

10-301/10-601 Introduction to Machine Learning

Machine Learning Department School of Computer Science Carnegie Mellon University

Hidden Markov Models (Part II)

Matt Gormley Lecture 19 Mar. 27, 2023

Reminders

- Practice Problems: Exam 2
 - Out: Fri, Mar. 24
- Exam 2
 - Thu, Mar. 30, 6:30pm 8:30pm
- Homework 7: Hidden Markov Models
 - Out: Fri, Mar. 31
 - Due: Mon, Apr. 10 at 11:59pm

SUPERVISED LEARNING FOR HMMS

Recipe for Closed-form MLE

1. Assume data was generated i.i.d. from some model (i.e. write the generative story) $x^{(i)} \sim p(x|\theta)$

2. Write log-likelihood

$$\ell(\boldsymbol{\theta}) = \log p(\mathbf{x}^{(1)}|\boldsymbol{\theta}) + \dots + \log p(\mathbf{x}^{(N)}|\boldsymbol{\theta})$$

3. Compute partial derivatives (i.e. gradient)

$$\frac{\partial \ell(\boldsymbol{\theta})}{\partial \theta_1} = \dots$$
$$\frac{\partial \ell(\boldsymbol{\theta})}{\partial \theta_2} = \dots$$
$$\frac{\partial \ell(\boldsymbol{\theta})}{\partial \theta_M} = \dots$$

4. Set derivatives to zero and solve for θ

$$\partial \ell(\theta)/\partial \theta_{\rm m} = {\rm o \ for \ all \ m} \in \{1, ..., M\}$$

 $\theta^{\rm MLE} = {\rm solution \ to \ system \ of \ } M \ {\rm equations \ and \ } M \ {\rm variables}$

5. Compute the second derivative and check that $\ell(\theta)$ is concave down at θ^{MLE}

MLE of Categorical Distribution

1. Suppose we have a **dataset** obtained by repeatedly rolling a M-sided (weighted) die N times. That is, we have data

$$\mathcal{D} = \{x^{(i)}\}_{i=1}^{N}$$

where $x^{(i)} \in \{1, \dots, M\}$ and $x^{(i)} \sim \mathsf{Categorical}(\phi)$.

2. A random variable is **Categorical** written $X \sim \mathsf{Categorical}(\phi)$ iff

$$P(X=x) = p(x; \phi) = \phi_x$$

where $x \in \{1, \dots, M\}$ and $\sum_{m=1}^{M} \phi_m = 1$. The **log-likelihood** of the data becomes:

$$\ell(oldsymbol{\phi}) = \sum_{i=1}^N \log \phi_{x^{(i)}}$$
 s.t. $\sum_{m=1}^M \phi_m = 1$

3. Solving this constrained optimization problem yields the maximum likelihood estimator (MLE):

$$\phi_m^{MLE} = \frac{N_{x=m}}{N} = \frac{\sum_{i=1}^{N} \mathbb{I}(x^{(i)} = m)}{N}$$

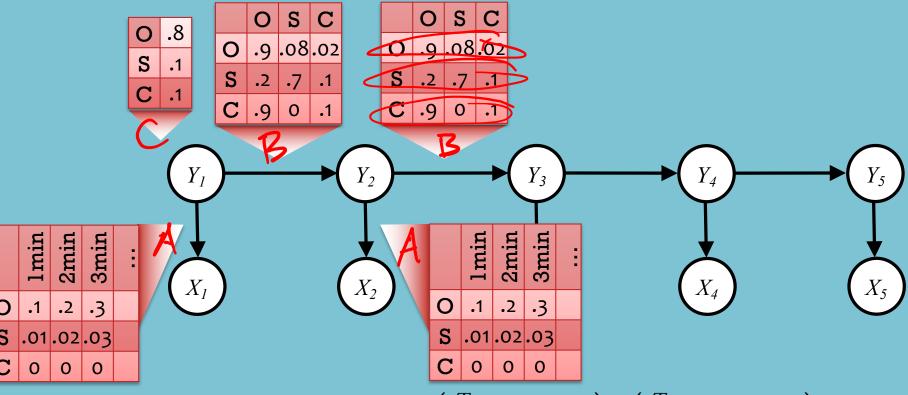


Hidden Markov Model (v1)

HMM Parameters:

Emission matrix, **A**, where $P(X_t = k | Y_t = j) = A_{j,k}, \forall t, k$ Transition matrix, **B**, where $P(Y_t = k | Y_{t-1} = j) = B_{j,k}, \forall t, k$

Initial probs, C, where $P(Y_1 = k) = C_k, \forall k$



$$P(\mathbf{X}, \mathbf{Y}) = P(Y_1) \left(\prod_{t=1}^{T} P(X_t | Y_t) \right) \left(\prod_{t=2}^{T} p(Y_t | Y_{t-1}) \right)$$

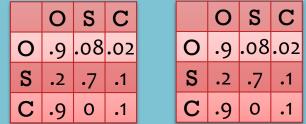
Hidden Markov Model (v1)

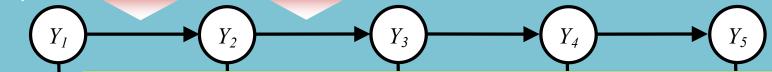
HMM Parameters:

Emission matrix, A, where $P(X_t = k | Y_t = j) = A_{i,k}, \forall t, k$ Transition matrix, **B**, where $P(Y_t = k | Y_{t-1} = j) = B_{j,k}, \forall t, k$

Initial probs, **C**, where $P(Y_1 = k) = C_k, \forall k$







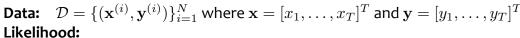
2min 3min .01.02.03 Joint Distribution (probability mass function):

$$\sum_{X_l} p(\mathbf{x}, \mathbf{y}) = p(y_1, C) \left(\prod_{t=1}^T p(x_t \mid y_t, A) \right) \left(\prod_{t=2}^T p(y_t \mid y_{t-1}, B) \right)$$

$$= C_{y_1} \left(\prod_{t=1}^{T} A_{y_t, x_t} \right) \left(\prod_{t=2}^{T} B_{y_{t-1}, y_t} \right)$$

Supervised Learning for HMM (v1)

Learning an HMM decomposes into solving two (independent) Mixture Models



$$\ell(\mathbf{A}, \mathbf{B}, \mathbf{C}) = \sum_{i=1}^{N} \log p(\mathbf{x}^{(i)}, \mathbf{y}^{(i)} \mid \mathbf{A}, \mathbf{B}, \mathbf{C})$$

$$= \sum_{i=1}^{N} \left[\underbrace{\log p(y_1^{(i)} \mid \mathbf{C})}_{\text{initial}} + \underbrace{\left(\sum_{t=2}^{T} \log p(y_t^{(i)} \mid y_{t-1}^{(i)}, \mathbf{B})\right)}_{\text{transition}} + \underbrace{\left(\sum_{t=1}^{T} \log p(x_t^{(i)} \mid y_t^{(i)}, \mathbf{A})\right)}_{\text{emission}} \right]$$

MLE:

$$\hat{\mathbf{A}}, \hat{\mathbf{B}}, \hat{\mathbf{C}} = \underset{\mathbf{A}, \mathbf{B}, \mathbf{C}}{\operatorname{argmax}} \ell(\mathbf{A}, \mathbf{B}, \mathbf{C})$$

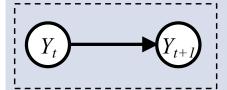
$$\Rightarrow \hat{\mathbf{C}} = \underset{\mathbf{C}}{\operatorname{argmax}} \sum_{i=1}^{N} \log p(y_1^{(i)} \mid \mathbf{C})$$

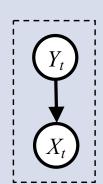
$$\hat{\mathbf{B}} = \underset{\mathbf{B}}{\operatorname{argmax}} \sum_{i=1}^{N} \sum_{t=2}^{T} \log p(y_t^{(i)} \mid y_{t-1}^{(i)}, \mathbf{B})$$

$$\hat{\mathbf{A}} = \underset{\mathbf{A}}{\operatorname{argmax}} \sum_{i=1}^{N} \sum_{t=1}^{T} \log p(x_t^{(i)} \mid y_t^{(i)}, \mathbf{A})$$

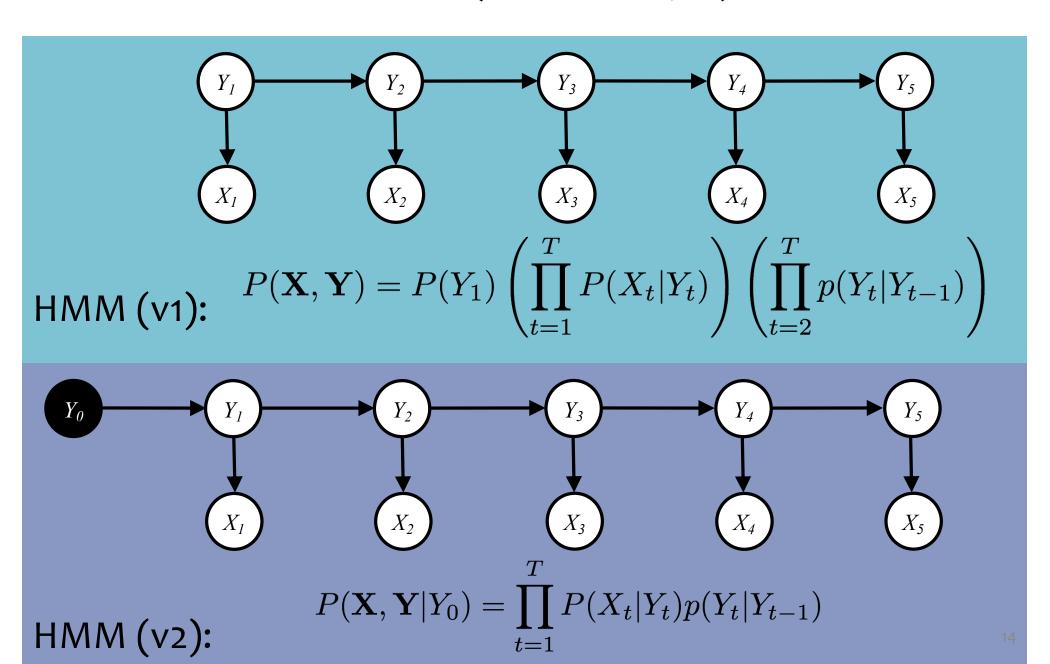
We can solve the above in closed form, which yields...

