



Probabilistic Graphical Models

Conditional Random Fields

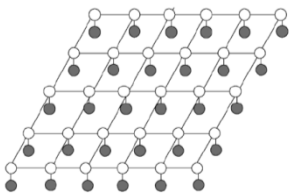
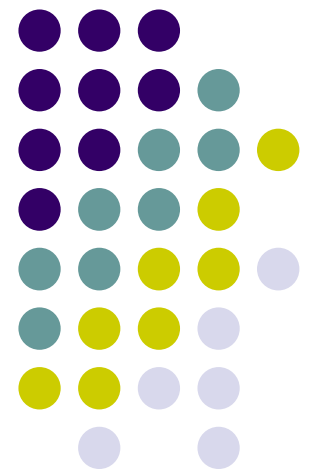
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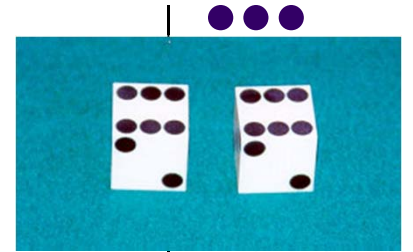
Case study I: image segmentation

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Lecture 11, February 21, 2014

Reading: See class website





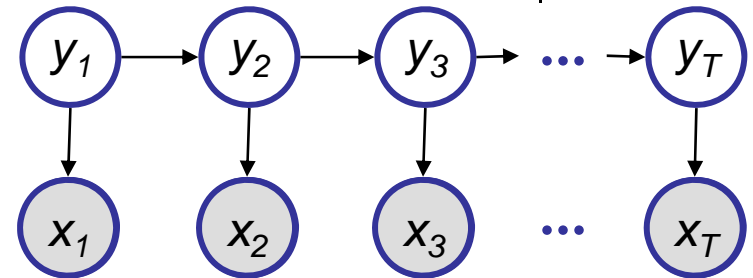
Hidden Markov Model revisit

- Transition probabilities between any two states

$$p(y_t^j = 1 | y_{t-1}^i = 1) = a_{i,j},$$

or

$$p(y_t | y_{t-1}^i = \mathbf{1}) \sim \text{Multinomial}(a_{i,1}, a_{i,2}, \dots, a_{i,M}), \forall i \in I.$$



- Start probabilities

$$p(y_1) \sim \text{Multinomial}(\pi_1, \pi_2, \dots, \pi_M).$$

- Emission probabilities associated with each state

$$p(x_t | y_t^i = \mathbf{1}) \sim \text{Multinomial}(b_{i,1}, b_{i,2}, \dots, b_{i,K}), \forall i \in I.$$

or in general:

$$p(x_t | y_t^i = \mathbf{1}) \sim f(\cdot | \theta_i), \forall i \in I.$$



Inference (review)

- Forward algorithm

$$\alpha_t^k \stackrel{\text{def}}{=} \mu_{t-1 \rightarrow t}(k) = P(x_1, \dots, x_{t-1}, x_t, y_t^k = 1)$$

$$\alpha_t^k = p(x_t | y_t^k = 1) \sum_i \alpha_{t-1}^i a_{i,k}$$

- Backward algorithm

$$\beta_t^k = \sum_i a_{k,i} p(x_{t+1} | y_{t+1}^i = 1) \beta_{t+1}^i$$

$$\beta_t^k \stackrel{\text{def}}{=} \mu_{t-1 \leftarrow t}(k) = P(x_{t+1}, \dots, x_T | y_t^k = 1)$$

$$\gamma_t^i \stackrel{\text{def}}{=} p(y_t^i = 1 | x_{1:T}) \propto \alpha_t^i \beta_t^i = \sum_j \xi_t^{i,j}$$

$$\xi_t^{i,j} \stackrel{\text{def}}{=} p(y_t^i = 1, y_{t+1}^j = 1, x_{1:T})$$

$$\propto \mu_{t-1 \rightarrow t}(y_t^i = 1) \mu_{t \leftarrow t+1}(y_{t+1}^j = 1) p(x_{t+1} | y_{t+1}) p(y_{t+1} | y_t)$$

$$\xi_t^{i,j} = \alpha_t^i \beta_{t+1}^j a_{i,j} p(x_{t+1} | y_{t+1}^i = 1)$$

The matrix-vector form:

$$B_t(i) \stackrel{\text{def}}{=} p(x_t | y_t^i = 1)$$

$$A(i, j) \stackrel{\text{def}}{=} p(y_{t+1}^j = 1 | y_t^i = 1)$$

$$\alpha_t = (A^T \alpha_{t-1}) .* B_t$$

$$\beta_t = A(\beta_{t+1} .* B_{t+1})$$

$$\xi_t = (\alpha_t (\beta_{t+1} .* B_{t+1})^T) .* A$$

$$\gamma_t = \alpha_t .* \beta_t$$



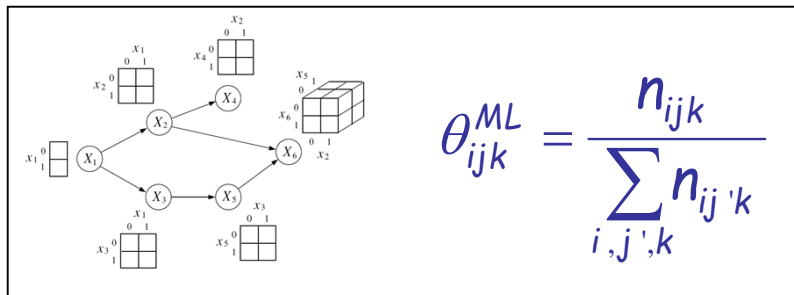
Learning HMM

- **Supervised learning**: estimation when the “right answer” is known
 - **Examples:**
 - GIVEN:** a genomic region $x = x_1 \dots x_{1,000,000}$ where we have good (experimental) annotations of the CpG islands
 - GIVEN:** the casino player allows us to observe him one evening, as he changes dice and produces 10,000 rolls
 - **Unsupervised learning**: estimation when the “right answer” is unknown
 - **Examples:**
 - GIVEN:** the porcupine genome; we don't know how frequent are the CpG islands there, neither do we know their composition
 - GIVEN:** 10,000 rolls of the casino player, but we don't see when he changes dice
- **QUESTION:** Update the parameters θ of the model to maximize $P(x|\theta)$ -
-- Maximal likelihood (ML) estimation



Learning HMM: two scenarios

- Supervised learning: if only we knew the true state path then ML parameter estimation would be trivial
 - E.g., recall that for complete observed tabular BN:



$$\theta_{ijk}^{ML} = \frac{n_{ijk}}{\sum_{i,j',k} n_{ij'k}}$$

$$a_{ij}^{ML} = \frac{\#(i \rightarrow j)}{\#(i \rightarrow \bullet)} = \frac{\sum_n \sum_{t=2}^T y_{n,t-1}^i y_{n,t}^j}{\sum_n \sum_{t=2}^T y_{n,t-1}^i}$$

$$b_{ik}^{ML} = \frac{\#(i \rightarrow k)}{\#(i \rightarrow \bullet)} = \frac{\sum_n \sum_{t=1}^T y_{n,t}^i x_{n,t}^k}{\sum_n \sum_{t=1}^T y_{n,t}^i}$$

- What if y is continuous? We can treat $\{(x_{n,t}, y_{n,t}) : t = 1:T, n = 1:N\}$ as $N \times T$ observations of, e.g., a GLIM, and apply learning rules for GLIM ...
- Unsupervised learning: when the true state path is unknown, we can fill in the missing values using inference recursions.
 - The Baum Welch algorithm (i.e., EM)
 - Guaranteed to increase the log likelihood of the model after each iteration
 - Converges to local optimum, depending on initial conditions



The Baum Welch algorithm

- The complete log likelihood

$$\ell_c(\boldsymbol{\theta}; \mathbf{x}, \mathbf{y}) = \log p(\mathbf{x}, \mathbf{y}) = \log \prod_n \left(p(y_{n,1}) \prod_{t=2}^T p(y_{n,t} | y_{n,t-1}) \prod_{t=1}^T p(x_{n,t} | x_{n,t}) \right)$$

- The expected complete log likelihood

$$\langle \ell_c(\boldsymbol{\theta}; \mathbf{x}, \mathbf{y}) \rangle = \sum_n \left(\langle y_{n,1}^i \rangle_{p(y_{n,1} | \mathbf{x}_n)} \log \pi_i \right) + \sum_n \sum_{t=2}^T \left(\langle y_{n,t-1}^i y_{n,t}^j \rangle_{p(y_{n,t-1}, y_{n,t} | \mathbf{x}_n)} \log a_{i,j} \right) + \sum_n \sum_{t=1}^T \left(x_{n,t}^k \langle y_{n,t}^i \rangle_{p(y_{n,t} | \mathbf{x}_n)} \log b_{i,k} \right)$$

- EM

- The E step

$$\gamma_{n,t}^i = \langle y_{n,t}^i \rangle = p(y_{n,t}^i = \mathbf{1} | \mathbf{x}_n)$$

$$\xi_{n,t}^{i,j} = \langle y_{n,t-1}^i y_{n,t}^j \rangle = p(y_{n,t-1}^i = \mathbf{1}, y_{n,t}^j = \mathbf{1} | \mathbf{x}_n)$$

