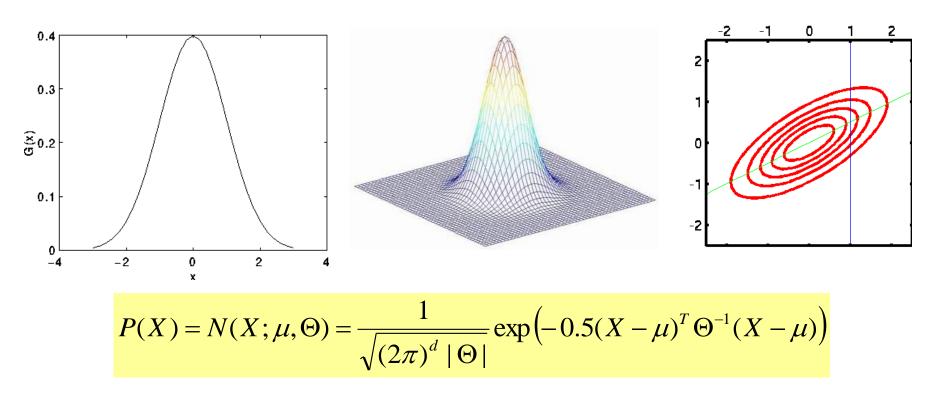
- Log() is a monotonic function
- $\operatorname{argmax}_{x} f(x) = \operatorname{argmax}_{x} \log(f(x))$
- Solving for the probabilities gives us
 - Requires constrained optimization to ensure probabilities sum to 1

$$p_i = \frac{n_i}{\sum_j n_j}$$



FVFNTUALLY

Segue: Gaussians



- Parameters of a Gaussian:
 - Mean μ , Covariance Θ

Maximum Likelihood: Gaussian

Given a collection of observations ($X_1, X_2,...$), estimate mean μ and covariance Θ

$$P(X_1, X_2, ...) = \prod_i \frac{1}{\sqrt{(2\pi)^d |\Theta|}} \exp\left(-0.5(X_i - \mu)^T \Theta^{-1}(X_i - \mu)\right)$$

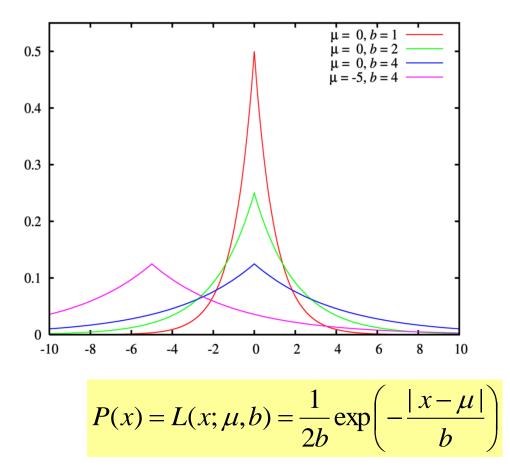
$$\log(P(X_1, X_2, ...)) = C - 0.5 \sum_{i} (\log(|\Theta|) + (X_i - \mu)^T \Theta^{-1} (X_i - \mu))$$

• Maximizing w.r.t μ and Θ gives us

$$\mu = \frac{1}{N} \sum_{i} X_{i} \qquad \Theta = \frac{1}{N} \sum_{i} (X_{i} - \mu) (X_{i} - \mu)^{T}$$

ITS STILL JUST COUNTING!

Laplacian



- Parameters: Median μ , scale *b* (*b* > 0)
 - $-\ \mu$ is also the mean, but is better viewed as the median

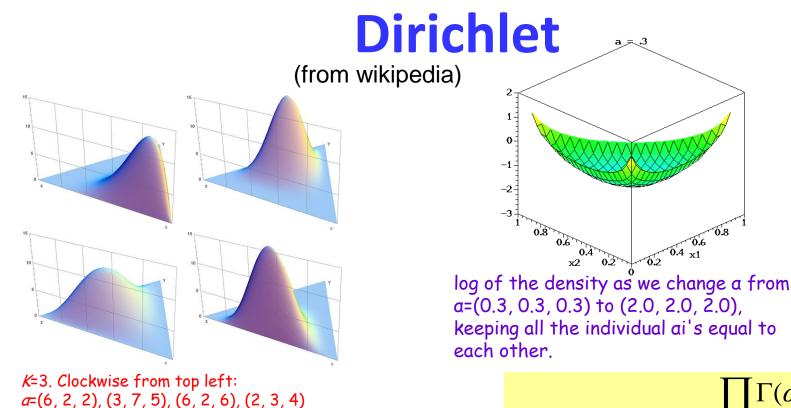
Maximum Likelihood: Laplacian

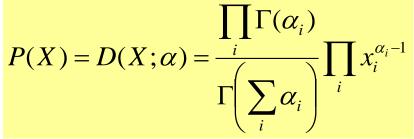
 Given a collection of observations (x₁, x₂,...), estimate mean μ and scale b

$$\log(P(x_1, x_2, ...)) = C - N \log(b) - \sum_i \frac{|x_i - \mu|}{b}$$

• Maximizing w.r.t μ and b gives us

$$\mu = median(\{x_i\}) \quad b = \frac{1}{N} \sum_i |x_i - \mu|$$





- Parameters are αs
 - Determine mode and curvature
- Defined only of probability vectors
 - $\label{eq: 1.1.1} \textbf{-} X = [x_1 \; x_2 \; .. \; x_K] \textbf{,} \; \Sigma_i \; x_i = 1, \;\; x_i >= 0 \; \text{for all} \; i$