$$\hat{C}_k = rac{\#(y_1^{(i)} = k)}{N}, \, orall k$$
 $\hat{B}_{j,k} = rac{\#(y_t^{(i)} = k ext{ and } y_{t-1}^{(i)} = j)}{\#(y_{t-1}^{(i)} = j)}, \, orall j, k$
 $\hat{A}_{j,k} = rac{\#(x_t^{(i)} = k ext{ and } y_t^{(i)} = j)}{\#(y_t^{(i)} = j)}, \, orall j, k$

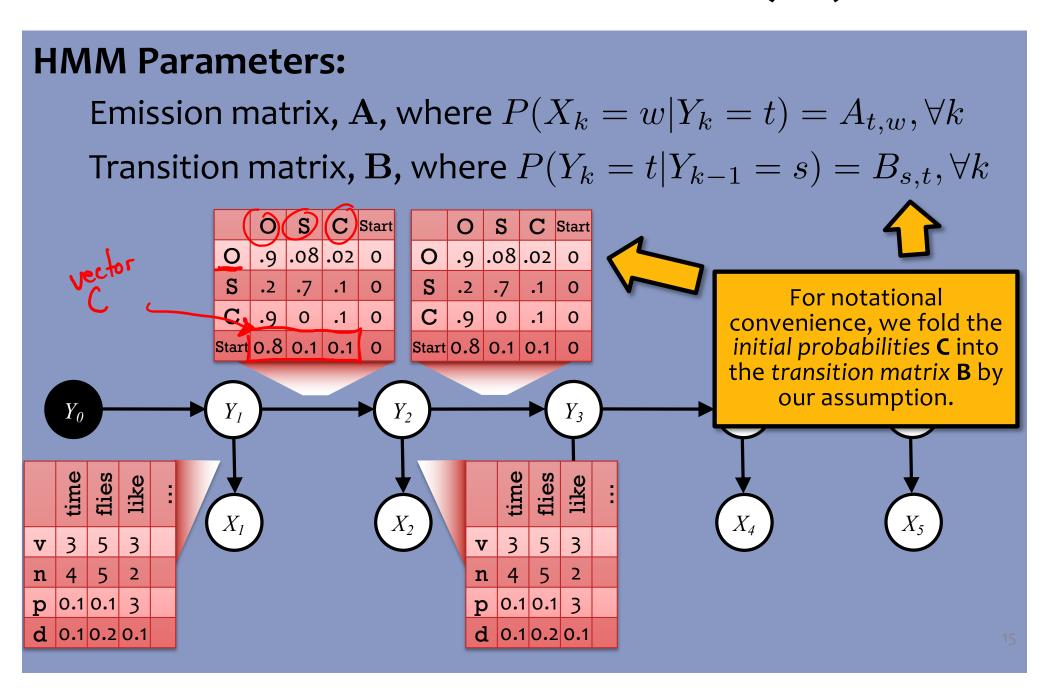




HMM (two ways)



Hidden Markov Model (v2)



Hidden Markov Model (v2)

HMM Parameters:

Emission matrix, **A**, where $P(X_t = k | Y_t = j) = A_{j,k}, \forall t, k$ Transition matrix, **B**, where $P(Y_t = k | Y_{t-1} = j) = B_{j,k}, \forall t, k$

Assumption: $y_0 = START$

Generative Story:

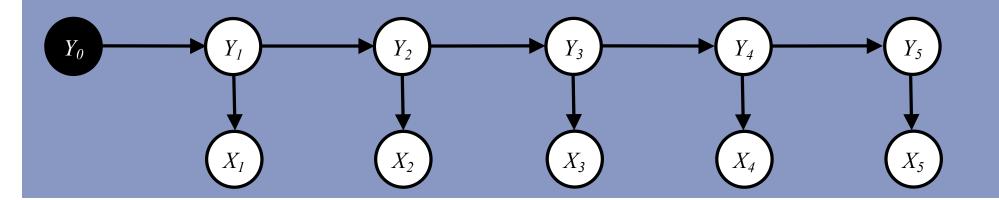
 $Y_t \sim \text{Multinomial}(\mathbf{B}_{Y_{t-1}}) \ \forall t$

 $X_t \sim \mathsf{Multinomial}(\mathbf{A}_{Y_t}) \ \forall t$

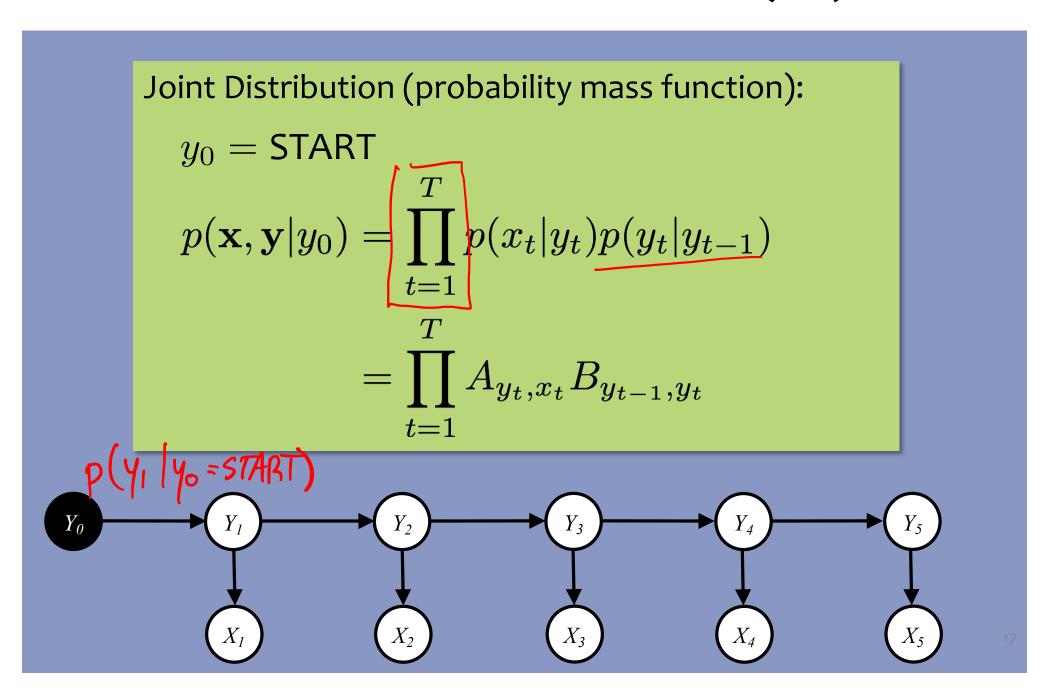




For notational convenience, we fold the initial probabilities **C** into the transition matrix **B** by our assumption.



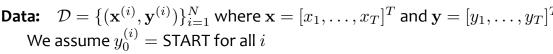
Hidden Markov Model (v2)



Supervised Learning for HMM (v2) Data: $\mathcal{D} = \{(\mathbf{x}^{(i)}, \mathbf{y}^{(i)})\}_{i=1}^{N}$ where $\mathbf{x} = [x_1, \dots, x_T]^T$ and $\mathbf{y} = [y_1, \dots, y_T]^T$

Learning an **HMM** decomposes into solving two (independent) Mixture Models

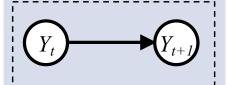


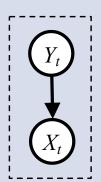


Likelihood:

$$\ell(\mathbf{A}, \mathbf{B}) = \sum_{i=1}^{N} \log p(\mathbf{x}^{(i)}, \mathbf{y}^{(i)} \mid \mathbf{A}, \mathbf{B})$$

$$= \sum_{i=1}^{N} \left[\sum_{t=1}^{T} \underbrace{\log p(y_t^{(i)} \mid y_{t-1}^{(i)}, \mathbf{B})}_{\text{transition}} + \underbrace{\log p(x_t^{(i)} \mid y_t^{(i)}, \mathbf{A})}_{\text{emission}} \right]$$





MLE:

$$\hat{\mathbf{A}}, \hat{\mathbf{B}} = \underset{\mathbf{A}, \mathbf{B}, \mathbf{C}}{\operatorname{argmax}} \ell(\mathbf{A}, \mathbf{B})$$

$$\Rightarrow \hat{\mathbf{B}} = \underset{\mathbf{B}}{\operatorname{argmax}} \sum_{i=1}^{N} \sum_{t=1}^{T} \log p(y_t^{(i)} \mid y_{t-1}^{(i)}, \mathbf{B})$$