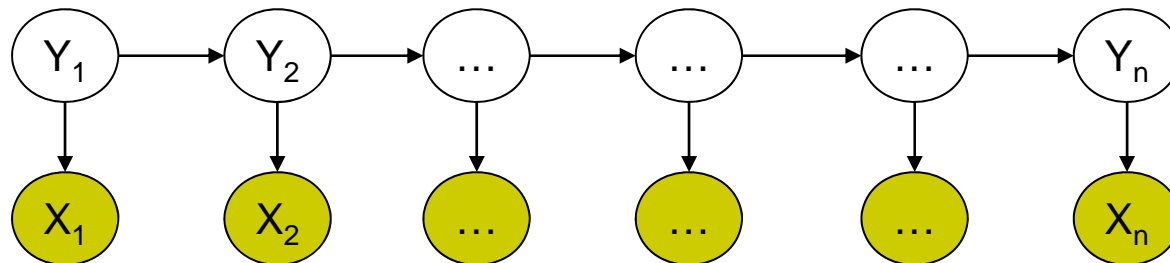
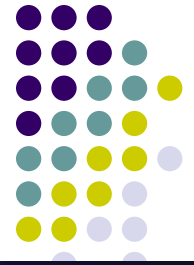
- The M step ("symbolically" identical to MLE)

$$\pi_i^{ML} = \frac{\sum_n \gamma_{n,1}^i}{N}$$

$$a_{ij}^{ML} = \frac{\sum_n \sum_{t=2}^T \xi_{n,t}^{i,j}}{\sum_n \sum_{t=1}^{T-1} \gamma_{n,t}^i}$$

$$b_{ik}^{ML} = \frac{\sum_n \sum_{t=1}^T \gamma_{n,t}^i x_{n,t}^k}{\sum_n \sum_{t=1}^{T-1} \gamma_{n,t}^i}$$

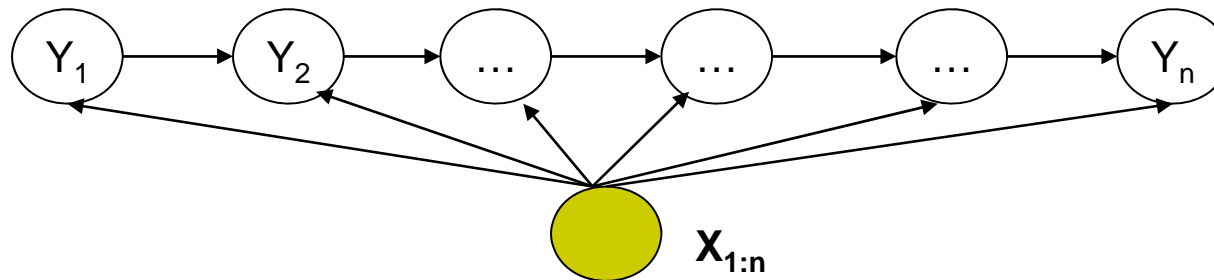
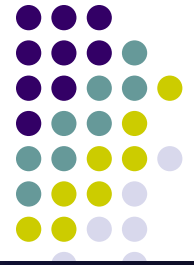
Shortcomings of Hidden Markov Model (1): locality of features



- HMM models capture dependences between each state and **only** its corresponding observation
 - NLP example: In a sentence segmentation task, each segmental state may depend not just on a single word (and the adjacent segmental stages), but also on the (non-local) features of the whole line such as line length, indentation, amount of white space, etc.
- Mismatch between learning objective function and prediction objective function
 - HMM learns a joint distribution of states and observations $P(\mathbf{Y}, \mathbf{X})$, but in a prediction task, we need the conditional probability $P(\mathbf{Y}|\mathbf{X})$

Solution:

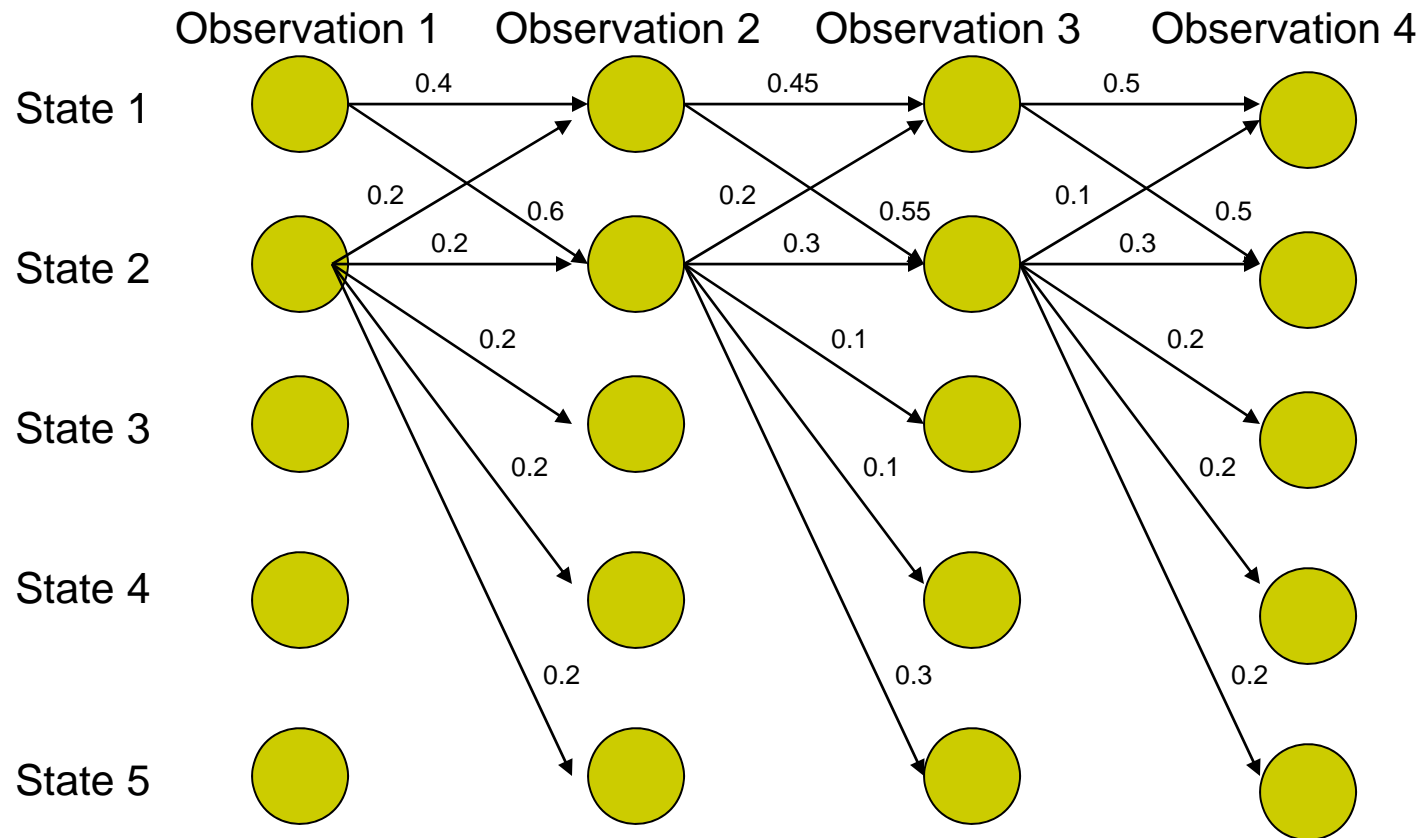
Maximum Entropy Markov Model (MEMM)



$$P(\mathbf{y}_{1:n}|\mathbf{x}_{1:n}) = \prod_{i=1}^n P(y_i|y_{i-1}, \mathbf{x}_{1:n}) = \prod_{i=1}^n \frac{\exp(\mathbf{w}^T \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_{1:n}))}{Z(y_{i-1}, \mathbf{x}_{1:n})}$$

- Models dependence between each state and the **full observation** sequence explicitly
 - More expressive than HMMs
- Discriminative model
 - Completely ignores modeling $P(\mathbf{X})$: saves modeling effort
 - Learning objective function consistent with predictive function: $P(\mathbf{Y}|\mathbf{X})$