Maximum Likelihood: Dirichlet

 Given a collection of observations (X₁, X₂,...), estimate α

$$\log(P(X_1, X_2, \dots)) = \sum_{j} \sum_{i} (\alpha_i - 1) \log(X_{j,i}) + N \sum_{i} \log(\Gamma(\alpha_i)) - N \log\left(\Gamma\left(\sum_{i} \alpha_i\right)\right)$$

- No closed form solution for α s.
 - Needs gradient ascent
- Several distributions have this property: the ML estimate of their parameters have no closed form solution

Continuing the Thought Experiment





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- Two persons shoot loaded dice repeatedly

 The dice are differently loaded for the two of them
- We observe the series of outcomes for both persons
- How to determine the probability distributions of the two dice?

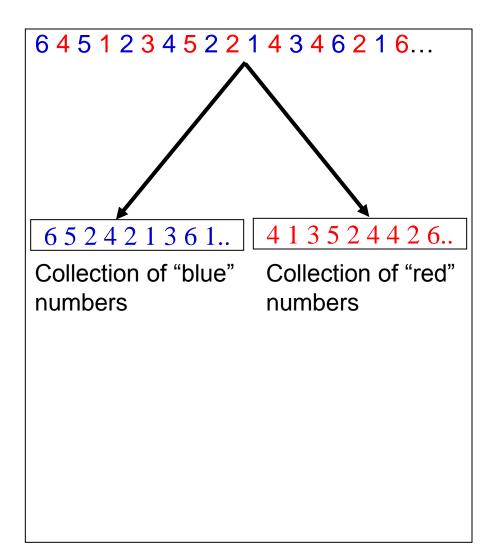
Estimating Probabilities

- Observation: The sequence of numbers from the two dice
 - As indicated by the colors, we know who rolled what number

6 4 5 1 2 3 4 5 2 2 1 4 3 4 6 2 1 6...

Estimating Probabilities

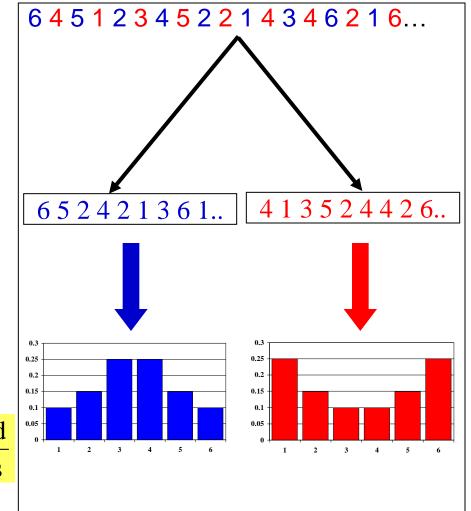
- Observation: The sequence of numbers from the two dice
 - As indicated by the colors, we know who rolled what number
- Segregation: Separate the blue observations from the red



Estimating Probabilities

- Observation: The sequence of numbers from the two dice
 - As indicated by the colors, we know who rolled what number
- Segregation: Separate the blue observations from the red
- From each set compute probabilities for each of the 6 possible outcomes

 $P(number) = \frac{\text{no. of times number was rolled}}{\text{total number of observed rolls}}$



A Thought Experiment



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44163212...

- Now imagine that you cannot observe the dice yourself
- Instead there is a "caller" who randomly calls out the outcomes
 - 40% of the time he calls out the number from the left shooter, and 60% of the time, the one from the right (and you know this)
- At any time, you do not know which of the two he is calling out
- How do you determine the probability distributions for the two dice?

A Thought Experiment



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44163212...

- How do you now determine the probability distributions for the two sets of dice ...
- If you do not even know what fraction of time the blue numbers are called, and what fraction are red?

A Mixture Multinomial

- The caller will call out a number X in any given callout IF
 - He selects "RED", and the Red die rolls the number X
 - OR
 - He selects "BLUE" and the Blue die rolls the number X
- P(X) = P(Red)P(X | Red) + P(Blue)P(X | Blue)
 E.g. P(6) = P(Red)P(6 | Red) + P(Blue)P(6 | Blue)
- A distribution that *combines* (or *mixes*) multiple multinomials is a *mixture* multinomial

$$P(X) = \sum_{Z} P(Z)P(X \mid Z)$$

Mixture weights Component multinomials

Mixture Distributions

Mixture Gaussian

 $P(X) = \sum P(Z)N(X; \mu_z, \Theta_z)$

$$P(X) = \sum_{Z} P(Z)P(X \mid Z)$$

Mixture weights Component distributions

Mixture of Gaussians and Laplacians

$$P(X) = \sum_{Z} P(Z)N(X; \mu_z, \Theta_z) + \sum_{Z} P(Z)\prod_{i} L(X_i; \mu_z, b_{z,i})$$

- Mixture distributions mix several component distributions
 - Component distributions may be of varied type
- Mixing weights must sum to 1.0
- Component distributions integrate to 1.0
- Mixture distribution integrates to 1.0

Maximum Likelihood Estimation

• For our problem: $P(X) = \sum_{Z} P(Z)P(X \mid Z)$ - Z = color of dice

$$P(n_1, n_2, n_3, n_4, n_5, n_6) = Const \prod_X P(X)^{n_X} = Const \prod_X \left(\sum_Z P(Z) P(X \mid Z) \right)^{n_X}$$

• Maximum likelihood solution: Maximize

$$\log(P(n_1, n_2, n_3, n_4, n_5, n_6)) = \log(Const) + \sum_X n_X \log\left(\sum_Z P(Z)P(X \mid Z)\right)$$

- No closed form solution (summation inside log)!
 - In general ML estimates for mixtures do not have a closed form
 - USE EM!