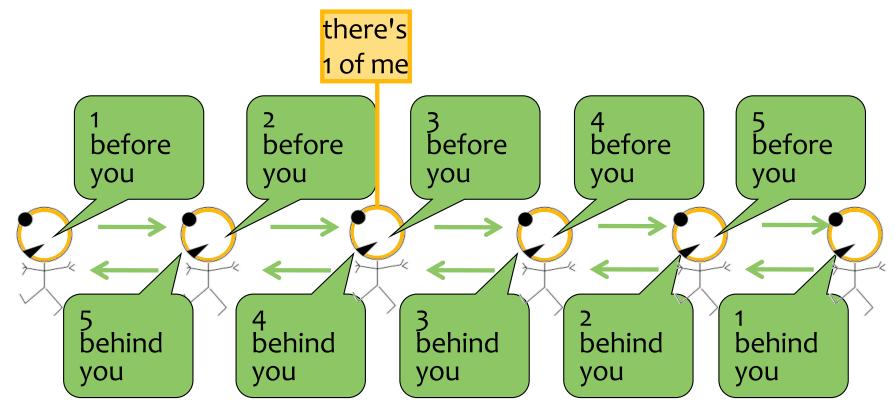
$$\hat{\mathbf{A}} = \underset{\mathbf{A}}{\operatorname{argmax}} \sum_{i=1}^{N} \sum_{t=1}^{T} \log p(x_t^{(i)} \mid y_t^{(i)}, \mathbf{A})$$

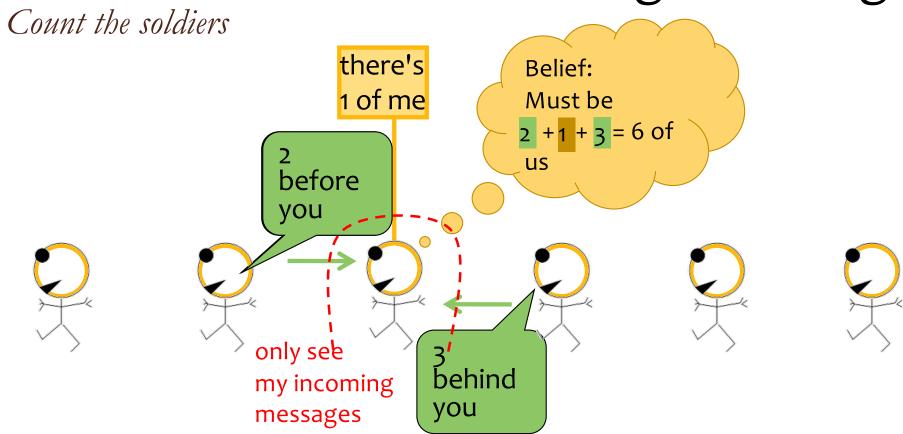
We can solve the above in closed form, which yields...

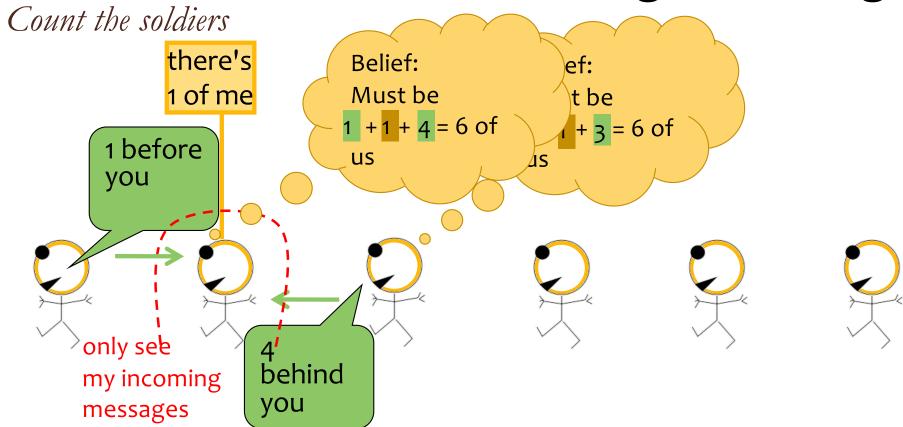
$$\hat{B}_{j,k} = rac{\#(oldsymbol{y}_t^{(i)} = k ext{ and } oldsymbol{y}_{t-1}^{(i)} = j)}{\#(oldsymbol{y}_{t-1}^{(i)} = j)}, \, orall j, k$$
 $\hat{A}_{j,k} = rac{\#(oldsymbol{x}_t^{(i)} = k ext{ and } oldsymbol{y}_t^{(i)} = j)}{\#(oldsymbol{y}_t^{(i)} = j)}, \, orall j, k$