Then, shortcomings of MEMM (and HMM) (2): the Label bias problem

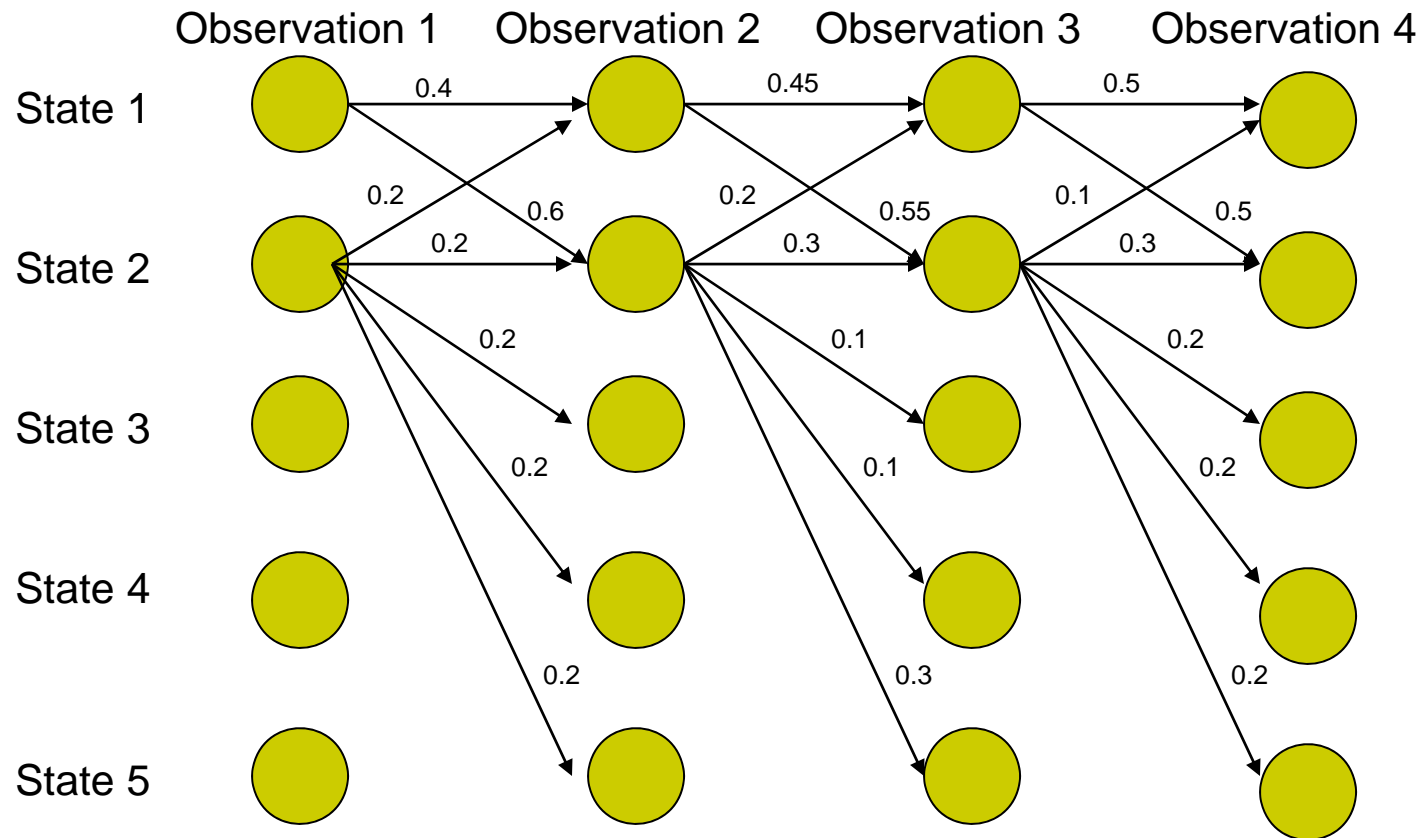


What the local transition probabilities say:

- State 1 almost always prefers to go to state 2
- State 2 almost always prefer to stay in state 2



MEMM: the Label bias problem

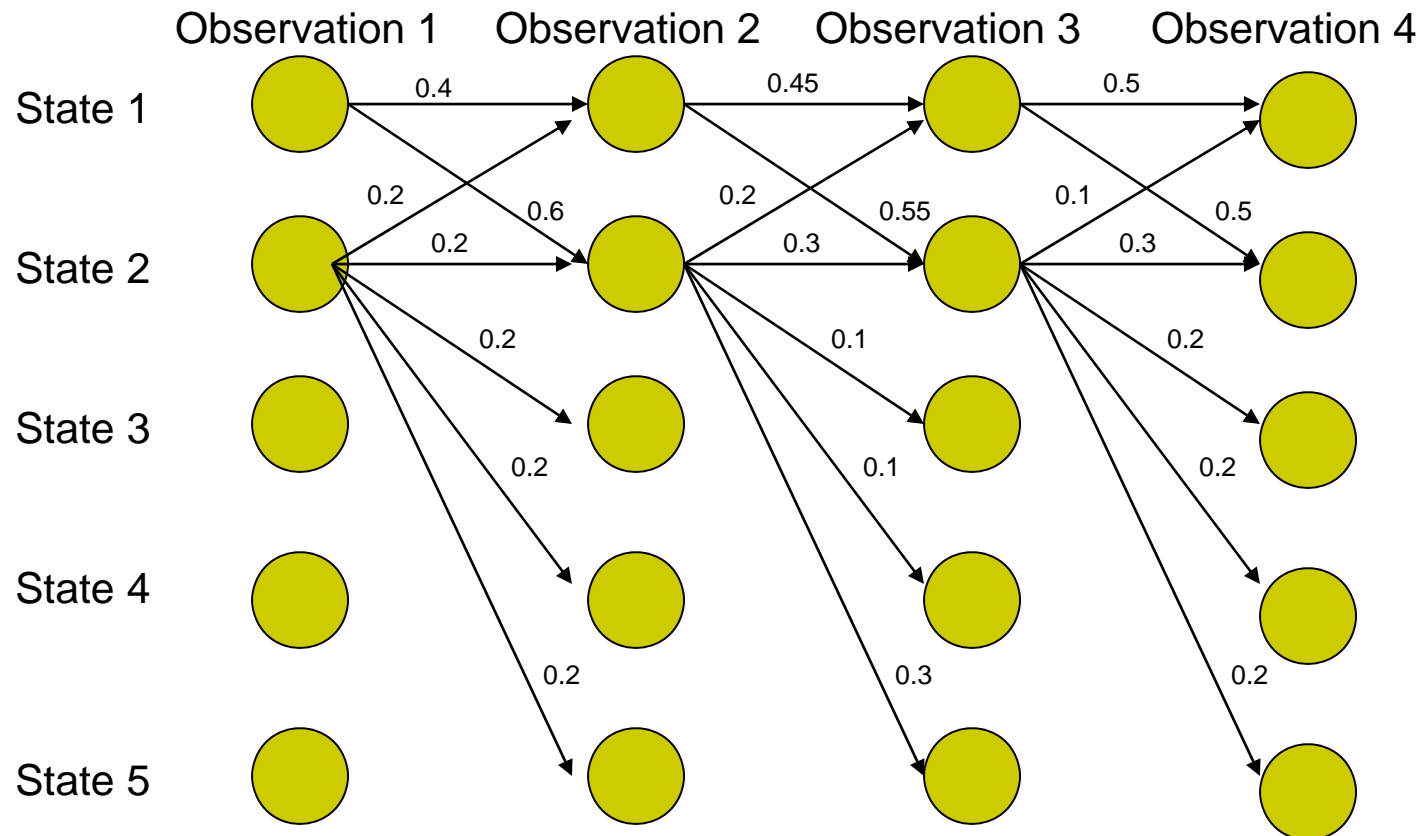


Probability of path 1-> 1-> 1-> 1:

- $0.4 \times 0.45 \times 0.5 = 0.09$



MEMM: the Label bias problem



Probability of path 2->2->2->2 :