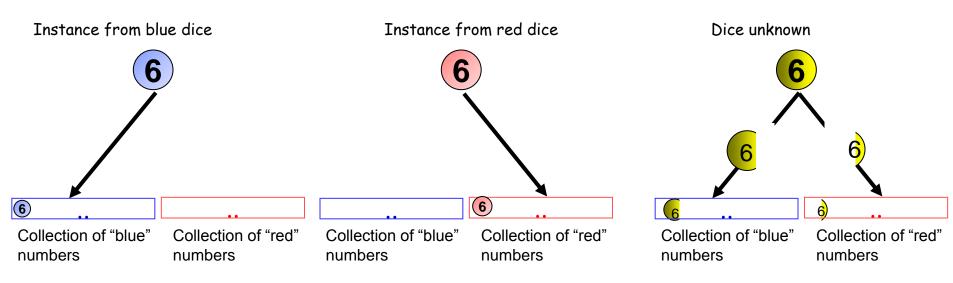
- It is possible to estimate all parameters in this setup using the Expectation Maximization (or EM) algorithm
- First described in a landmark paper by Dempster, Laird and Rubin
 - Maximum Likelihood Estimation from incomplete data, via the EM Algorithm, Journal of the Royal Statistical Society, Series B, 1977
- Much work on the algorithm since then
- The principles behind the algorithm existed for several years prior to the landmark paper, however.

- Iterative solution
- Get some initial estimates for all parameters
 - Dice shooter example: This includes probability distributions for dice AND the probability with which the caller selects the dice
- Two steps that are iterated:
 - *Expectation Step:* Estimate statistically, the values of unseen variables
 - Maximization Step: Using the estimated values of the unseen variables as truth, estimates of the model parameters

EM: The auxiliary function

- EM iteratively optimizes the following auxiliary function
- $Q(\theta, \theta') = \Sigma_Z P(Z|X, \theta') \log(P(Z, X | \theta))$
 - $-\mathbf{Z}$ are the unseen variables
 - Assuming Z is discrete (may not be)
- θ' are the parameter estimates from the previous iteration
- $\boldsymbol{\theta}$ are the estimates to be obtained in the current iteration

Expectation Maximization as counting



- Hidden variable: Z
 - Dice: The identity of the dice whose number has been called out
- If we knew Z for every observation, we could estimate all terms

 By adding the observation to the right bin
- Unfortunately, we do not know Z it is hidden from us!
- Solution: FRAGMENT THE OBSERVATION

Fragmenting the Observation

• EM is an iterative algorithm

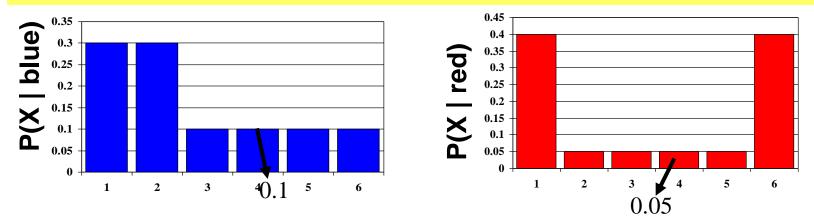
At each time there is a *current* estimate of parameters

- The "size" of the fragments is proportional to the *a posteriori probability* of the component distributions
 - The *a posteriori* probabilities of the various values of *Z* are computed using Bayes' rule:

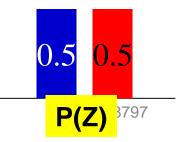
$$P(Z \mid X) = \frac{P(X \mid Z)P(Z)}{P(X)} = CP(X \mid Z)P(Z)$$

• Every dice gets a fragment of size P(dice | number)

- Hypothetical Dice Shooter Example:
- We obtain an initial estimate for the probability distribution of the two sets of dice (somehow):



• We obtain an initial estimate for the probability with which the caller calls out the two shooters (somehow)



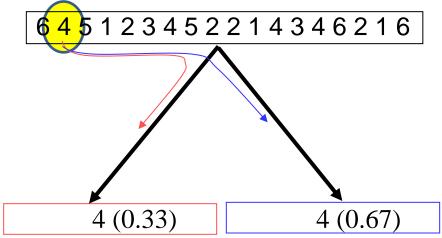
- Hypothetical Dice Shooter Example:
- Initial estimate:
 - P(blue) = P(red) = 0.5
 - P(4 | blue) = 0.1, for P(4 | red) = 0.05
- Caller has just called out 4
- Posterior probability of colors:

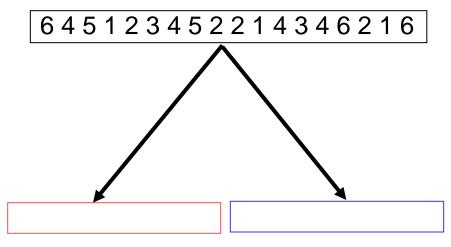
 $P(red | X = 4) = CP(X = 4 | Z = red)P(Z = red) = C \times 0.05 \times 0.5 = C 0.025$ $P(blue | X = 4) = CP(X = 4 | Z = blue)P(Z = blue) = C \times 0.1 \times 0.5 = C 0.05$

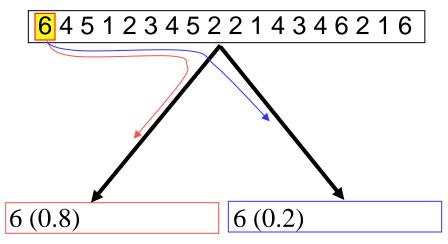
$$P(red | X = 4) = \frac{C \ 0.025}{C \ 0.025 + C \ 0.05}$$

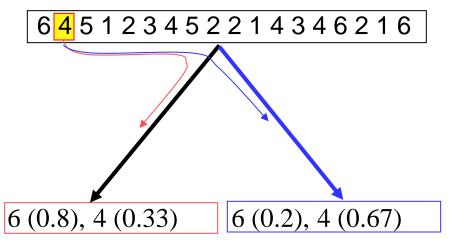
$$P(red | X = 4) = 0.33$$

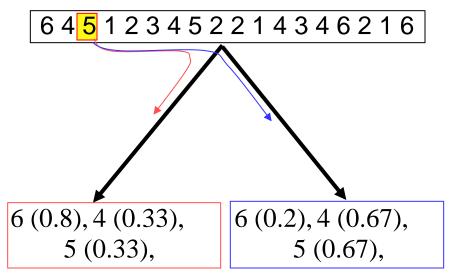
$$P(blue | X = 4) = 0.67$$

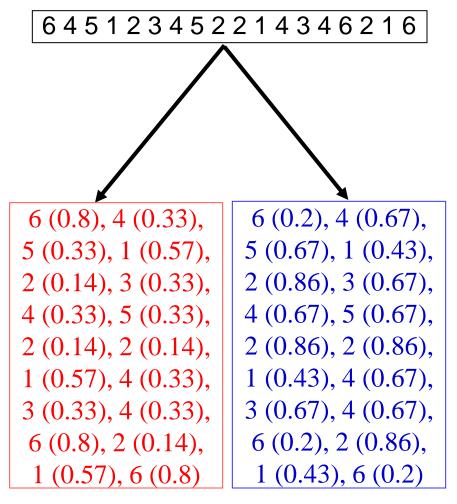












- Every observed roll of the dice contributes to both "Red" and "Blue"
- Total count for "Red" is the sum of all the posterior probabilities in the red column

- 7.31

- Total count for "Blue" is the sum of all the posterior probabilities in the blue column
 - 10.69
 - Note: 10.69 + 7.31 = 18 = the total number of instances

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6	.8	.2
4 5 1 2 3 4 5 2 2 2 1 4 3 4 6 2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Red" : 7.31
- Red:
 - Total count for 1: 1.71

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 4 5 1 2 3 4 5 2 2 2 2 1 4 3 4 3 4 6 2 2 1	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Red" : 7.31
- Red:
 - Total count for 1: 1.71
 - Total count for 2: 0.56

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 4 5 1 2 3 4 5 2 2 2 1 4 5 2 2 1 4 3 4 5 2 2 1 4 6 2 1 6	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Red" : 7.31
- Red:
 - Total count for 1: 1.71
 - Total count for 2: 0.56
 - Total count for 3: 0.66

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
6 4 5 1 2 3 4 5 2 2 2 1 4 3 4	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 2 1	.8	.2
2	.14	.86
	.57	.43
6	.8	.2



- Total count for "Red" : 7.31
- Red:
 - Total count for 1: 1.71
 - Total count for 2: 0.56
 - Total count for 3: 0.66
 - Total count for 4: 1.32

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6	.8	.2
6 4 5 1 2 3 4 5 2 2 2 1 4 3 4 3 4 3 4 6 2 1 6 2 1 6	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Red" : 7.31
- Red:
 - Total count for 1: 1.71
 - Total count for 2: 0.56
 - Total count for 3: 0.66
 - Total count for 4: 1.32
 - Total count for 5: 0.66

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6	.8	.2
6 4 5 1 2 3 4 5 2 2 2 1 4 3 4 3 4 3 4 6 2 1 6	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Red" : 7.31
- Red:
 - Total count for 1: 1.71
 - Total count for 2: 0.56
 - Total count for 3: 0.66
 - Total count for 4: 1.32
 - Total count for 5: 0.66
 - Total count for 6: 2.4

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 4 5 1 2 3 4 5 2 1 4 5 2 1 4 3 4 3 4 6 2 1	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Red" : 7.31
- Red:
 - Total count for 1: 1.71
 - Total count for 2: 0.56
 - Total count for 3: 0.66
 - Total count for 4: 1.32
 - Total count for 5: 0.66
 - Total count for 6: 2.4

Updated probability of Red dice:

- P(1 | Red) = 1.71/7.31 = 0.234
- P(2 | Red) = 0.56/7.31 = 0.077
- P(3 | Red) = 0.66/7.31 = 0.090
- P(4 | Red) = 1.32/7.31 = 0.181
- P(5 | Red) = 0.66/7.31 = 0.090
- P(6 | Red) = 2.40/7.31 = 0.328