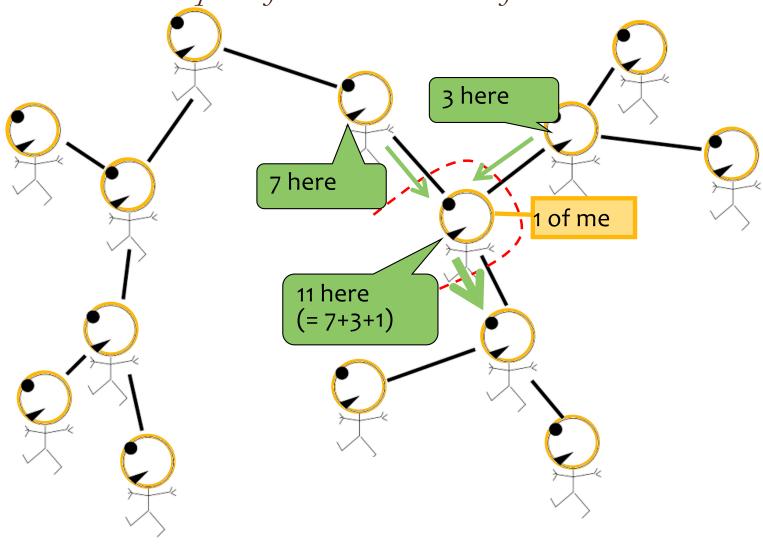
BACKGROUND: MESSAGE PASSING

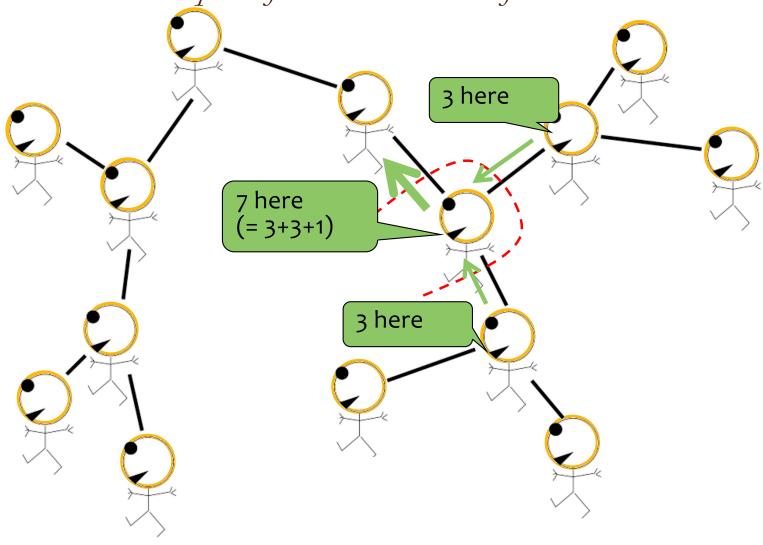
Count the soldiers

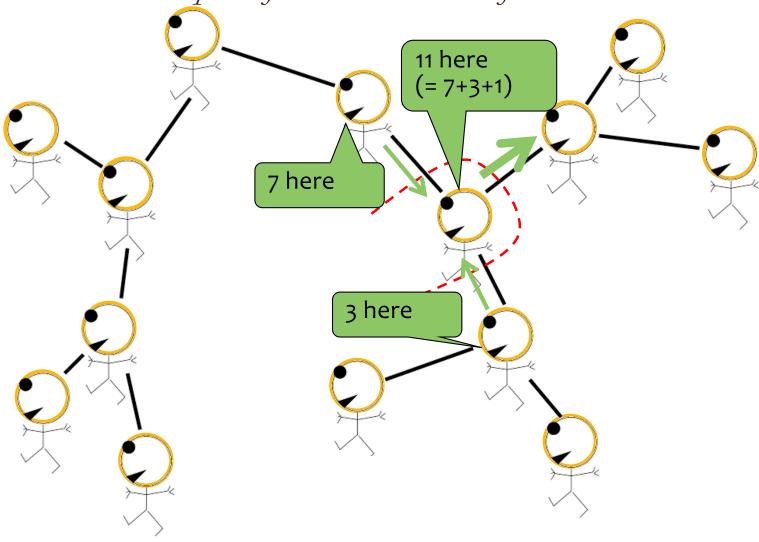


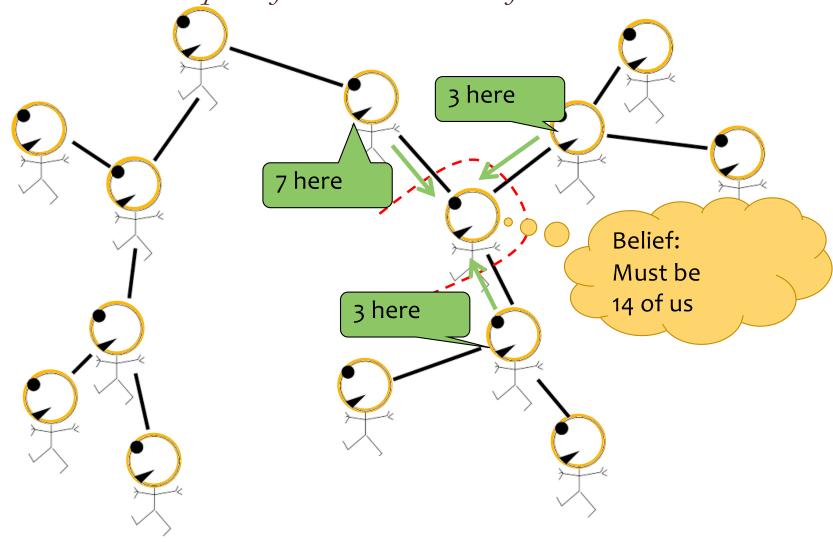


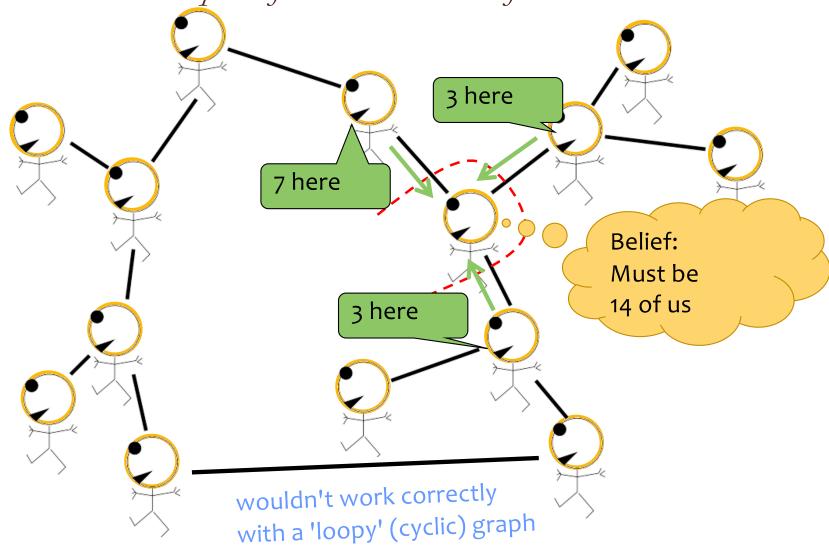












INFERENCE FOR HMMS



True or False: The joint probability of the observations and the hidden states in an HMM is given by:

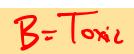
$$P(\mathbf{X} = \mathbf{x}, \mathbf{Y} = \mathbf{y}) = C_{y_1} \left[\prod_{t=1}^{T} A_{y_t, x_t} \right] \left[\prod_{t=1}^{T-1} B_{y_t, y_{t+1}} \right]$$

Recall:

Emission matrix, **A**, where $P(X_t = k | Y_t = j) = A_{j,k}, \forall t, k$ Transition matrix, **B**, where $P(Y_t = k | Y_{t-1} = j) = B_{j,k}, \forall t, k$ Initial probs, C, where $P(Y_1 = k) = C_k, \forall k$

Inference



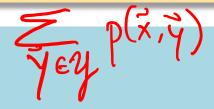


Question: QZ A=Tre B=Toxic C= False

True or False: The probability of the observations in an HMM is given by:

$$P(\mathbf{X} = \mathbf{x}) = \prod_{t=1}^{T} A_{x_t, x_{t-1}}$$

Recall:



Emission matrix, **A**, where $P(X_t = k | Y_t = j) = A_{i,k}, \forall t, k$ Transition matrix, **B**, where $P(Y_t = k | Y_{t-1} = j) = B_{j,k}, \forall t, k$ Initial probs, C, where $P(Y_1 = k) = C_k, \forall k$

Inference for HMMs

Whiteboard

- Three Inference Problems for an HMM
 - Evaluation: Compute the probability of a given sequence of observations
 - 2. Viterbi Decoding: Find the most-likely sequence of hidden states, given a sequence of observations
 - 3. Marginals: Compute the marginal distribution for a hidden state, given a sequence of observations

THE SEARCH SPACE FOR FORWARD-BACKWARD

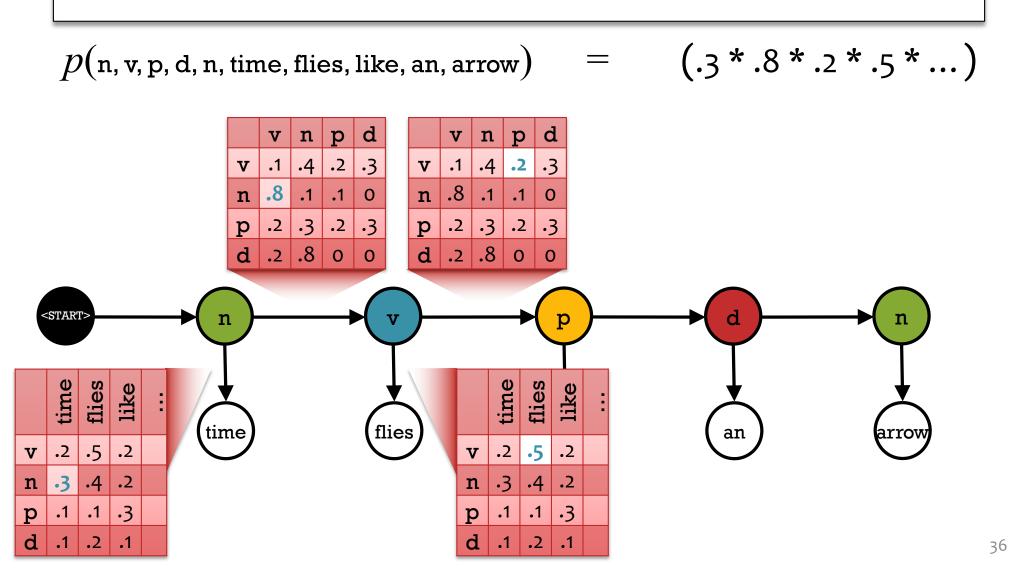
Dataset for Supervised Part-of-Speech (POS) Tagging

Data: $\mathcal{D} = \{oldsymbol{x}^{(n)}, oldsymbol{y}^{(n)}\}_{n=1}^N$

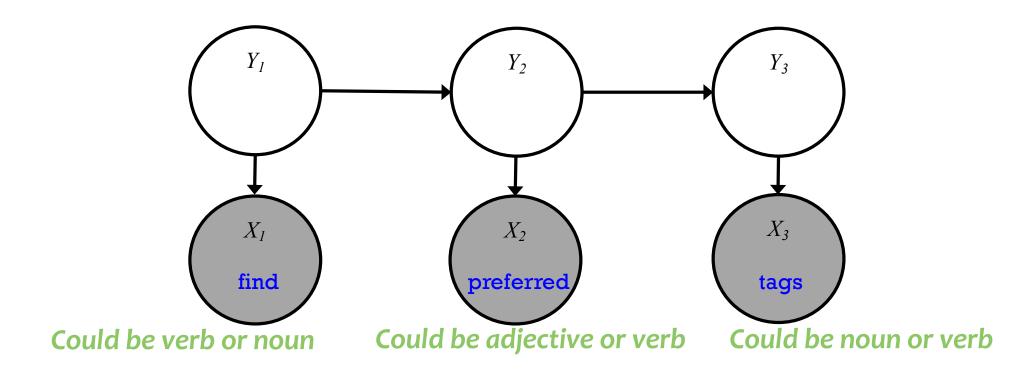
Sample 1:	n	flies	p like	an	$\begin{array}{c c} & & \\ & &$
Sample 2:	n	n	like	d	$y^{(2)}$ $x^{(2)}$
Sample 3:	n	fly	with	n	$\begin{cases} n \\ y^{(3)} \\ x^{(3)} \end{cases}$
Sample 4:	with	time	you	will	$\begin{cases} \mathbf{v} \\ \mathbf{see} \end{cases} y^{(4)}$

Example: HMM for POS Tagging

A Hidden Markov Model (HMM) provides a joint distribution over the the sentence/tags with an assumption of dependence between adjacent tags.



Example: HMM for POS Tagging



Inference for HMMs

Whiteboard

- Brute Force Evaluation
- Forward-backward search space

THE FORWARD-BACKWARD ALGORITHM

How is efficient computation even possible?

- The short answer is dynamic programming!
- The key idea is this:
 - We first come up with a recursive definition for the quantity we want to compute
 - We then observe that many of the recursive intermediate terms are reused across timesteps and tags
 - We then perform bottom-up dynamic programming by running the recursion in reverse, storing the intermediate quantities along the way!
- This enables us to search the exponentially large space in polynomial time!