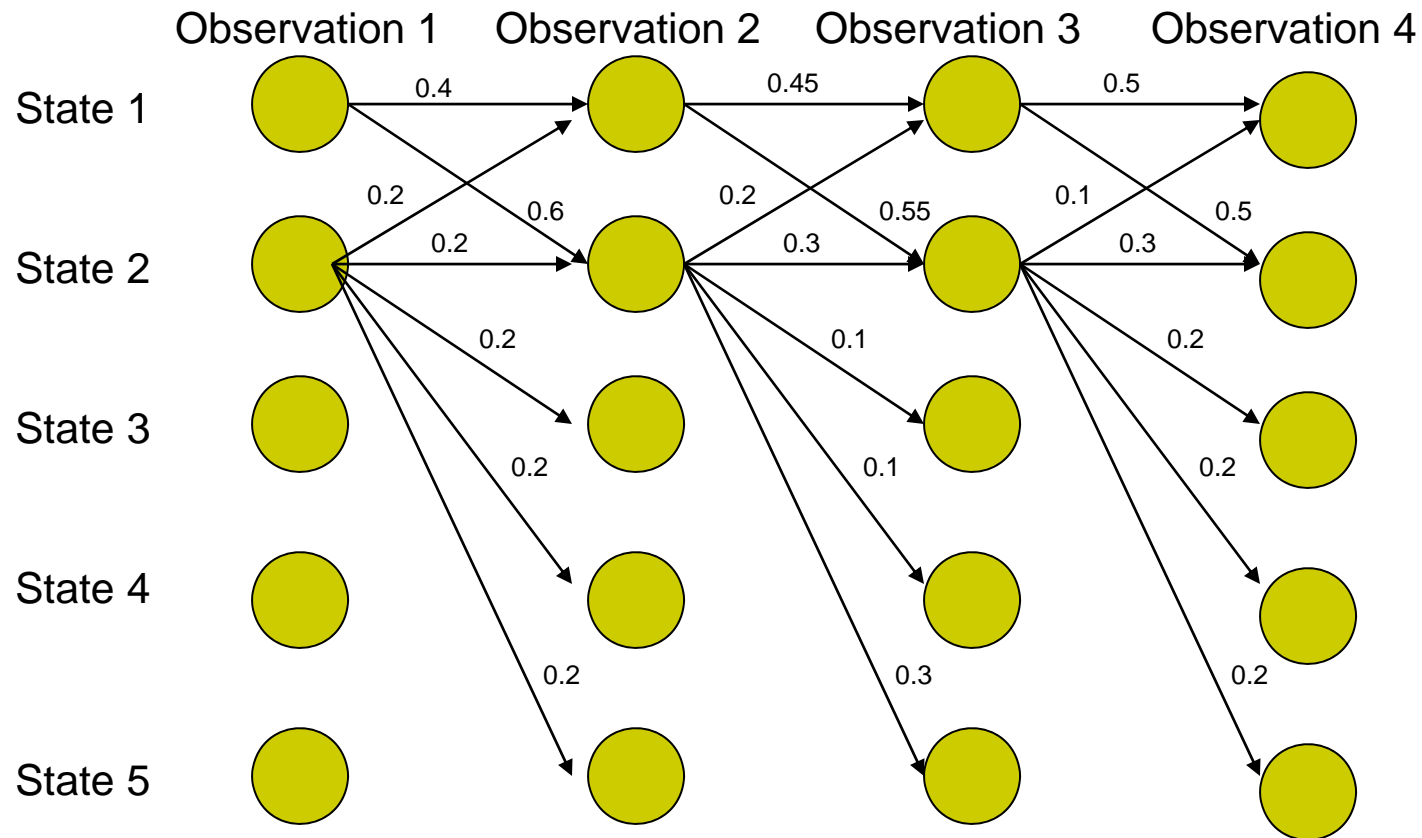
- $0.2 \times 0.3 \times 0.3 = 0.018$

Other paths:

1-> 1-> 1-> 1: 0.09



MEMM: the Label bias problem



Probability of path 1->2->1->2:

- $0.6 \times 0.2 \times 0.5 = 0.06$

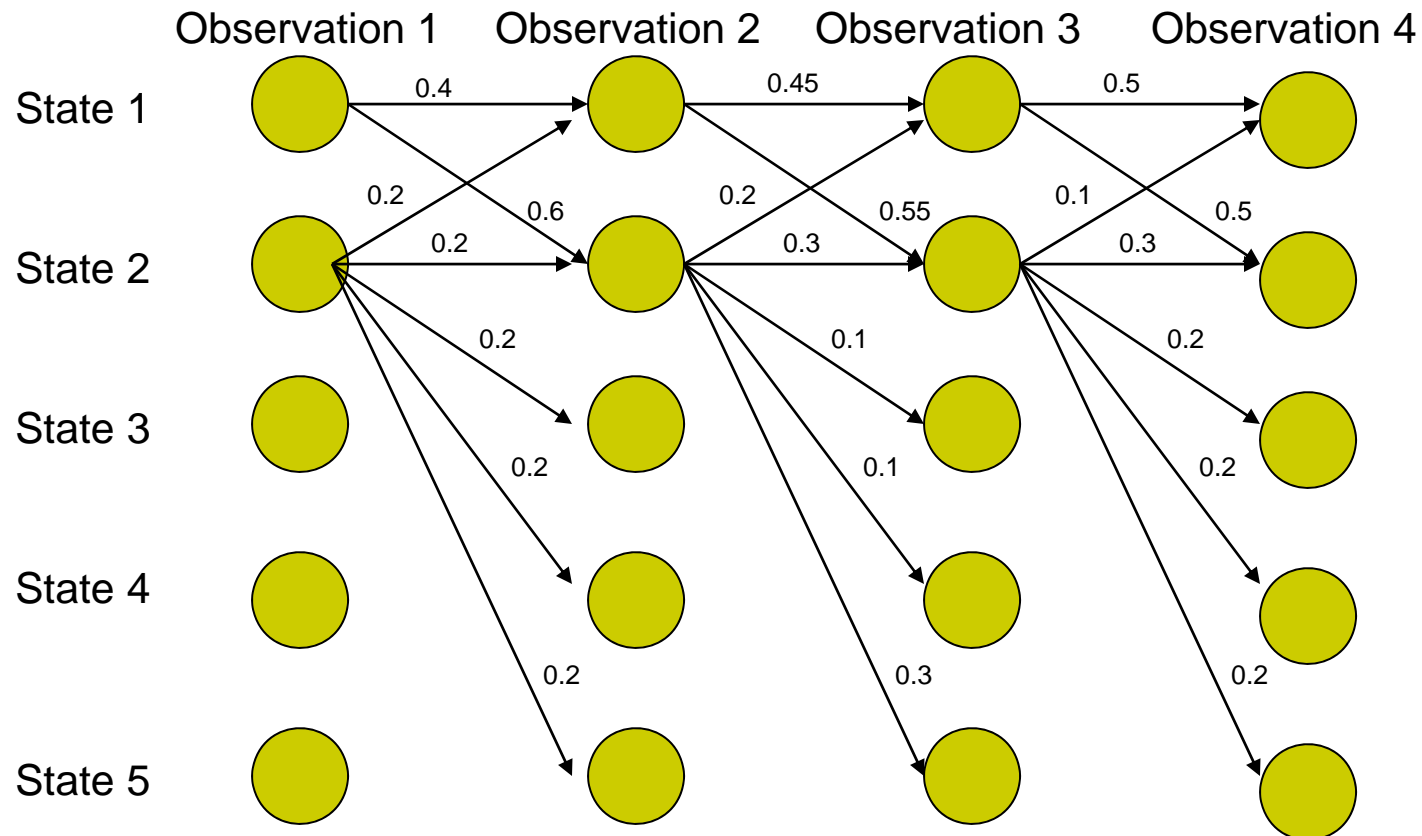
Other paths:

1->1->1->1: 0.09

2->2->2->2: 0.018



MEMM: the Label bias problem



Probability of path 1->1->2->2:

- $0.4 \times 0.55 \times 0.3 = 0.066$

Other paths:

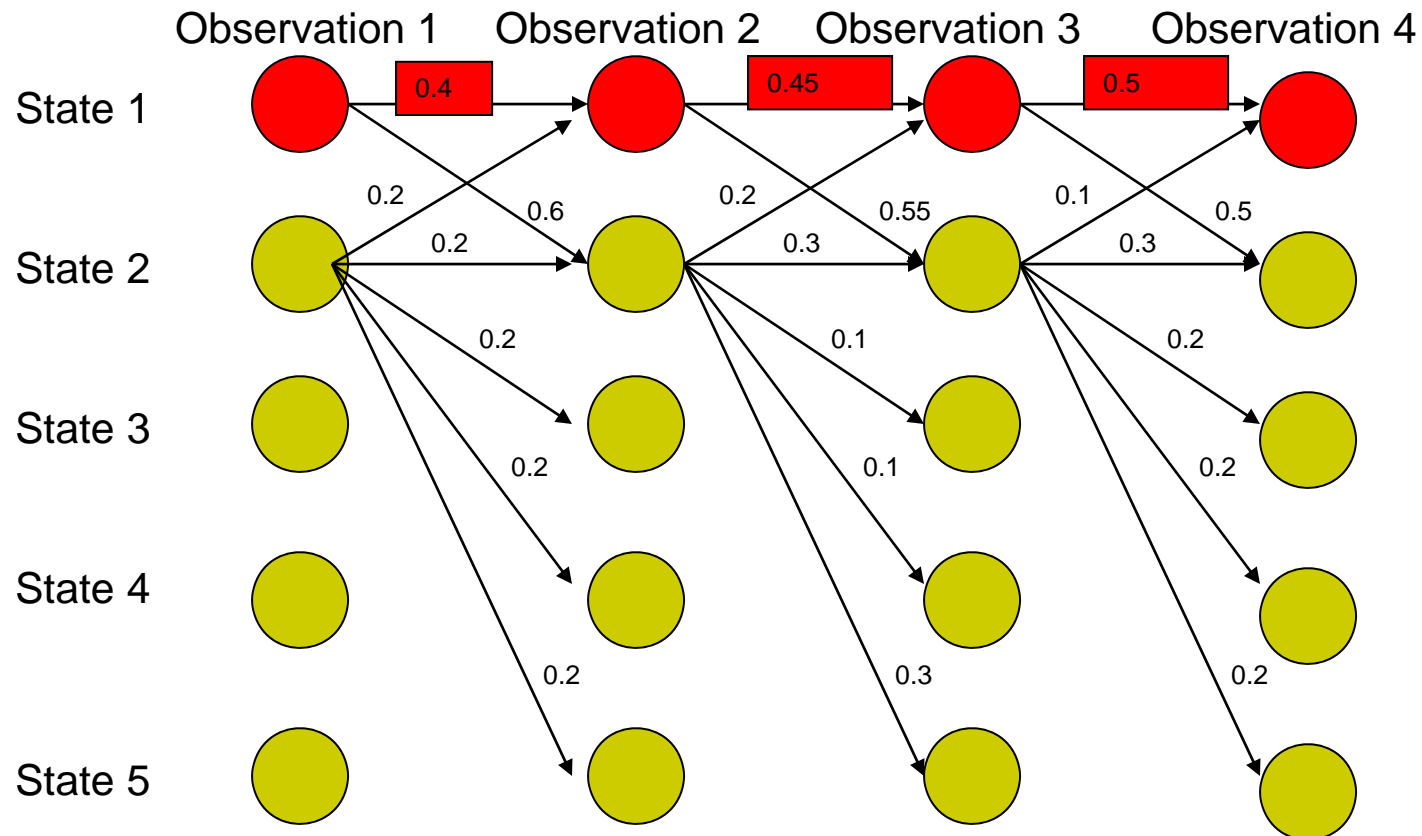
1->1->1->1: 0.09

2->2->2->2: 0.018

1->2->1->2: 0.06



MEMM: the Label bias problem

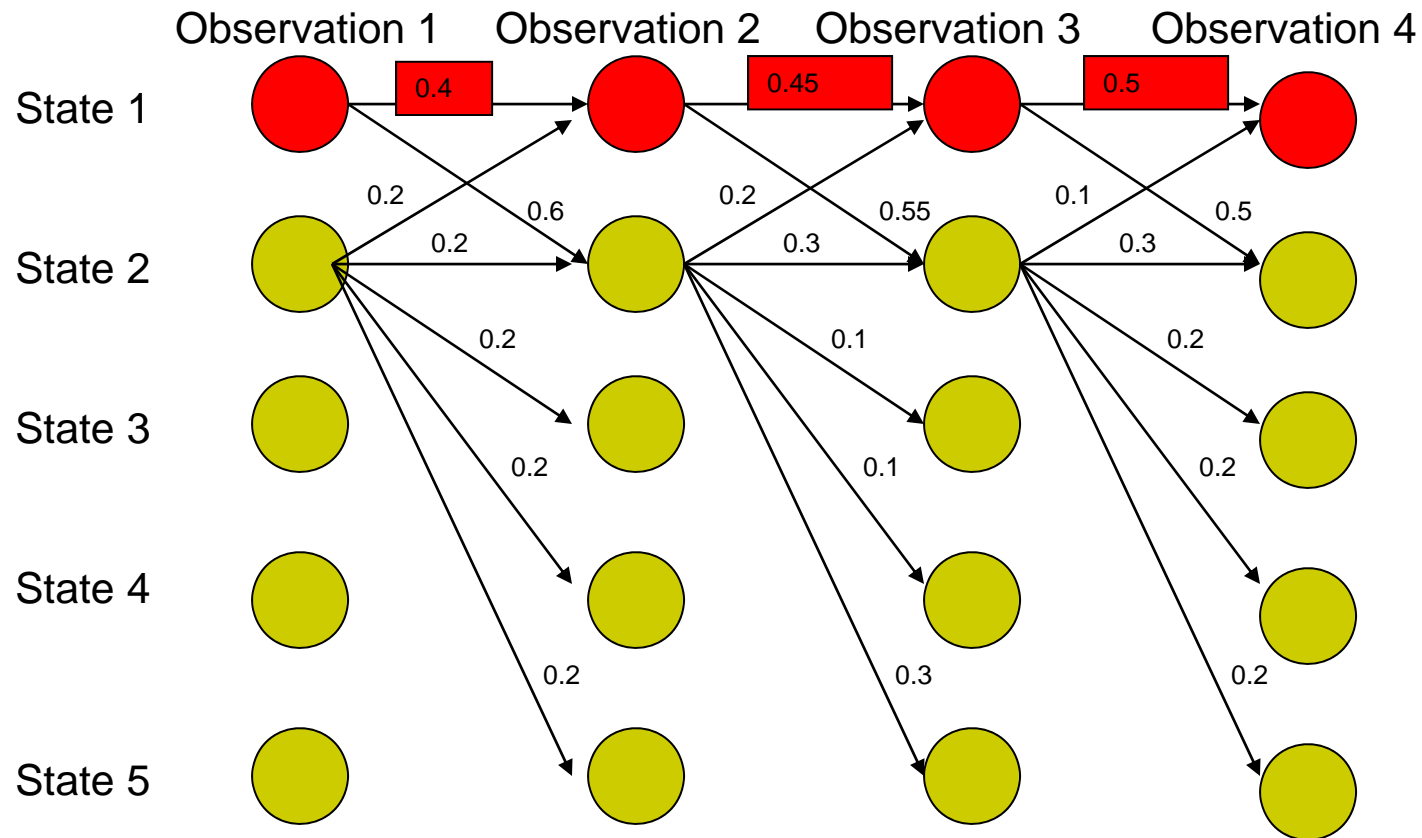


Most Likely Path: 1-> 1-> 1-> 1

- Although **locally** it seems state 1 wants to go to state 2 and state 2 wants to remain in state 2.
- **why?**



MEMM: the Label bias problem

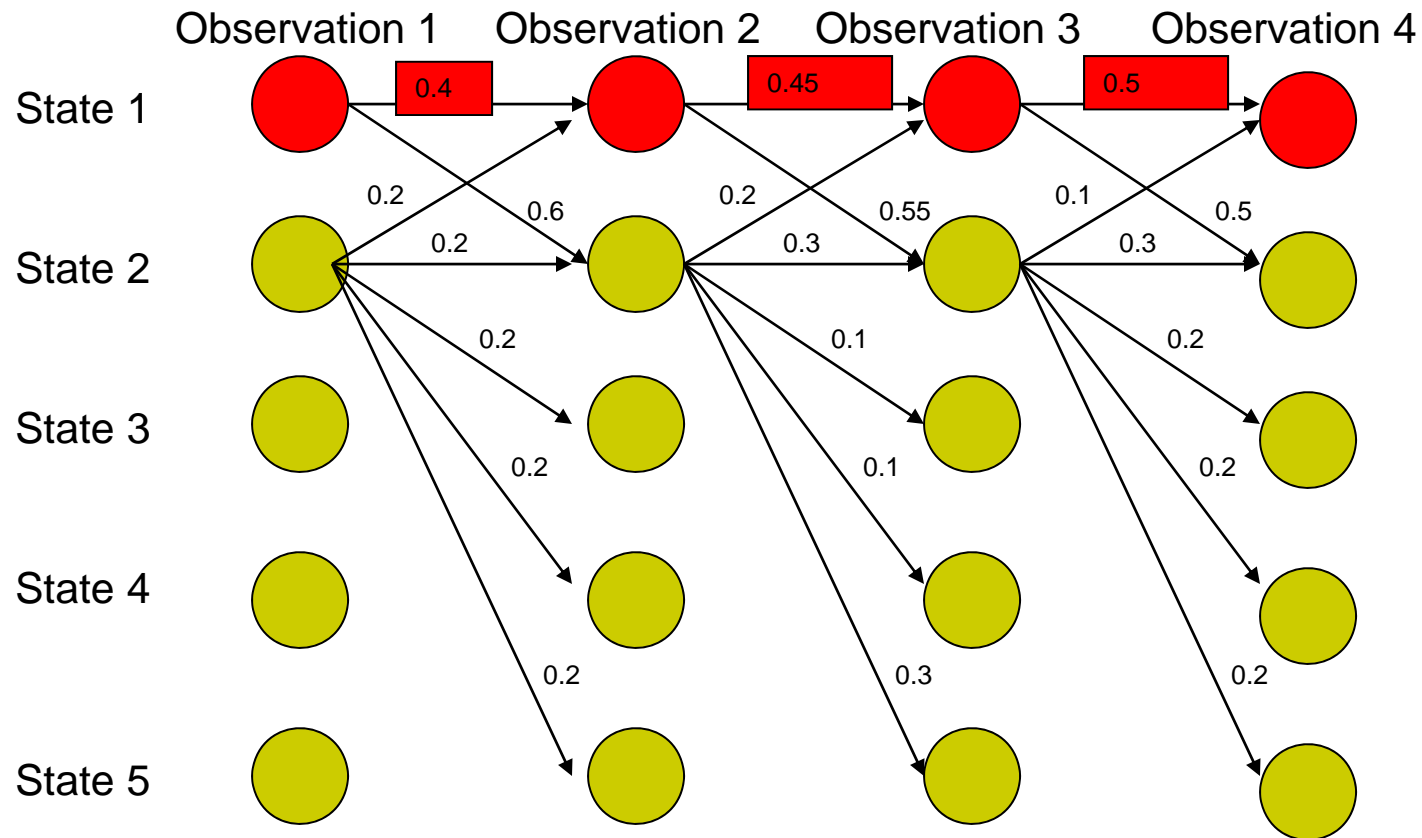


Most Likely Path: 1-> 1-> 1-> 1

- State 1 has only two transitions but state 2 has 5:
 - Average transition probability from state 2 is lower