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
5 1 2 3 4 5 2 2 2 1 4	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6	.8	.2
6 2 1	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Blue" : 10.69
- Blue:
 - Total count for 1: 1.29

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 4 5 1 2 3 4 5 2 2 2 2 1 4 3 4 3 4 6 2 2 1	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Blue" : 10.69
- Blue:
 - Total count for 1: 1.29
 - Total count for 2: 3.44

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
4 5 1 2 3 4 5 2 2 2 1 4 3 4 3 4 6 2 1 6	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Blue" : 10.69
- Blue:
 - Total count for 1: 1.29
 - Total count for 2: 3.44
 - Total count for 3: 1.34

Called	P(red X)	P(blue X)
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1	.57	.43
2	.14	.86
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4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 4 5 1 2 3 4 5 2 2 2 1 4 5 2 2 1 4 5 2 2 1 4 6 2 1 6	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Blue" : 10.69
- Blue:
 - Total count for 1: 1.29
 - Total count for 2: 3.44
 - Total count for 3: 1.34
 - Total count for 4: 2.68

Called	P(red X)	P(blue X)
	.8	.2
6 4 5 1 2 3 4 5 2 2 2 1 4 3 4 3 4 6 2 1 6 2 1 6	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Blue" : 10.69
- Blue:
 - Total count for 1: 1.29
 - Total count for 2: 3.44
 - Total count for 3: 1.34
 - Total count for 4: 2.68
 - Total count for 5: 1.34

Called	P(red X)	P(blue X)
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1	.57	.43
2	.14	.86
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4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
5 1 2 3 4 5 2 2 1 4 3 4 3 4 6 2 1 4 3 4 1	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Blue" : 10.69
- Blue:
 - Total count for 1: 1.29
 - Total count for 2: 3.44
 - Total count for 3: 1.34
 - Total count for 4: 2.68
 - Total count for 5: 1.34
 - Total count for 6: 0.6

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 4 5 1 2 3 4 5 2 2 2 1 4 3 4 5 2 1 4 3 4 6 2 1 6 2 1 6	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



- Total count for "Blue" : 10.69
- Blue:
 - Total count for 1: 1.29
 - Total count for 2: 3.44
 - Total count for 3: 1.34
 - Total count for 4: 2.68
 - Total count for 5: 1.34
 - Total count for 6: 0.6
- Updated probability of Blue dice:
 - P(1 | Blue) = 1.29/11.69 = 0.122
 - P(2 | Blue) = 0.56/11.69 = 0.322
 - P(3 | Blue) = 0.66/11.69 = 0.125
 - P(4 | Blue) = 1.32/11.69 = 0.250
 - P(5 | Blue) = 0.66/11.69 = 0.125
 - P(6 | Blue) = 2.40/11.69 = 0.056

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5 1 2 3 4 5 2 2 2 1 4	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
	.33	.67
6 2 1 6	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2

7.31



10.69

- Total count for "Red" : 7.31
- Total count for "Blue" : 10.69
- Total instances = 18
 - Note 7.31+10.69 = 18
- We also revise our estimate for the probability that the caller calls out Red or Blue
 - i.e the fraction of times that he calls Red and the fraction of times he calls Blue
- P(Z=Red) = 7.31/18 = 0.41
- P(Z=Blue) = 10.69/18 = 0.59

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
6 4 5 1 2 3 4 5 2 2 2 1 4 3 4	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6 2 1	.8	.2
2	.14	.86
1	.57	.43
6	.8	.2



The updated values

Probability of Red dice: P(1 | Red) = 1.71/7.31 = 0.234 P(2 | Red) = 0.56/7.31 = 0.077P(3 | Red) = 0.66/7.31 = 0.090P(4 | Red) = 1.32/7.31 = 0.181 P(5 | Red) = 0.66/7.31 = 0.090P(6 | Red) = 2.40/7.31 = 0.328Probability of Blue dice: P(1 | Blue) = 1.29/11.69 = 0.122 P(2 | Blue) = 0.56/11.69 = 0.322P(3 | Blue) = 0.66/11.69 = 0.125P(4 | Blue) = 1.32/11.69 = 0.250 P(5 | Blue) = 0.66/11.69 = 0.125P(6 | Blue) = 2.40/11.69 = 0.056

- P(Z=Red) = 7.31/18 = 0.41
- P(Z=Blue) = 10.69/18 = 0.59

THE UPDATED VALUES CAN BE USED TO REPEAT THE PROCESS. ESTIMATION IS AN ITERATIVE PROCESS

Called	P(red X)	P(blue X)
6	.8	.2
4	.33	.67
5	.33	.67
1	.57	.43
2	.14	.86
3	.33	.67
4	.33	.67
5	.33	.67
2	.14	.86
2	.14	.86
1	.57	.43
4	.33	.67
3	.33	.67
4	.33	.67
6	.8	.2
2	.14	.86
4 5 1 2 3 4 5 2 2 2 1 4 3 4 3 4 6 2 1 6 2 1 6	.57	.43
6	.8	.2

The Dice Shooter Example



63154124...

44163212...