Inference for HMMs

Whiteboard

Forward-backward algorithm (edge weights version)

Forward-Backward Algorithm

Definitions

$$\alpha_t(k) \triangleq p(x_1, \dots, x_t, y_t = k)$$

$$\beta_t(k) \triangleq p(x_{t+1}, \dots, x_T \mid y_t = k)$$

Assume

$$y_0 = \mathsf{START}$$

$$y_{T+1} = END$$

1. Initialize

$$lpha_0({\sf START}) = 1$$
 $lpha_0(k) = 0, \, \forall k
eq {\sf START}$ $eta_{T+1}({\sf END}) = 1$ $eta_{T+1}(k) = 0, \, \forall k
eq {\sf END}$

2. Forward Algorithm

for
$$t = 1, ..., T+1$$
:
for $k = 1, ..., K$:

$$\alpha_t(k) = \sum_{j=1}^K p(x_t \mid y_t = k) \alpha_{t-1}(j) p(y_t = k \mid y_{t-1} = j)$$

3. Backward Algorithm

for
$$t = T, ..., 0$$
:
for $k = 1, ..., K$:

$$\beta_t(k) = \sum_{j=1}^K p(x_{t+1} \mid y_{t+1} = j) \beta_{t+1}(j) p(y_{t+1} = j \mid y_t = k)$$

- 4. Evaluation $p(\mathbf{x}) = \alpha_{T+1}(\mathsf{END})$
- 5. Marginals $p(y_t = k \mid \mathbf{x}) = \frac{\alpha_t(k)\beta_t(k)}{p(\mathbf{x})}$

Forward-Backward Algorithm

1. Initialize

$$lpha_0({\sf START}) = 1$$
 $lpha_0(k) = 0, \, \forall k
eq {\sf START}$ $eta_{T+1}({\sf END}) = 1$ $eta_{T+1}(k) = 0, \, \forall k
eq {\sf END}$

Definitions

$$\alpha_t(k) \triangleq p(x_1, \dots, x_t, y_t = k)$$

$$\beta_t(k) \triangleq p(x_{t+1}, \dots, x_T \mid y_t = k)$$

2. Forward Algorithm

for
$$t = 1, ..., T + 1$$
:

for $k = 1, ..., K$:
$$\alpha_t(k) = \sum_{j=1}^K p(x_t \mid y_t = k) \alpha_{t-1}(j) p(y_t = k \mid y_{t-1} = j)$$

Assume

$$y_0 = \mathsf{START}$$

$$y_{T+1} = END$$

 $O(K^2T)$

O(K) kward Algorithm

for
$$t = T, \dots, 0$$
:

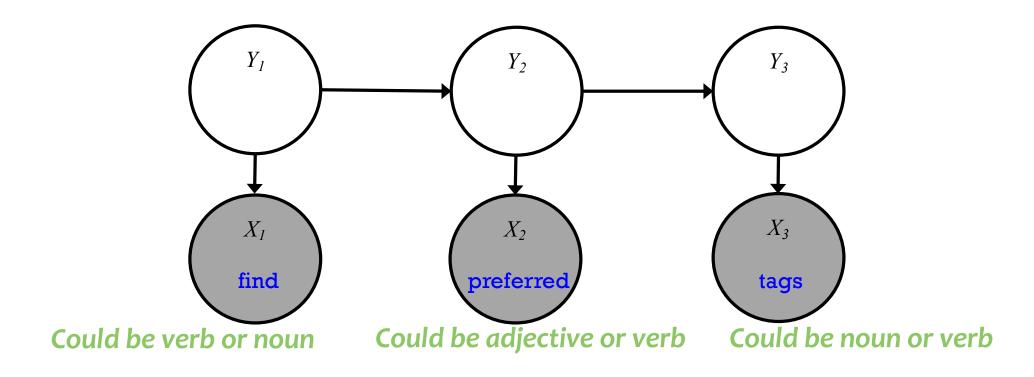
Brute force $O(K^T)$

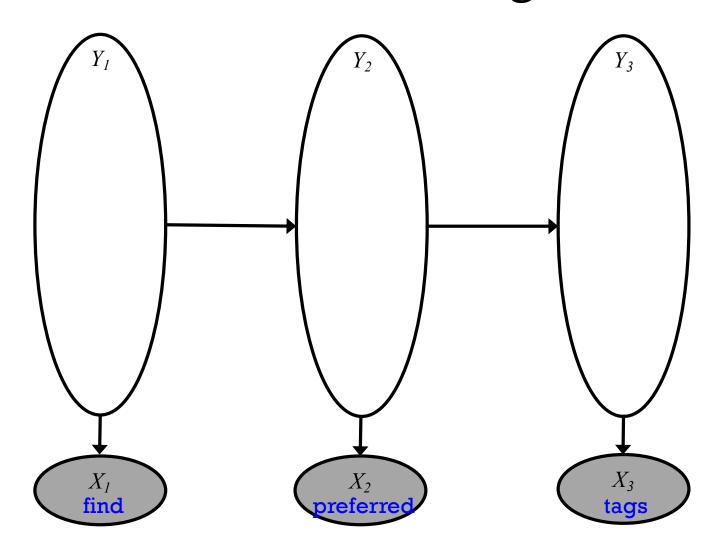
or
$$k=1,\ldots,K$$
:

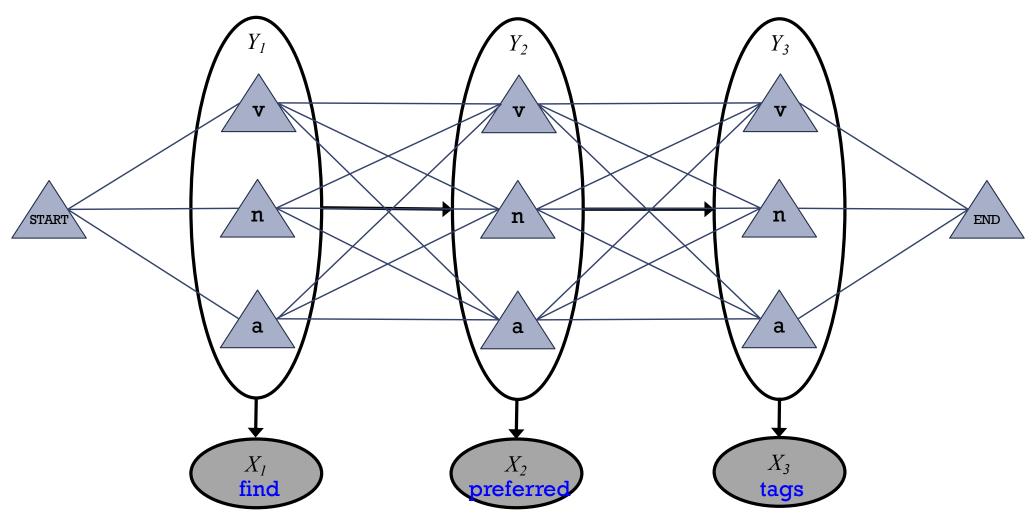
algorithm would be
$$O(\mathsf{K}^\mathsf{T}) \beta_t(k) = \sum_{j=1}^K p(x_{t+1} \mid y_{t+1} = j) \beta_{t+1}(j) p(y_{t+1} = j \mid y_t = k)$$

- 4. Evaluation $p(\mathbf{x}) = \alpha_{T+1}(\mathsf{END})$
- 5. Marginals $p(y_t = k \mid \mathbf{x}) = \frac{\alpha_t(k)\beta_t(k)}{p(\mathbf{x})}$

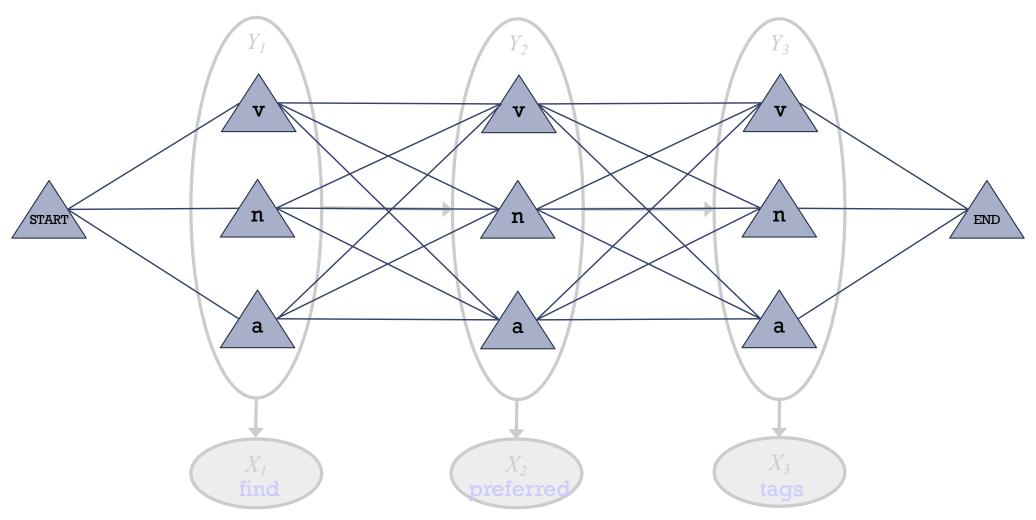
EXAMPLE: FORWARD-BACKWARD ON THREE WORDS



