# Learning P-maps <br> <br> Param. Learning 

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Graphical Models - 10708
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## Perfect maps (P-maps)

- I-maps are not unique and often not simple enough

■ Define "simplest" $G$ that is I-map for $P$
$\square \mathrm{ABN}$ structure $G$ is a perfect map for a distribution $P$ if $I(P)=I(G)$

- Our goal:
$\square$ Find a perfect map!
$\square$ Must address equivalent BNs


## Inexistence of P-maps 1

- XOR (this is a hint for the homework)
$A=B X O R C$
$A \perp B|7 A \perp B| C$
$B \perp C$
$C \perp A>B \perp C \mid A$

P-map?
extra credit

## Obtaining a P-map

- Given the independence assertions that are true for $P$
- Assume that there exists a perfect map $\mathrm{G}^{*}$
$\square$ Want to find G*

- Many structures may encode same independencies as $\mathrm{G}^{*}$, when are we done?

Find all equivalent structures simultaneously!

## I-Equivalence

- Two graphs $G_{1}$ and $G_{2}$ are I-equivalent if $\mathrm{I}\left(G_{1}\right)=\mathrm{I}\left(G_{2}\right)$
- Equivalence class of BN structures

Mutually-exclusive and exhaustive partition of graphs

- How do we characterize these equivalence classes?


## Skeleton of a BN

Skeleton of a BN structure G is an undirected graph over the same variables that has an edge $X-Y$ for every $X \rightarrow Y$ or $Y \rightarrow X$ in $G$

- (Little) Lemma: Two I -equivalent BN structures must have the same skeleton



# What about V-structures? <br> - V-structures are key property of BN structure 



- Theorem: If $G_{1}$ and $G_{2}$ have the same skeleton and $V$-structures, then $G_{1}$ and $G_{2}$ are I-equivalent


## Same V-structures not necessary

- Theorem: If $G_{1}$ and $G_{2}$ have the same skeleton and $V$-structures, then $G_{1}$ and $G_{2}$ are I-equivalent
- Though sufficient, same V-structures not necessary


## Immoralities \& I-Equivalence

- Key concept not V-structures, but "immoralities" (unmarried parents $)$ ) $X \rightarrow Z \leftarrow Y$, with no arrow between $X$ and $Y$
$\square$ Important pattern: X and Y independent given their parents, but not given $Z$
$\square$ (If edge exists between X and Y , we have covered the V-structure)
- Theorem: $G_{1}$ and $G_{2}$ have the same skeleton and immoralities if and only if $G_{1}$ and $G_{2}$ are I-equivalent


## Obtaining a P-map

- Given the independence assertions that are true for $P$
$\square$ Obtain skeleton
$\square$ Obtain immoralities
- From skeleton and immoralities, obtain every (and any) BN structure from the equivalence class


## Identifying the skeleton 1

When is there an edge between X and Y ?

- When is there no edge between $X$ and $Y$ ?


## Identifying the skeleton 2

- Assume d is max number of parents ( d could be n )
- For each $X_{i}$ and $X_{j}$
$\square \mathrm{E}_{\mathrm{ij}} \leftarrow$ true
$\square$ For each $\mathbf{U} \subseteq \mathbf{X}-\left\{\mathrm{X}_{\mathrm{i}}, \mathrm{X}_{\mathrm{j}}\right\},|\mathbf{U}| \leq \mathrm{d}$
- Is $\left(X_{i} \perp X_{j} \mid \mathrm{U}\right)$ ?
$\square \mathrm{E}_{\mathrm{ij}} \leftarrow$ false
$\square$ If $\mathrm{E}_{\mathrm{ij}}$ is true
- Add edge X - Y to skeleton


## Identifying immoralities

- Consider $\mathrm{X}-\mathrm{Z}-\mathrm{Y}$ in skeleton, when should it be an immorality?