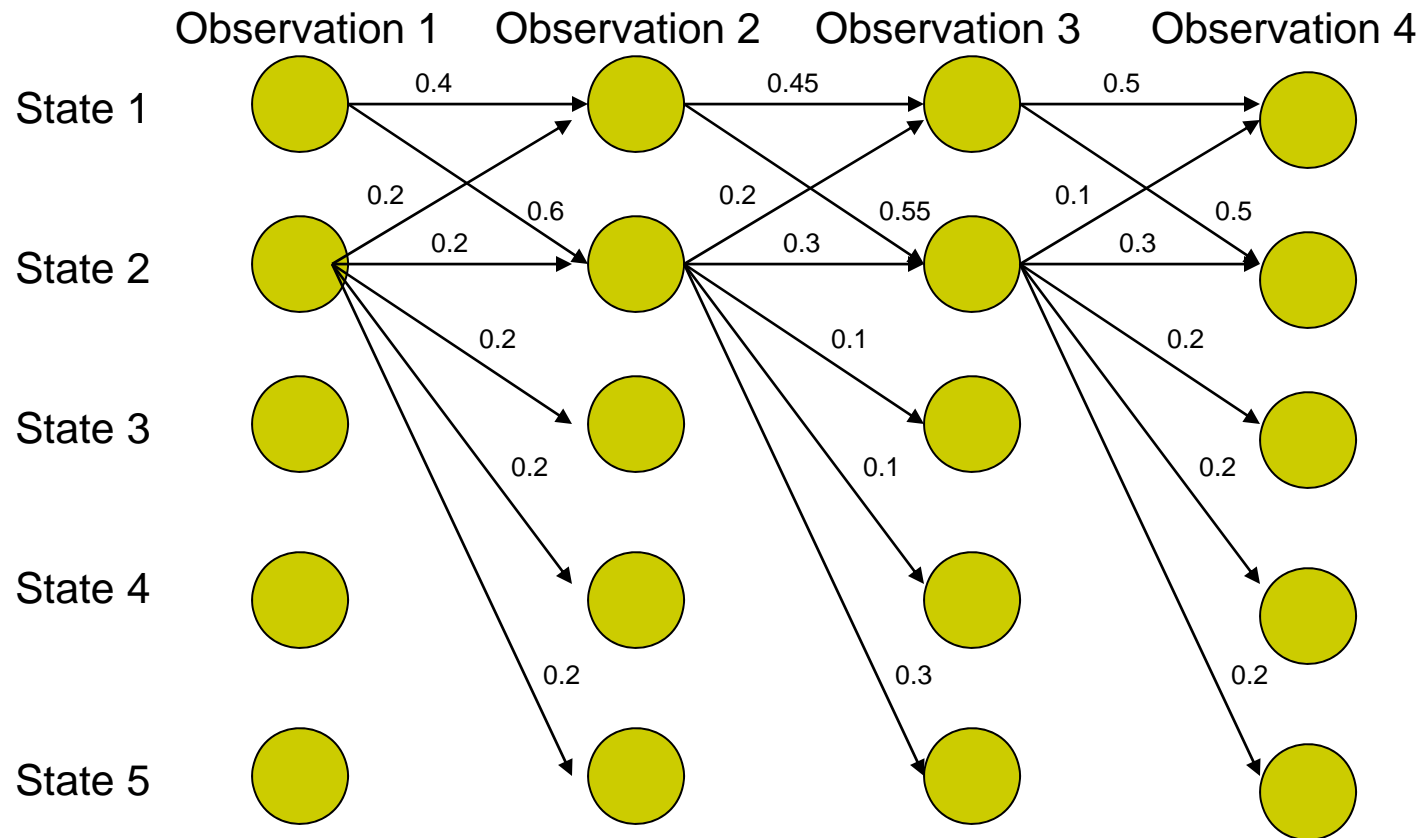
MEMM: the Label bias problem



Label bias problem in MEMM:

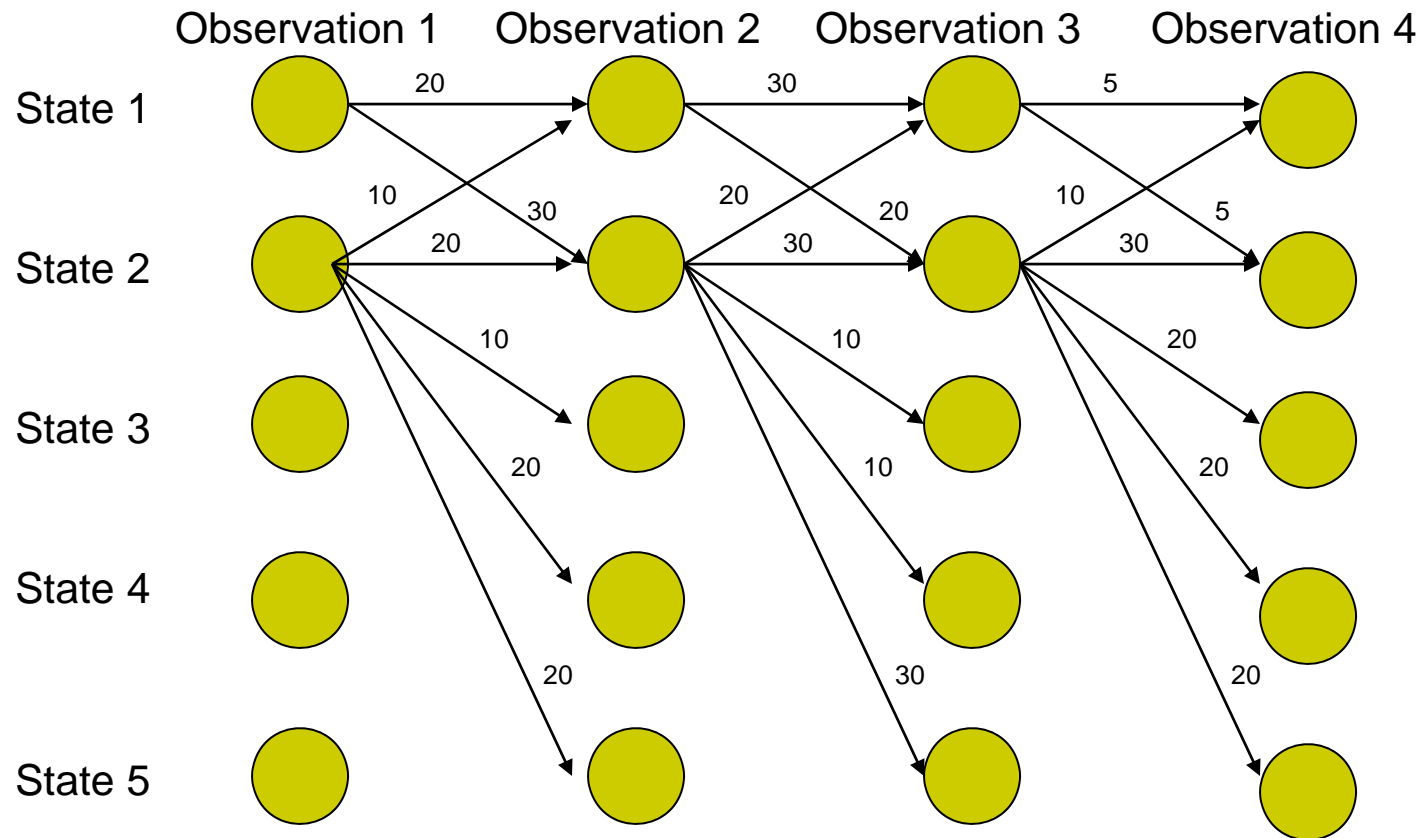
- Preference of states with lower number of transitions over others

Solution: Do not normalize probabilities locally



From local probabilities

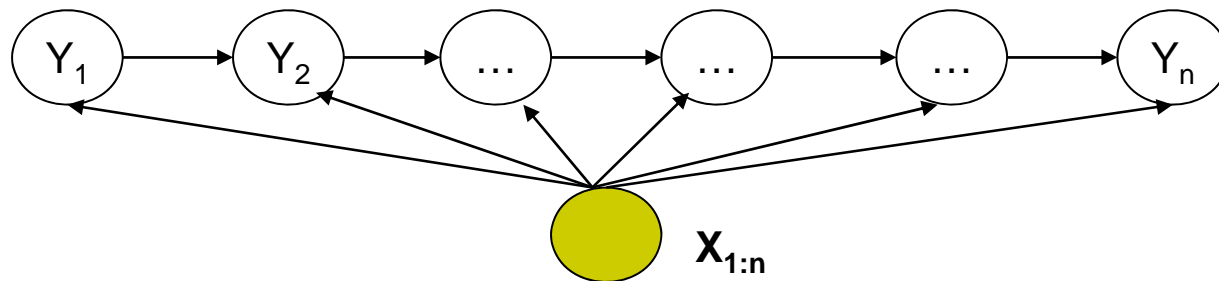
Solution: Do not normalize probabilities locally



From local probabilities to local potentials

- States with lower transitions do not have an unfair advantage!