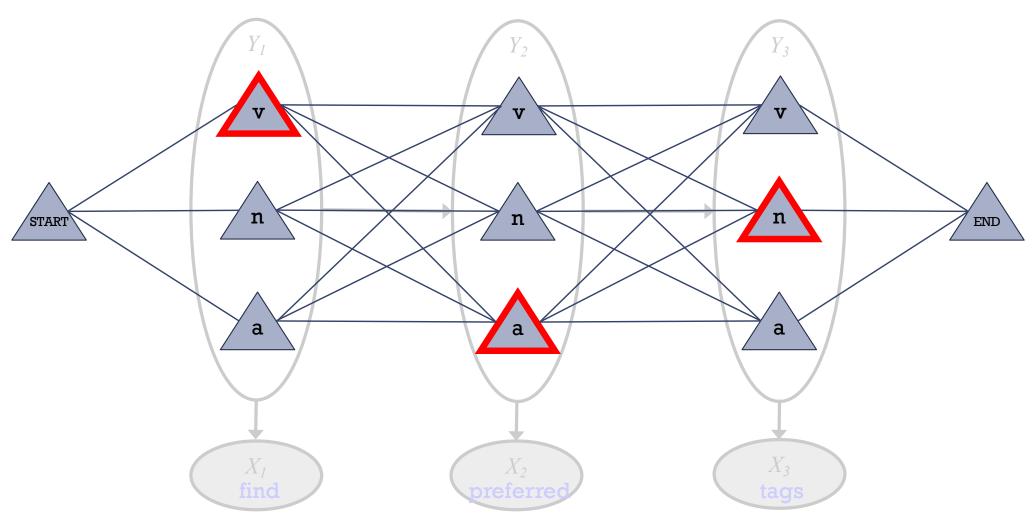




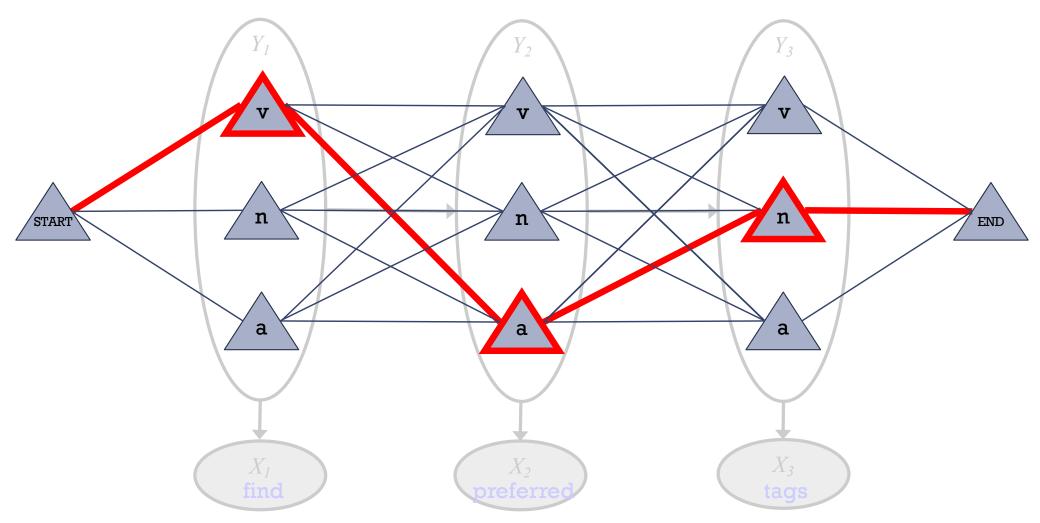
• Let's show the possible values for each variable



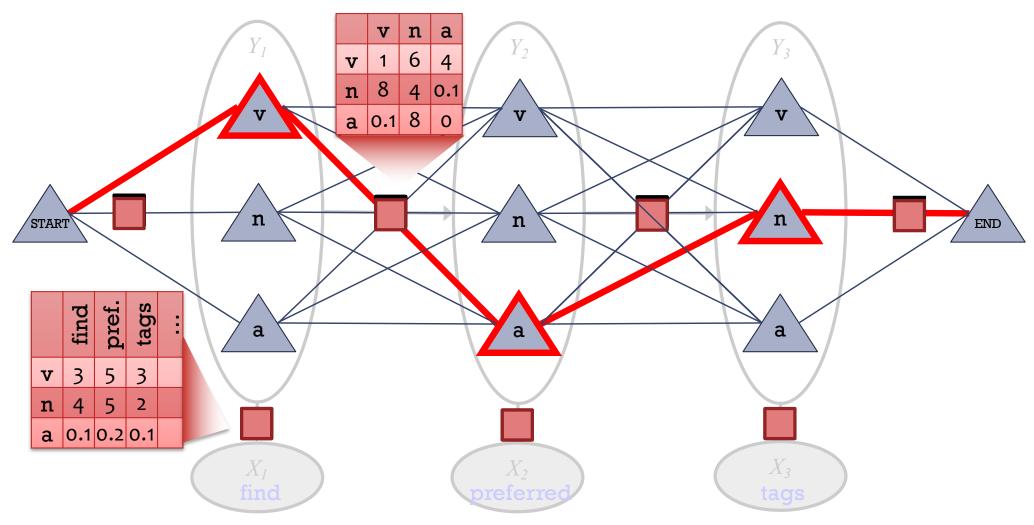
Let's show the possible values for each variable



- Let's show the possible values for each variable
- One possible assignment



- Let's show the possible values for each variable
- One possible assignment
- And what the 7 transition / emission factors think of it ...

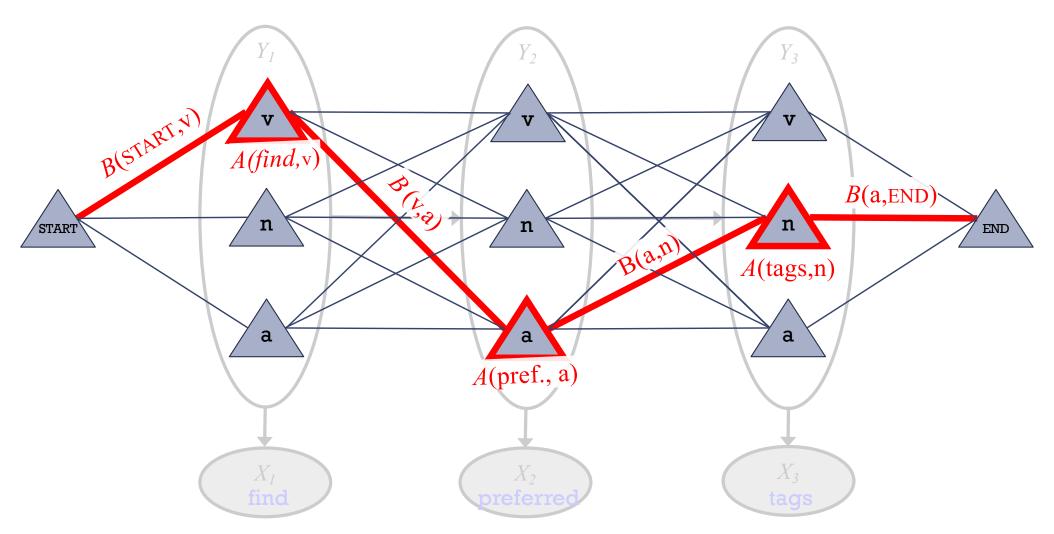


- Let's show the possible values for each variable
- One possible assignment
- And what the 7 transition / emission factors think of it ...

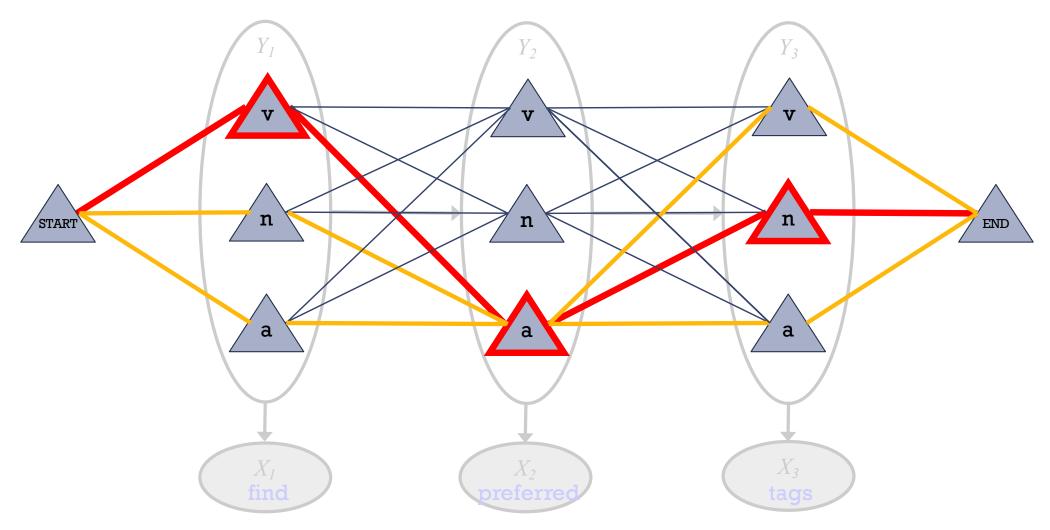
Viterbi Algorithm: Most Probable Assignment B(START,V) A(find,v)B(a,END)END B(a,n)A(tags,n)A(pref., a)

- So $p(\mathbf{v} \mathbf{a} \mathbf{n}) = (1/Z) * product of 7 numbers$
- Numbers associated with edges and nodes of path
- Most probable assignment = path with highest product