■ Must be $X \rightarrow Z \leftarrow Y$ (immorality):
$\square$ When $X$ and $Y$ are never independent given $\mathbf{U}$, if $Z \in \mathbf{U}$

- Must not be $X \rightarrow Z \leftarrow Y$ (not immorality):
$\square$ When there exists $\mathbf{U}$ with $\mathrm{Z} \in \mathbf{U}$, such that X and Y are independent given $\mathbf{U}$


## From immoralities and skeleton to BN structures

- Representing BN equivalence class as a partially-directed acyclic graph (PDAG)
- Immoralities force direction on some other BN edges
- Full (polynomial-time) procedure described in reading


## What you need to know

Minimal I-map
$\square$ every $P$ has one, but usually many

- Perfect map
$\square$ better choice for BN structure
not every $P$ has one
can find one (if it exists) by considering l-equivalence
$\square$ Two structures are l-equivalent if they have same skeleton and immoralities


## Announcements

- Recitation tomorrow
$\square$ Don't miss it!
- No class on Monday © $^{\circ}$


## Review

- Bayesian Networks

Compact representation for probability distributions
$\square$ Exponential reduction in number of parametersExploits independencies


- Next - Learn BNs
parameters
$\square$ structure


## Thumbtack - Binomial Distribution

- $P($ Heads $)=\theta, P($ Tails $)=1-\theta$

Flips are i.i.d.:
$\square$ Independent events
$\square$ Identically distributed according to Binomial distribution

- Sequence $D$ of $\alpha_{H}$ Heads and $\alpha_{T}$ Tails

$$
P(\mathcal{D} \mid \theta)=\theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}}
$$

## Maximum Likelihood Estimation

- Data: Observed set $D$ of $\alpha_{H}$ Heads and $\alpha_{T}$ Tails
- Hypothesis: Binomial distribution
- Learning $\theta$ is an optimization problem
$\square$ What's the objective function?
- MLE: Choose $\theta$ that maximizes the probability of observed data:

$$
\begin{aligned}
\hat{\theta} & =\arg \max _{\theta} P(\mathcal{D} \mid \theta) \\
& =\arg \max _{\theta} \ln P(\mathcal{D} \mid \theta)
\end{aligned}
$$

## Your first learning algorithm

$\begin{aligned} \hat{\theta} & =\arg \max _{\theta} \ln P(\mathcal{D} \mid \theta) \\ & =\arg \max _{\theta} \ln \theta^{\alpha_{H}}(1-\theta)^{\alpha_{T}}\end{aligned}$

- Set derivative to zero: $\quad \frac{d}{d \theta} \ln P(\mathcal{D} \mid \theta)=0$



Maximum likelihood estimation (MLE) of BN parameters - example

- Given structure, log likelihood of data:
$\log P\left(\mathcal{D} \mid \theta_{\mathcal{G}}, \mathcal{G}\right)$

Maximum likelihood estimation (MLE) of BN parameters - General case

- Data: $\mathbf{x}^{(1)}, \ldots, \mathbf{x}^{(m)}$
- Restriction: $\mathbf{x}^{(\mathrm{i})}\left[\mathrm{Pa}_{\mathrm{x}_{\mathrm{i}}}\right] \rightarrow$ assignment to $\mathrm{Pa}_{\mathrm{xi}_{\mathrm{i}}}$ in $\mathbf{x}^{(\mathrm{j})}$
- Given structure, log likelihood of data:
$\log P\left(\mathcal{D} \mid \theta_{\mathcal{G}}, \mathcal{G}\right)$

Taking derivatives of MLE of BN
parameters - General case
$\log P\left(\mathcal{D} \mid \theta_{G}, \mathcal{G}\right)=\sum_{j=1}^{m} \sum_{i=1}^{n} \log P\left(X_{i}=x_{i}^{(j)} \mid \mathbf{P a}_{X_{i}}=\mathbf{x}^{(j)}\left[{ }^{\left[\mathrm{Pa} x_{i}\right]}\right)\right.$