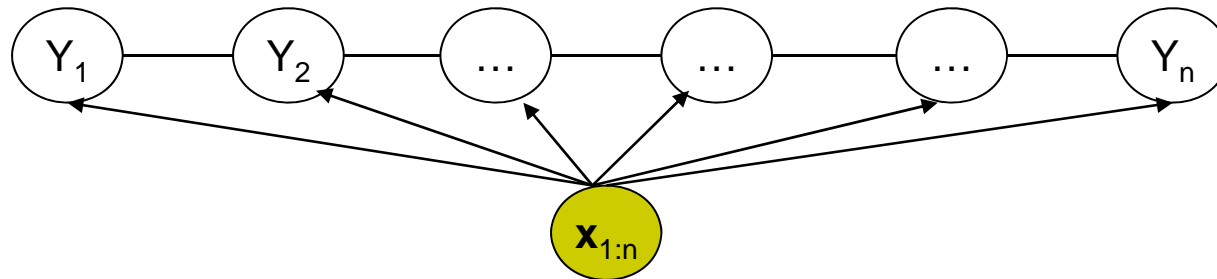
From MEMM



$$P(\mathbf{y}_{1:n} | \mathbf{x}_{1:n}) = \prod_{i=1}^n P(y_i | y_{i-1}, \mathbf{x}_{1:n}) = \prod_{i=1}^n \frac{\exp(\mathbf{w}^T \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_{1:n}))}{Z(y_{i-1}, \mathbf{x}_{1:n})}$$



From MEMM to CRF



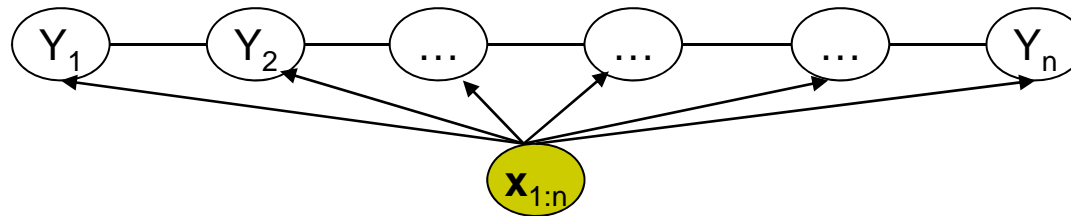
$$P(\mathbf{y}_{1:n} | \mathbf{x}_{1:n}) = \frac{1}{Z(\mathbf{x}_{1:n})} \prod_{i=1}^n \phi(y_i, y_{i-1}, \mathbf{x}_{1:n}) = \frac{1}{Z(\mathbf{x}_{1:n}, \mathbf{w})} \prod_{i=1}^n \exp(\mathbf{w}^T \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_{1:n}))$$

- CRF is a partially directed model
 - Discriminative model like MEMM
 - Usage of global normalizer $Z(\mathbf{x})$ overcomes the label bias problem of MEMM
 - Models the dependence between each state and the entire observation sequence (like MEMM)



Conditional Random Fields

- General parametric form:



$$\begin{aligned} P(\mathbf{y}|\mathbf{x}) &= \frac{1}{Z(\mathbf{x}, \lambda, \mu)} \exp\left(\sum_{i=1}^n \left(\sum_k \lambda_k f_k(y_i, y_{i-1}, \mathbf{x}) + \sum_l \mu_l g_l(y_i, \mathbf{x})\right)\right) \\ &= \frac{1}{Z(\mathbf{x}, \lambda, \mu)} \exp\left(\sum_{i=1}^n (\lambda^T \mathbf{f}(y_i, y_{i-1}, \mathbf{x}) + \mu^T \mathbf{g}(y_i, \mathbf{x}))\right) \end{aligned}$$

$$\text{where } Z(\mathbf{x}, \lambda, \mu) = \sum_{\mathbf{y}} \exp\left(\sum_{i=1}^n (\lambda^T \mathbf{f}(y_i, y_{i-1}, \mathbf{x}) + \mu^T \mathbf{g}(y_i, \mathbf{x}))\right)$$

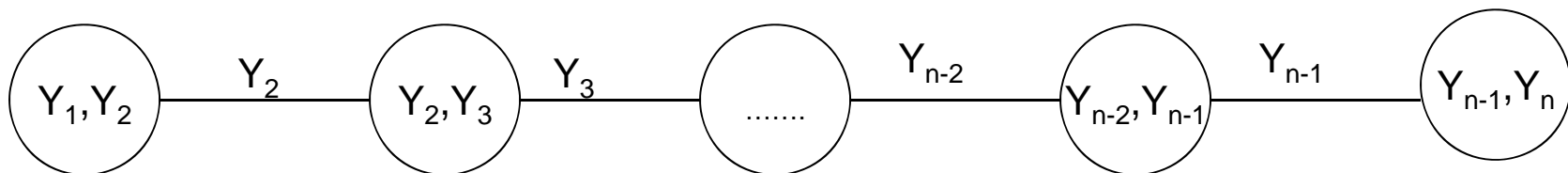
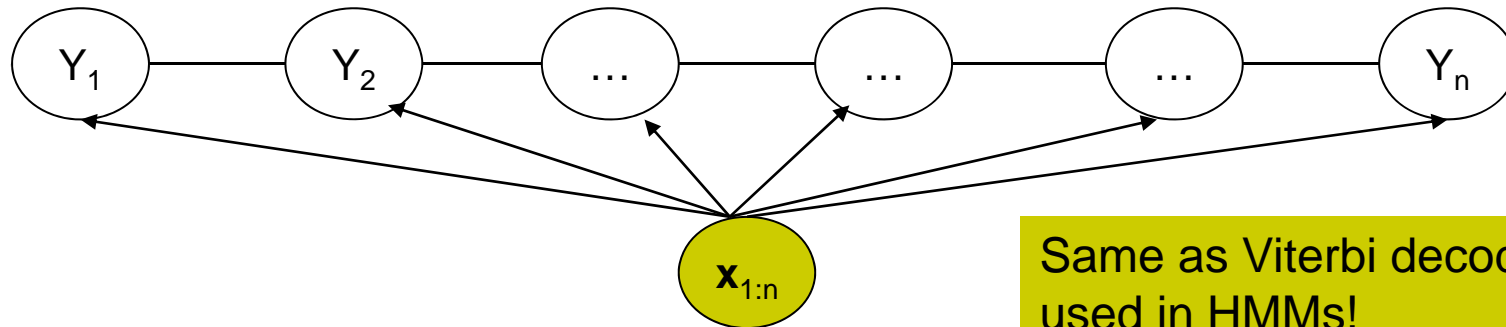


CRFs: Inference

- Given CRF parameters λ and μ , find the \mathbf{y}^* that maximizes $P(\mathbf{y}|\mathbf{x})$

$$\mathbf{y}^* = \arg \max_{\mathbf{y}} \exp\left(\sum_{i=1}^n (\lambda^T \mathbf{f}(y_i, y_{i-1}, \mathbf{x}) + \mu^T \mathbf{g}(y_i, \mathbf{x}))\right)$$

- Can ignore $Z(\mathbf{x})$ because it is not a function of \mathbf{y}
- Run the max-product algorithm on the junction-tree of CRF:





CRF learning

- Given $\{(\mathbf{x}_d, \mathbf{y}_d)\}_{d=1}^N$, find λ^*, μ^* such that

$$\begin{aligned}\lambda^*, \mu^* &= \arg \max_{\lambda, \mu} L(\lambda, \mu) = \arg \max_{\lambda, \mu} \prod_{d=1}^N P(\mathbf{y}_d | \mathbf{x}_d, \lambda, \mu) \\ &= \arg \max_{\lambda, \mu} \prod_{d=1}^N \frac{1}{Z(\mathbf{x}_d, \lambda, \mu)} \exp\left(\sum_{i=1}^n (\lambda^T \mathbf{f}(y_{d,i}, y_{d,i-1}, \mathbf{x}_d) + \mu^T \mathbf{g}(y_{d,i}, \mathbf{x}_d))\right) \\ &= \arg \max_{\lambda, \mu} \sum_{d=1}^N \left(\sum_{i=1}^n (\lambda^T \mathbf{f}(y_{d,i}, y_{d,i-1}, \mathbf{x}_d) + \mu^T \mathbf{g}(y_{d,i}, \mathbf{x}_d)) - \log Z(\mathbf{x}_d, \lambda, \mu)\right)\end{aligned}$$

- Computing the gradient w.r.t λ :

Gradient of the log-partition function in an exponential family is the expectation of the sufficient statistics.

$$\nabla_{\lambda} L(\lambda, \mu) = \sum_{d=1}^N \left(\sum_{i=1}^n \mathbf{f}(y_{d,i}, y_{d,i-1}, \mathbf{x}_d) - \sum_{\mathbf{y}} (P(\mathbf{y} | \mathbf{x}_d) \sum_{i=1}^n \mathbf{f}(y_{d,i}, y_{d,i-1}, \mathbf{x}_d))\right)$$



CRF learning

$$\nabla_{\lambda} L(\lambda, \mu) = \sum_{d=1}^N \left(\sum_{i=1}^n \mathbf{f}(y_{d,i}, y_{d,i-1}, \mathbf{x}_d) - \sum_{\mathbf{y}} (P(\mathbf{y} | \mathbf{x}_d) \sum_{i=1}^n \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d)) \right)$$

- Computing the model expectations:

- Requires exponentially large number of summations: Is it intractable?

$$\begin{aligned} \sum_{\mathbf{y}} (P(\mathbf{y} | \mathbf{x}_d) \sum_{i=1}^n \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d)) &= \sum_{i=1}^n \left(\sum_{\mathbf{y}} \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d) P(\mathbf{y} | \mathbf{x}_d) \right) \\ &= \sum_{i=1}^n \sum_{y_i, y_{i-1}} \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d) P(y_i, y_{i-1} | \mathbf{x}_d) \end{aligned}$$

Expectation of \mathbf{f} over the corresponding marginal probability of neighboring nodes!!