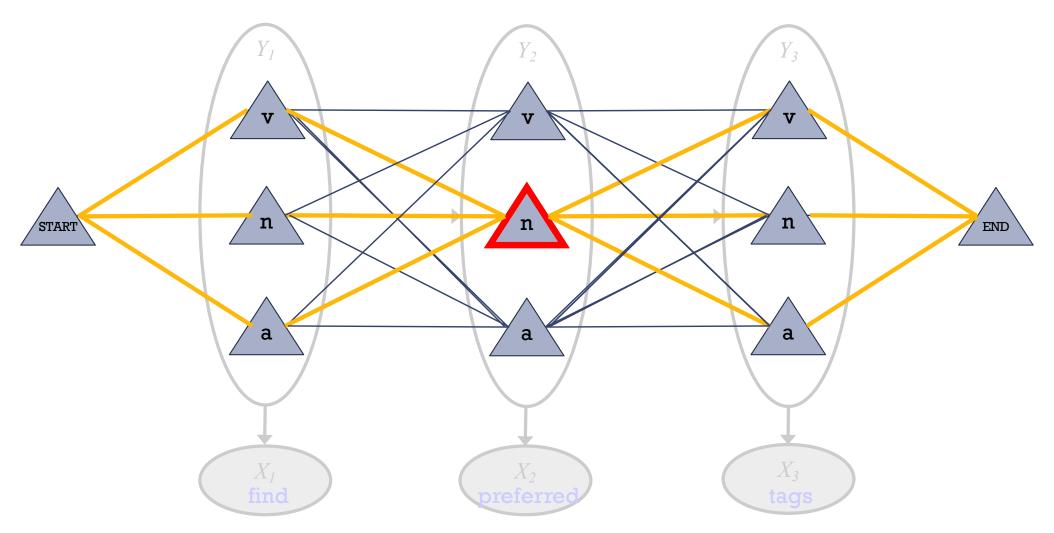
Viterbi Algorithm: Most Probable Assignment



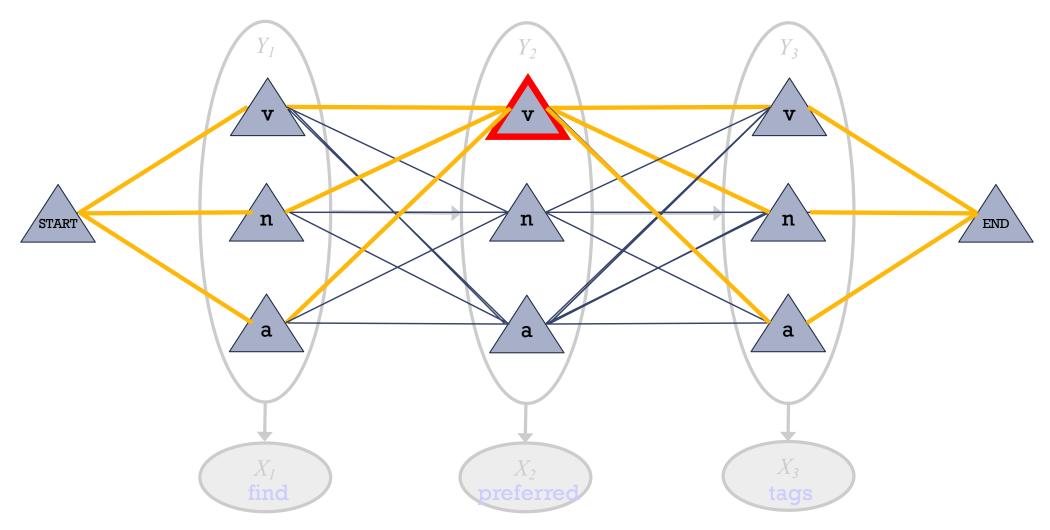
• So $p(\mathbf{v} \mathbf{a} \mathbf{n}) = (1/Z) * product weight of one path$



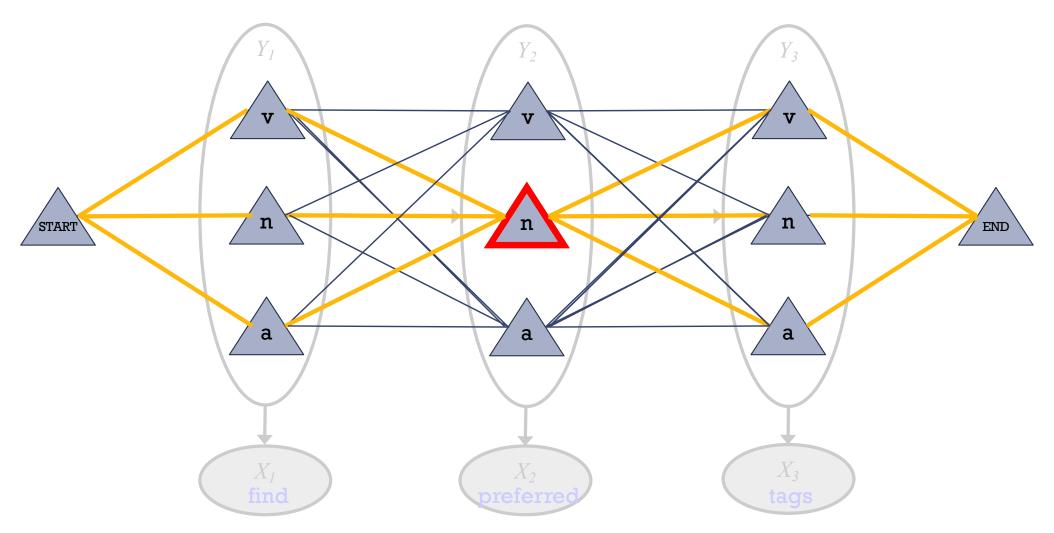
- So $p(\mathbf{v} \mathbf{a} \mathbf{n}) = (1/Z) * product weight of one path$
- Marginal probability $p(Y_2 = a) = (1/Z) * total weight of all paths through a$



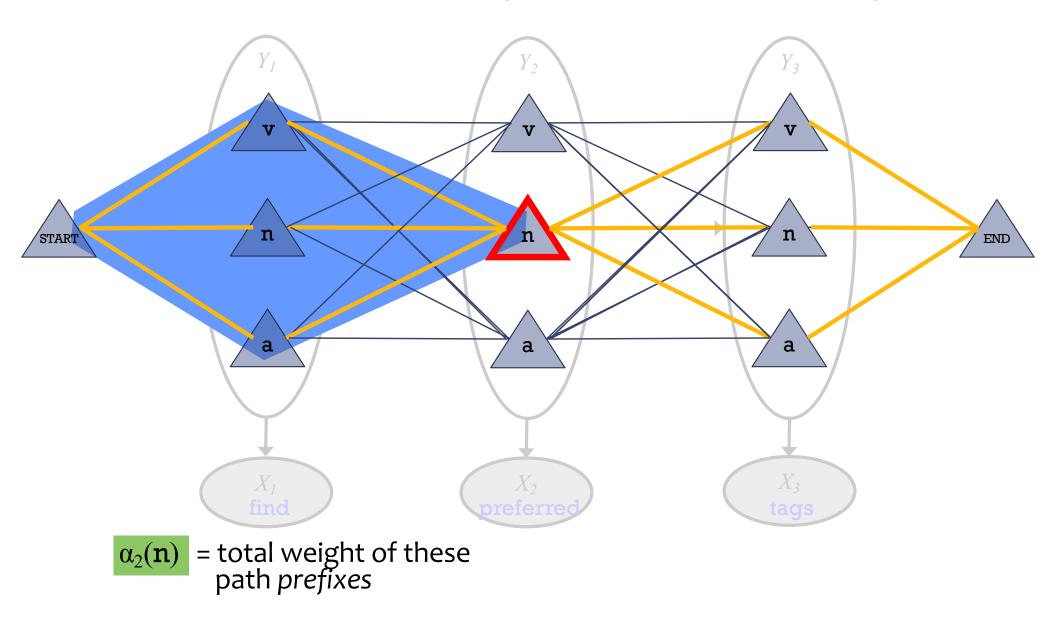
- So $p(\mathbf{v} \mathbf{a} \mathbf{n}) = (1/Z) * product weight of one path$
- Marginal probability $p(Y_2 = n)$ = (1/Z) * total weight of all paths through n



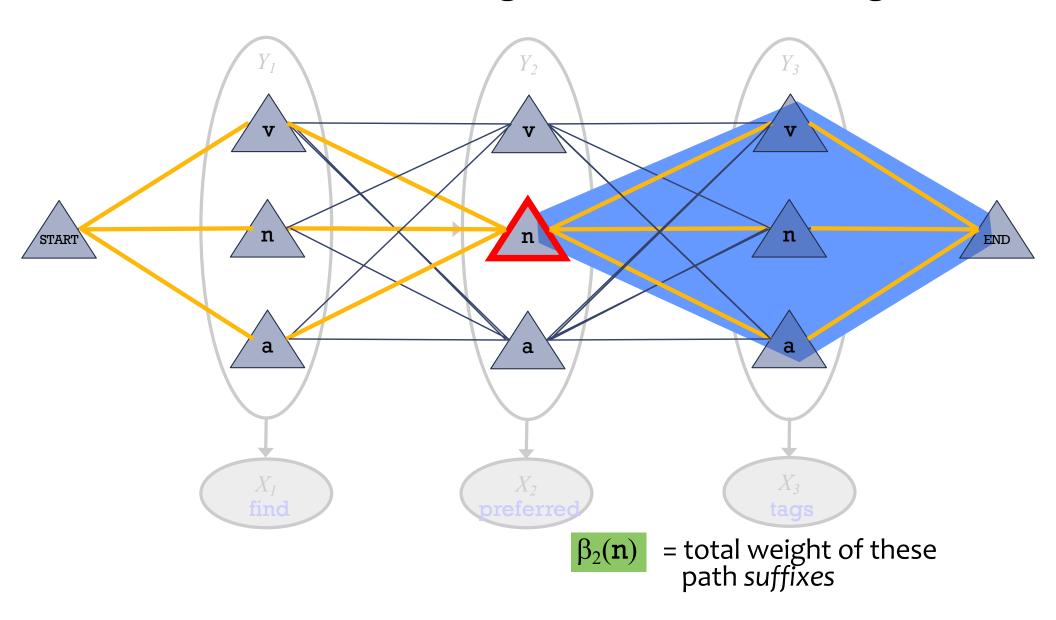
- So $p(\mathbf{v} \mathbf{a} \mathbf{n}) = (1/Z) * product weight of one path$
- Marginal probability $p(Y_2 = v)$ = (1/Z) * total weight of all paths through