## General MLE for a CPT

- Take a CPT: P(X|U)

■ Log likelihood term for this CPT

- Parameter $\theta_{X=x \mid \mathbf{U}=\mathrm{u}}$ :

MLE: $\quad P(X=x \mid \mathbf{U}=\mathbf{u})=\theta_{X=x \mid \mathbf{U}=\mathbf{u}}=\frac{\operatorname{Count}(X=x, \mathbf{U}=\mathbf{u})}{\operatorname{Count}(\mathbf{U}=\mathbf{u})}$

## Can we really trust MLE?

- What is better?
$\square 3$ heads, 2 tails
$\square 30$ heads, 20 tails
$\square 3 \times 10^{23}$ heads, $2 \times 10^{23}$ tails

- Many possible answers, we need distributions over possible parameters


## Bayesian Learning

- Use Bayes rule:
$P(\theta \mid \mathcal{D})=\frac{P(\mathcal{D} \mid \theta) P(\theta)}{P(\mathcal{D})}$
- Or equivalently:
$P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta) P(\theta)$


## Bayesian Learning for Thumbtack

$$
P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta) P(\theta)
$$

- Likelihood function is simply Binomial:

$$
P(\mathcal{D} \mid \theta)=\theta^{m_{H}}(1-\theta)^{m_{T}}
$$

- What about prior?
$\square$ Represent expert knowledge
$\square$ Simple posterior form
- Conjugate priors:
$\square$ Closed-form representation of posterior (more details soon)
$\square$ For Binomial, conjugate prior is Beta distribution


## Beta prior distribution - $P(\theta)$

$$
P(\theta)=\frac{\theta^{\alpha_{H}-1}(1-\theta)^{\alpha_{T}-1}}{B\left(\alpha_{H}, \alpha_{T}\right)} \sim \operatorname{Beta}\left(\alpha_{H}, \alpha_{T}\right)
$$



- Likelihood function: $P(\mathcal{D} \mid \theta)=\theta^{m_{H}}(1-\theta)^{m_{T}}$
- Posterior: $P(\theta \mid \mathcal{D}) \propto P(\mathcal{D} \mid \theta) P(\theta)$


## Posterior distribution

- Prior: $\operatorname{Beta}\left(\alpha_{H}, \alpha_{T}\right)$
- Data: $m_{H}$ heads and $m_{T}$ tails
- Posterior distribution:

$$
P(\theta \mid \mathcal{D}) \sim \operatorname{Beta}\left(m_{H}+\alpha_{H}, m_{T}+\alpha_{T}\right)
$$






## Conjugate prior

- Prior: $\operatorname{Beta}\left(\alpha_{H}, \alpha_{T}\right)$
- Data: $m_{H}$ heads and $m_{T}$ tails (binomial likelihood)
- Posterior distribution:

$$
P(\theta \mid \mathcal{D}) \sim \operatorname{Beta}\left(m_{H}+\alpha_{H}, m_{T}+\alpha_{T}\right)
$$

- Given likelihood function $P(D \mid \theta)$
- (Parametric) prior of the form $\mathrm{P}(\theta \mid \alpha)$ is conjugate to likelihood function if posterior is of the same parametric family, and can be written as:
$\square \mathrm{P}\left(\theta \mid \alpha^{\prime}\right)$, for some new set of parameters $\alpha^{\prime}$


## Using Bayesian posterior

- Posterior distribution:


$$
P(\theta \mid \mathcal{D}) \sim \operatorname{Beta}\left(m_{H}+\alpha_{H}, m_{T}+\alpha_{T}\right)
$$

- Bayesian inference:
$\square$ No longer single parameter:

$$
E[f(\theta)]=\int_{0}^{1} f(\theta) P(\theta \mid \mathcal{D}) d \theta
$$

$\square$ Integral is often hard to compute

## Bayesian prediction of a new coin flip

- Prior:

- Observed $m_{H}$ heads, $m_{T}$ tails, what is probability of $m+1$ flip is heads?