- 1. Initialize P(Z), $P(X \mid Z)$
- 2. Estimate *P*(*Z* | *X*) for each *Z*, for each called out number
 - Associate X with each value of Z, with weight P(Z | X)
- 3. Re-estimate *P*(*X* | *Z*) for every value of *X* and *Z*
- 4. Re-estimate P(Z)
- 5. If not converged, return 20 297

In Squiggles

- Given a sequence of observations O₁, O₂, ...
 - $-N_{\rm X}$ is the number of observations of number X
- Initialize P(Z), P(X|Z) for dice Z and numbers X
- Iterate:

– For each number X:

$$P(Z \mid X) = \frac{P(X \mid Z)P(Z)}{\sum_{Z'} P(Z')P(X \mid Z')}$$

- Update:

$$P(X \mid Z) = \frac{\sum_{\substack{O \text{ such that } O = = X\\O}} P(Z \mid X)}{\sum_{\substack{O \in X\\O}} P(Z \mid O)} = \frac{N_X P(Z \mid X)}{\sum_{\substack{X\\V}} N_X P(Z \mid X)} P(Z \mid X)$$

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Solutions may not be unique

- The EM algorithm will give us one of many solutions, all equally valid!
 - The probability of 6 being called out:

 $P(6) = \alpha P(6 | red) + \beta P(6 | blue) = \alpha P_r + \beta P_b$

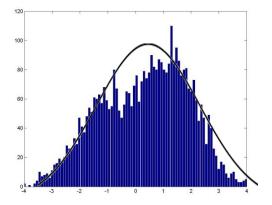
- Assigns P_r as the probability of 6 for the red die
- Assigns P_b as the probability of 6 for the blue die
- The following too is a valid solution [FIX]

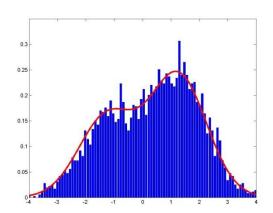
 $P(6) = 1.0(\alpha P_r + \beta P_b) + 0.0 anything$

- Assigns 1.0 as the a priori probability of the red die
- Assigns 0.0 as the probability of the blue die
- The solution is NOT unique

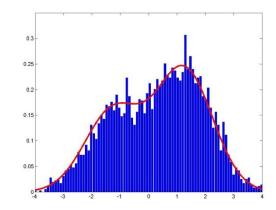
A more complex model: Gaussian mixtures

- A Gaussian mixture can represent data distributions far better than a simple Gaussian
- The two panels show the histogram of an unknown random variable
- The first panel shows how it is modeled by a simple Gaussian
- The second panel models the histogram by a mixture of two Gaussians
- Caveat: It is hard to know the optimal number of Gaussians in a mixture



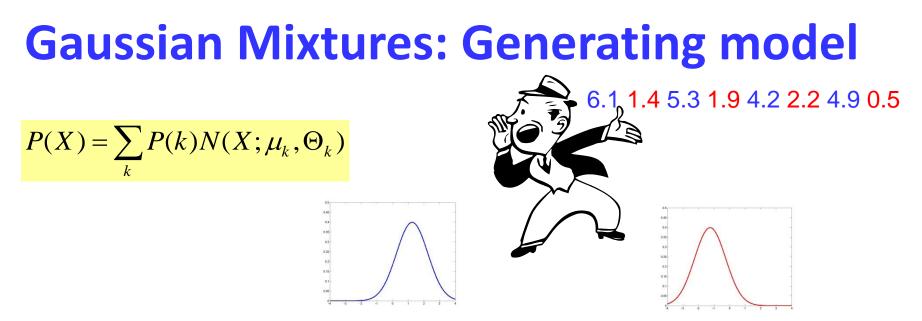


A More Complex Model



$$P(X) = \sum_{k} P(k)N(X; \mu_{k}, \Theta_{k}) = \sum_{k} \frac{P(k)}{\sqrt{(2\pi)^{d} |\Theta_{k}|}} \exp\left(-0.5(X - \mu_{k})^{T} \Theta_{k}^{-1}(X - \mu_{k})\right)$$

- Gaussian mixtures are often good models for the distribution of multivariate data
- Problem: Estimating the parameters, given a collection of data



- The caller now has two Gaussians
 - At each draw he randomly selects a Gaussian, by the mixture weight distribution
 - He then draws an observation from that Gaussian
 - Much like the dice problem (only the outcomes are now real numbers and can be anything)

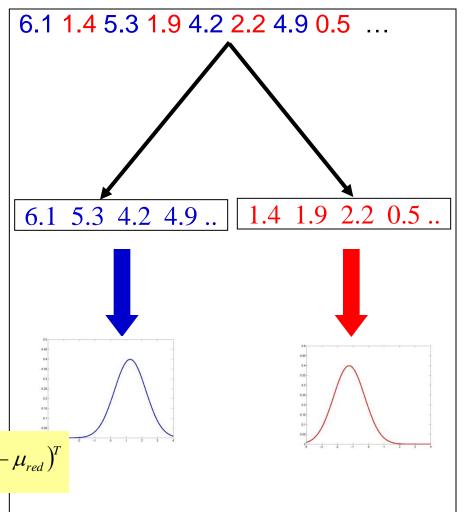
Estimating GMM with complete information