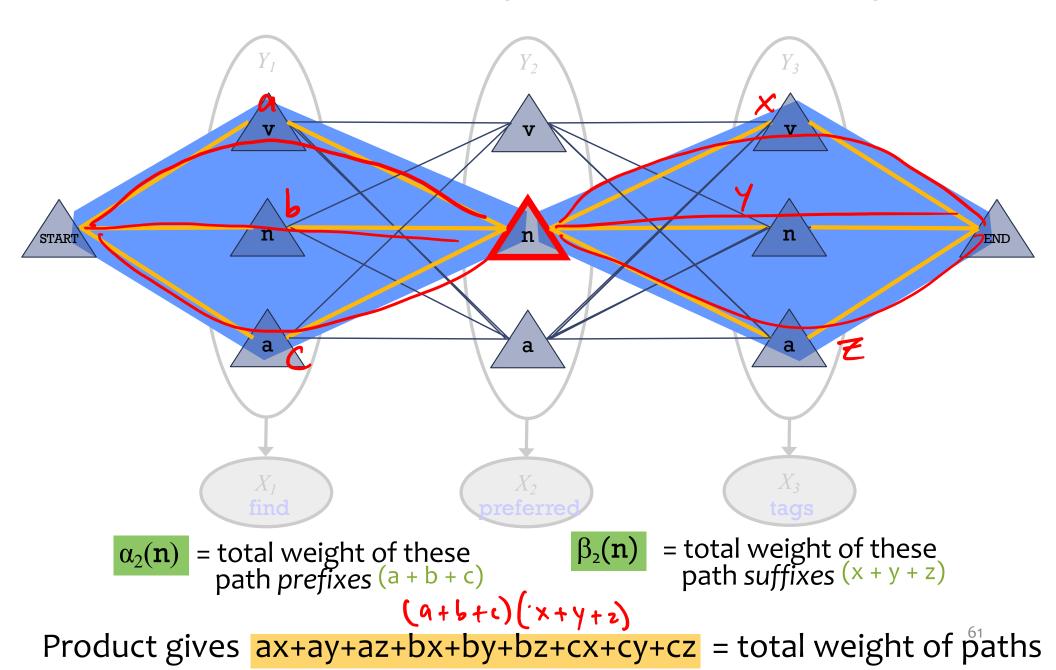


- So $p(\mathbf{v} \mathbf{a} \mathbf{n}) = (1/Z) * product weight of one path$
- Marginal probability $p(Y_2 = n)$ = (1/Z) * total weight of all paths through n



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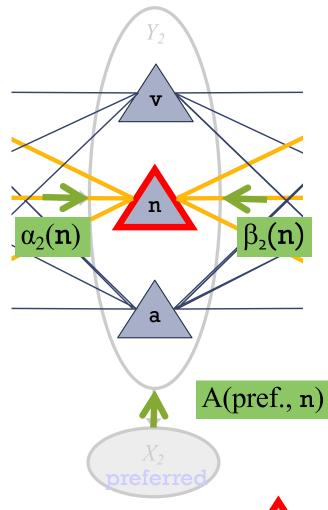




Oops! The weight of a path through a state also includes a weight at that state.

So $\alpha(\mathbf{n}) \cdot \beta(\mathbf{n})$ isn't enough.

The extra weight is the opinion of the emission probability at this variable.

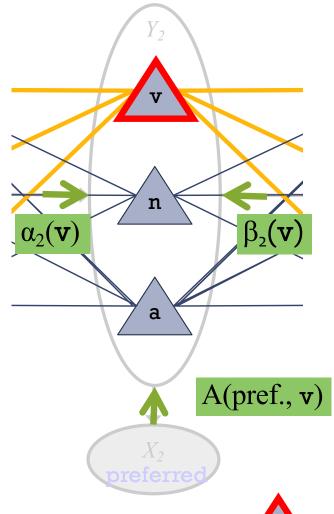


"belief that $Y_2 = \mathbf{n}$ "

total weight of all paths through



 $\alpha_2(\mathbf{n})$ A(pref., \mathbf{n}) $\beta_2(\mathbf{n})$

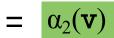


"belief that $Y_2 = \mathbf{v}$ "

"belief that $Y_2 = \mathbf{n}$ "

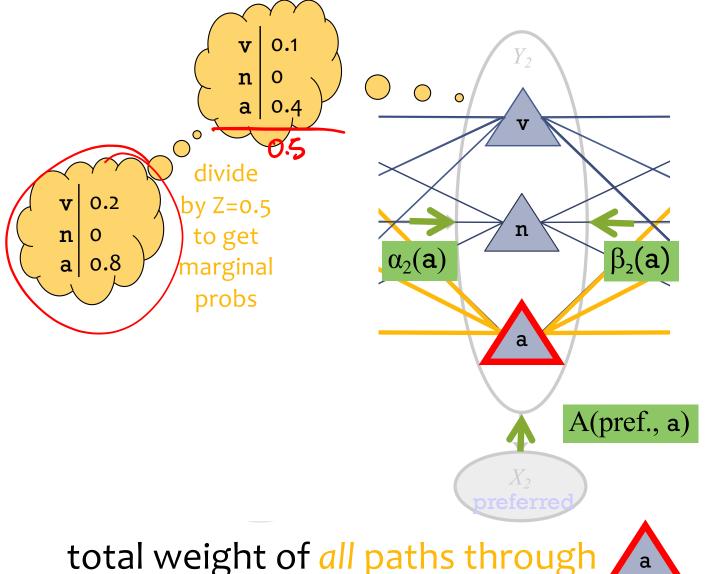
total weight of all paths through





 $\alpha_2(\mathbf{v})$ A(pref., \mathbf{v}) $\beta_2(\mathbf{v})$





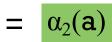
"belief that $Y_2 = \mathbf{v}$ "

"belief that $Y_2 = \mathbf{n}$ "

"belief that $Y_2 = \mathbf{a}$ "

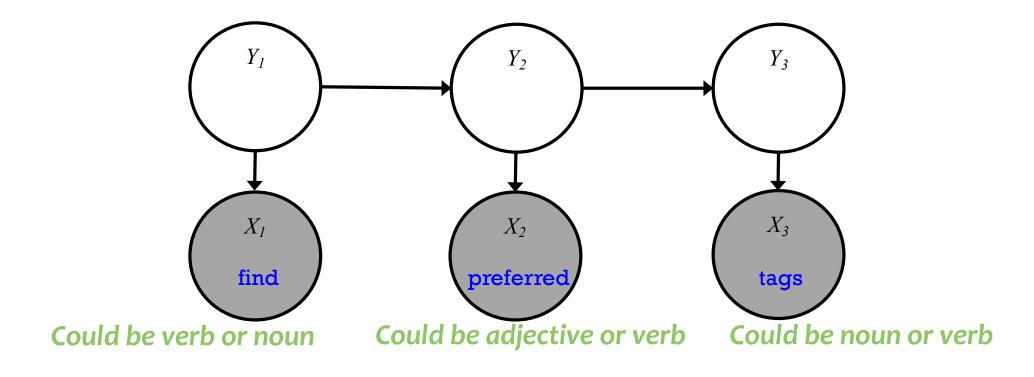
sum = Z(total weight of all paths)

total weight of all paths through



 $\alpha_2(a)$ A(pref., a) $\beta_2(a)$





THE FORWARD-BACKWARD ALGORITHM

Definitions

$$\alpha_t(k) \triangleq p(x_1, \dots, x_t, y_t = k)$$

$$\beta_t(k) \triangleq p(x_{t+1}, \dots, x_T \mid y_t = k)$$

Assume

$$y_0 = \mathsf{START}$$

$$y_{T+1} = \mathsf{END}$$

1. Initialize

$$lpha_0({\sf START}) = 1$$
 $lpha_0(k) = 0, \, \forall k
eq {\sf START}$ $eta_{T+1}({\sf END}) = 1$ $eta_{T+1}(k) = 0, \, \forall k
eq {\sf END}$

2. Forward Algorithm

for
$$t=1,\ldots,T+1$$
:
for $k=1,\ldots,K$:

$$\alpha_t(k)=\sum_{j=1}^K p(x_t\mid y_t=k)\alpha_{t-1}(j)p(y_t=k\mid y_{t-1}=j)$$

3. Backward Algorithm

for
$$t = T, ..., 0$$
:
for $k = 1, ..., K$:

$$\beta_t(k) = \sum_{j=1}^K p(x_{t+1} \mid y_{t+1} = j) \beta_{t+1}(j) p(y_{t+1} = j \mid y_t = k)$$

- 4. Evaluation $p(\mathbf{x}) = \alpha_{T+1}(\mathsf{END})$
- 5. Marginals $p(y_t = k \mid \mathbf{x}) = \frac{\alpha_t(k)\beta_t(k)}{p(\mathbf{x})}$