## Asymptotic behavior and equivalent

 sample size- Beta prior equivalent to extra thumbtack flips:
$E[\theta]=\frac{m_{H}+\alpha_{H}}{m_{H}+\alpha_{H}+m_{T}+\alpha_{T}}$
- As $m \rightarrow \infty$, prior is "forgotten"
- But, for small sample size, prior is important!
- Equivalent sample size:
$\square$ Prior parameterized by $\alpha_{H}, \alpha_{T}$, or
$\square \mathrm{m}^{\prime}$ (equivalent sample size) and $\alpha$
$E[\theta]=\frac{m_{H}+\alpha m^{\prime}}{m_{H}+m_{T}+m^{\prime}}$



## Bayesian learning corresponds to

 smoothing$E[\theta]=\frac{m_{H}+\alpha m^{\prime}}{m_{H}+m_{T}+m^{\prime}}$


- $\mathrm{m}=0 \Rightarrow$ prior parameter
- $\mathrm{m} \rightarrow \infty \Rightarrow \mathrm{MLE}$


## Bayesian learning for multinomial

- What if you have a $k$ sided coin???
- Likelihood function if multinomial:
$\square$
- Conjugate prior for multinomial is Dirichlet:
$\square \theta \sim \operatorname{Dirichlet}\left(\alpha_{1}, \ldots, \alpha_{k}\right) \sim \prod_{i} \theta_{i}^{\alpha_{i}-1}$
- Observe $m$ data points, $m_{i}$ from assignment i , posterior:
- Prediction:


## Bayesian learning for two-node BN

- Parameters $\theta_{\mathrm{X}}, \theta_{\mathrm{Y} \mid \mathrm{X}}$
- Priors:
$\square \mathrm{P}\left(\theta_{\mathrm{x}}\right)$ :
$\square \mathrm{P}\left(\theta_{Y \mid X}\right):$



## Global parameter independence, d-separation and local prediction

- Independencies in meta BN:

Proposition: For fully observable data $D$, if prior satisfies global parameter independence, then
$P(\theta \mid \mathcal{D})=\prod_{i} P\left(\theta_{X_{i} \mid \mathrm{Pa}_{X_{i}}}\right.$
D)


## Within a CPT

- Meta BN including CPT parameters:
- Are $\theta_{\mathrm{Y} \mid \mathrm{X}=\mathrm{t}}$ and $\theta_{\mathrm{Y} \mid \mathrm{X}=\mathrm{f}}$ d-separated given $D$ ?
- Are $\theta_{Y \mid X=t}$ and $\theta_{Y \mid X=f}$ independent given $D$ ?
$\square$ Context-specific independence!!!
- Posterior decomposes:


## Priors for BN CPTs

(more when we talk about structure learning)

- Consider each CPT: $\mathrm{P}(\mathrm{X} \mid \mathrm{U}=\mathbf{u})$
- Conjugate prior:
$\square \operatorname{Dirichlet}\left(\alpha_{X=1 \mid U=u}, \ldots, \alpha_{X=k \mid U=u}\right)$
- More intuitive:
"prior data set" $D$ ' with $m$ ' equivalent sample size
"prior counts":
$\square$ prediction:



## What you need to know about parameter learning

- MLE:
score decomposes according to CPTs
$\square$ optimize each CPT separately
■ Bayesian parameter learning:
motivation for Bayesian approach
$\square$ Bayesian prediction
$\square$ conjugate priors, equivalent sample size
$\square$ Bayesian learning $\Rightarrow$ smoothing
- Bayesian learning for BN parameters
$\square$ Global parameter independence
$\square$ Decomposition of prediction according to CPTs
$\square$ Decomposition within a CPT