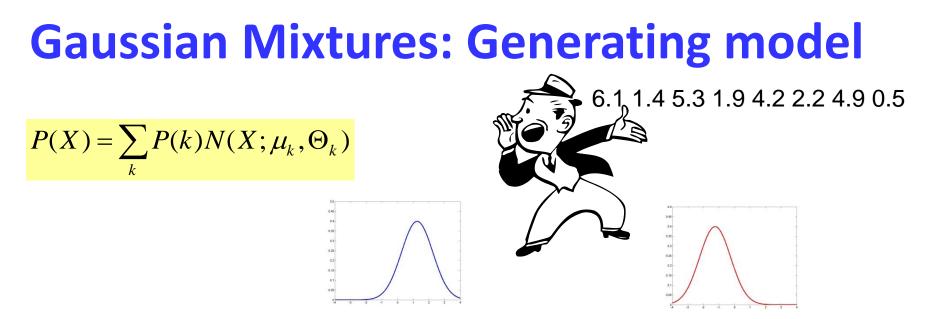
- Observation: A collection of numbers drawn from a mixture of 2 Gaussians
 - As indicated by the colors, we know which Gaussian generated what number
- Segregation: Separate the blue observations from the red
- From each set compute parameters for that Gaussian

 $P(red) = \frac{IN_{red}}{IN_{red}}$

$$\mu_{red} = \frac{1}{N_{red}} \sum_{i \in red} X_i$$

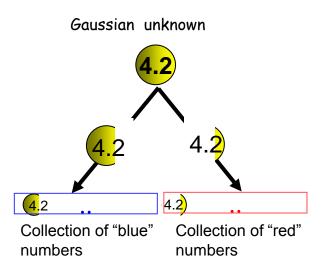
$$\Theta_{red} = \frac{1}{N_{red}} \sum_{i \in red} (X_i - \mu_{red}) (X_i - \mu_{red})^T$$





- Problem: In reality we will not know which Gaussian any observation was drawn from..
 - The color information is missing

Fragmenting the observation



- The identity of the Gaussian is not known!
- Solution: Fragment the observation
- Fragment size proportional to *a posteriori* probability

$$P(k \mid X) = \frac{P(X \mid k)P(k)}{\sum_{k'} P(k')P(X \mid k')} = \frac{P(k)N(X; \mu_k, \Theta_k)}{\sum_{k'} P(k')N(X; \mu_{k'}, \Theta_{k'})}$$

- Initialize P(k), μ_k and Θ_k for both Gaussians
 - Important how we do this
 - Typical solution: Initialize means randomly, Θ_k as the global covariance of the data and P(k) uniformly
- Compute fragment sizes for each Gaussian, for each observation

Number	P(red X)	P(blue X)
6.1	.81	.19
1.4	.33	.67
5.3	.75	.25
1.9	.41	.59
4.2	.64	.36
2.2	.43	.57
4.9	.66	.34
0.5	.05	.95

$$P(k \mid X) = \frac{P(k)N(X; \mu_k, \Theta_k)}{\sum_{k'} P(k')N(X; \mu_{k'}, \Theta_{k'})}$$

- Each observation contributes only as much as its fragment size to each statistic
- Mean(red) =

 (6.1*0.81 + 1.4*0.33 + 5.3*0.75 +
 1.9*0.41 + 4.2*0.64 + 2.2*0.43 + 4.9*0.66
 + 0.5*0.05) /
 (0.81 + 0.33 + 0.75 + 0.41 + 0.64 + 0.43 +
 0.66 + 0.05)
 = 17.05 / 4.08 = 4.18

Number	P(red X)	P(blue X)
6.1	.81	.19
1.4	.33	.67
5.3	.75	.25
1.9	.41	.59
4.2	.64	.36
2.2	.43	.57
4.9	.66	.34
0.5	.05	.95

4.08 3.92

Var(red) = $((6.1-4.18)^{2*}0.81 + (1.4-4.18)^{2*}0.33 + (5.3-4.18)^{2*}0.75 + (1.9-4.18)^{2*}0.41 + (4.2-4.18)^{2*}0.64 + (2.2-4.18)^{2*}0.43 + (4.9-4.18)^{2*}0.66 + (0.5-4.18)^{2*}0.05) / (0.81 + 0.33 + 0.75 + 0.41 + 0.64 + 0.43 + 0.66 + 0.05)$

$$P(red) = \frac{4.08}{8}$$

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EM for Gaussian Mixtures

- 1. Initialize P(k), μ_k and Θ_k for all Gaussians
- 2. For each observation X compute *a posteriori* probabilities for all Gaussian

$$P(k \mid X) = \frac{P(k)N(X; \mu_k, \Theta_k)}{\sum_{k'} P(k')N(X; \mu_{k'}, \Theta_{k'})}$$

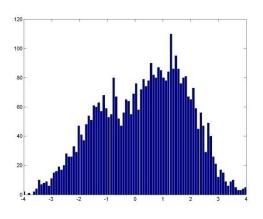
3. Update mixture weights, means and variances for all Gaussians

$$P(k) = \frac{\sum_{X} P(k|X)}{N} \qquad \mu_k = \frac{\sum_{X} P(k|X) X}{\sum_{X} P(k|X)} \qquad \Theta_k = \frac{\sum_{X} P(k|X) (X - \mu_k)^2}{\sum_{X} P(k|X)}$$

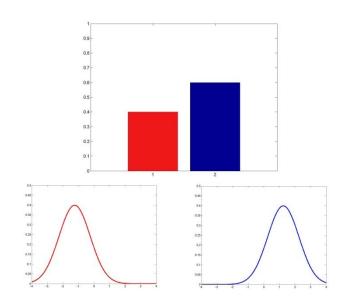
4. If not converged, return to 2

EM estimation of Gaussian Mixtures

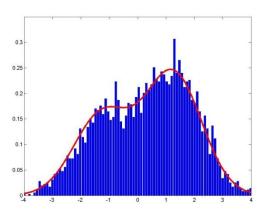
• An Example



Histogram of 4000 instances of a randomly generated data



Individual parameters of a two-Gaussian mixture estimated by EM



Two-Gaussian mixture estimated by EM

- The same principle can be extended to mixtures of other distributions.
- E.g. Mixture of Laplacians: Laplacian parameters become

$$\mu_{k} = median(P(k \mid x)) \quad b_{k} = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) \mid x - \mu_{k} \mid x = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) \mid x - \mu_{k} \mid x = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid x)} \sum_{x} P(k \mid x) = \frac{1}{\sum_{x} P(k \mid x)} \sum_{x} P(k \mid$$

 In a mixture of Gaussians and Laplacians, Gaussians use the Gaussian update rules, Laplacians use the Laplacian rule

- The EM algorithm is used whenever proper statistical analysis of a phenomenon requires the knowledge of a hidden or missing variable (or a set of hidden/missing variables)
 - The hidden variable is often called a "latent" variable
- Some examples:
 - Estimating mixtures of distributions
 - Only data are observed. The individual distributions and mixing proportions must both be learnt.
 - Estimating the distribution of data, when some attributes are missing
 - Estimating the dynamics of a system, based only on observations that may be a complex function of system state

Solve this problem:

- Problem 1:
 - Caller rolls a dice and flips a coin

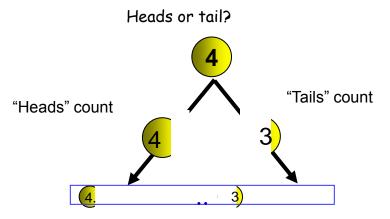


- Otherwise he calls the number+1
- Determine p(heads) and p(number) for the dice from a collection of outputs
- Problem 2:
 - Caller rolls two dice
 - He calls out the sum



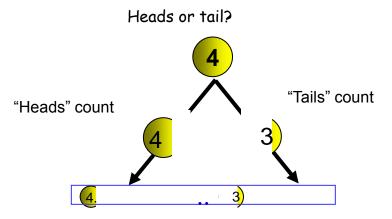


The dice and the coin



• Unknown: Whether it was head or tails

The dice and the coin

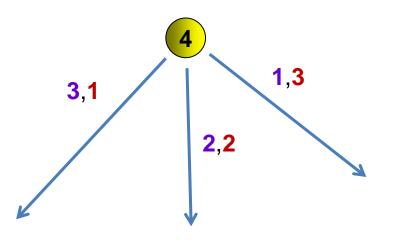


• Unknown: Whether it was head or tails

 $P(heads \mid N) = \frac{P(N)P(heads)}{P(N)P(heads) + P(N-1)P(tails)}$

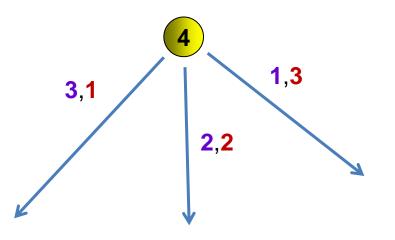
count(N) = #N.P(heads | N) + #(N-1).P(tails | N-1)

The two dice



- Unknown: How to partition the number
- Count_{blue}(3) += P(3,1 | 4)
- Count_{blue}(2) += P(2,2 | 4)
- Count_{blue}(1) += P(1,3 | 4)

The two dice

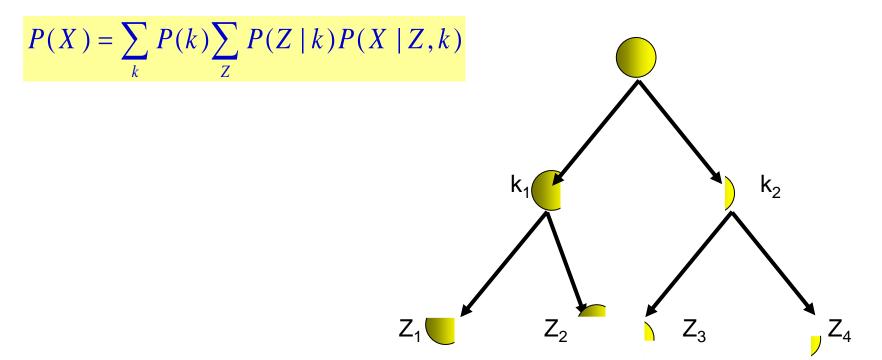


• Update rules

$$P(N, K-N \mid K) = \frac{P_1(N)P_2(K-N)}{\sum_{J=1}^{6} P_1(J)P_2(K-J)}$$

$$count_1(N) = \sum_{K=2}^{12} \# K \cdot P(N, K - N \mid K)$$
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Fragmentation can be hierarchical



- E.g. mixture of mixtures
- Fragments are further fragmented..
 - Work this out

More later

- Will see a couple of other instances of the use of EM
- EM for signal representation: PCA and factor analysis
- EM for signal separation
- EM for parameter estimation

• EM for homework..