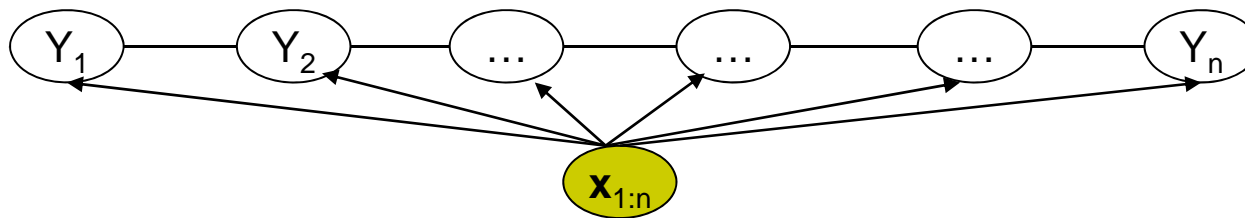
- Tractable!

- Can compute marginals using the sum-product algorithm on the chain



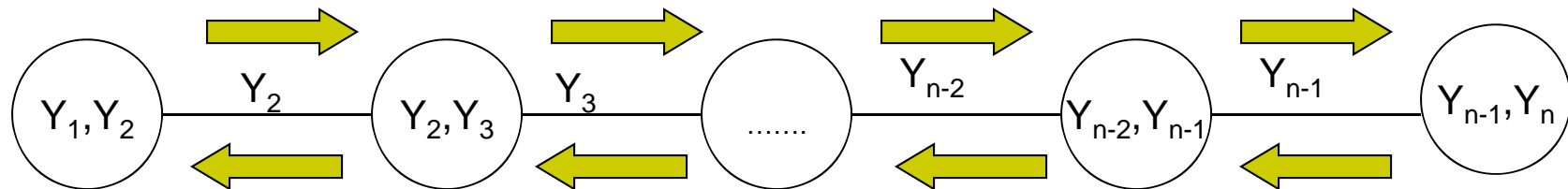
CRF learning

- Computing marginals using junction-tree calibration:



- Junction Tree Initialization:

$$\alpha^0(y_i, y_{i-1}) = \exp(\lambda^T \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d) + \mu^T \mathbf{g}(y_i, \mathbf{x}_d))$$



- After calibration:

$$P(y_i, y_{i-1} | \mathbf{x}_d) \propto \alpha(y_i, y_{i-1})$$

$$\Rightarrow P(y_i, y_{i-1} | \mathbf{x}_d) = \frac{\alpha(y_i, y_{i-1})}{\sum_{y_i, y_{i-1}} \alpha(y_i, y_{i-1})} = \alpha'(y_i, y_{i-1})$$

Also called forward-backward algorithm



CRF learning

- Computing feature expectations using calibrated potentials:

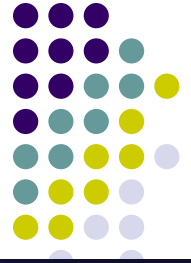
$$\sum_{y_i, y_{i-1}} \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d) P(y_i, y_{i-1} | \mathbf{x}_d) = \sum_{y_i, y_{i-1}} \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d) \alpha'(y_i, y_{i-1})$$

- Now we know how to compute $r_\lambda L(\lambda, \mu)$:

$$\begin{aligned} \nabla_\lambda L(\lambda, \mu) &= \sum_{d=1}^N \left(\sum_{i=1}^n \mathbf{f}(y_{d,i}, y_{d,i-1}, \mathbf{x}_d) - \sum_{\mathbf{y}} (P(\mathbf{y} | \mathbf{x}_d) \sum_{i=1}^n \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d)) \right) \\ &= \sum_{d=1}^N \left(\sum_{i=1}^n (\mathbf{f}(y_{d,i}, y_{d,i-1}, \mathbf{x}_d) - \sum_{y_i, y_{i-1}} \alpha'(y_i, y_{i-1}) \mathbf{f}(y_i, y_{i-1}, \mathbf{x}_d)) \right) \end{aligned}$$

- Learning can now be done using gradient ascent:

$$\begin{aligned} \lambda^{(t+1)} &= \lambda^{(t)} + \eta \nabla_\lambda L(\lambda^{(t)}, \mu^{(t)}) \\ \mu^{(t+1)} &= \mu^{(t)} + \eta \nabla_\mu L(\lambda^{(t)}, \mu^{(t)}) \end{aligned}$$



CRF learning

- In practice, we use a Gaussian Regularizer for the parameter vector to improve generalizability

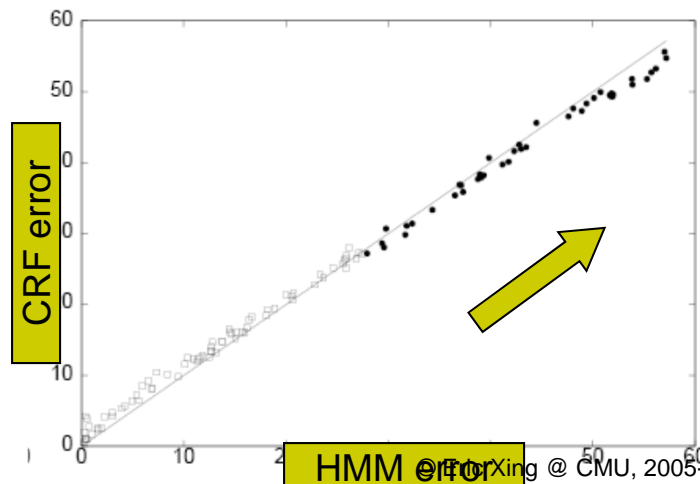
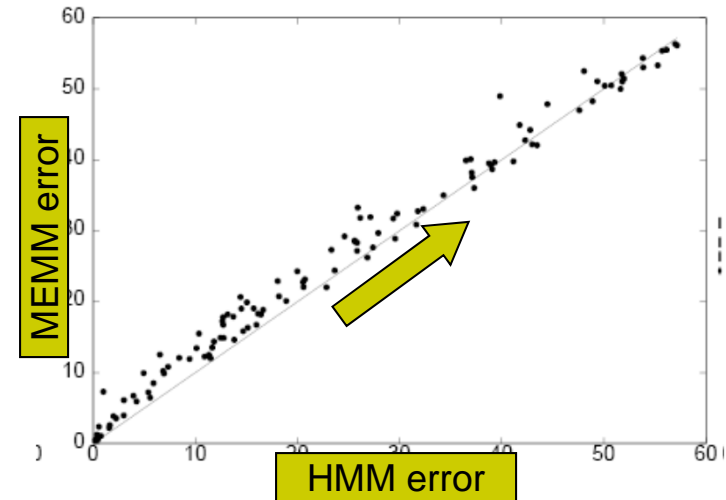
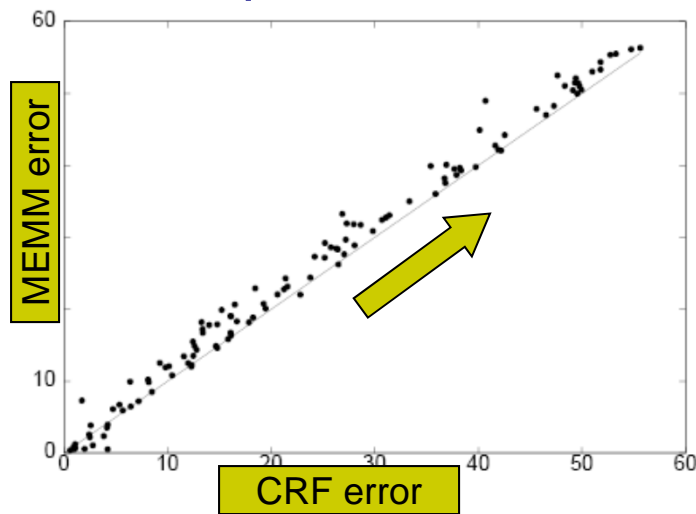
$$\lambda^*, \mu^* = \arg \max_{\lambda, \mu} \sum_{d=1}^N \log P(\mathbf{y}_d | \mathbf{x}_d, \lambda, \mu) - \frac{1}{2\sigma^2} (\lambda^T \lambda + \mu^T \mu)$$

- In practice, gradient ascent has very slow convergence
 - Alternatives:
 - Conjugate Gradient method
 - Limited Memory Quasi-Newton Methods



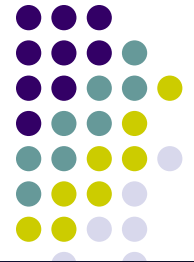
CRFs: some empirical results

- Comparison of error rates on synthetic data



Data is increasingly higher order in the direction of arrow

CRFs achieve the lowest error rate for higher order data



CRFs: some empirical results

- Parts of Speech tagging

<i>model</i>	<i>error</i>	<i>oov error</i>
HMM	5.69%	45.99%
MEMM	6.37%	54.61%
CRF	5.55%	48.05%
MEMM ⁺	4.81%	26.99%
CRF ⁺	4.27%	23.76%

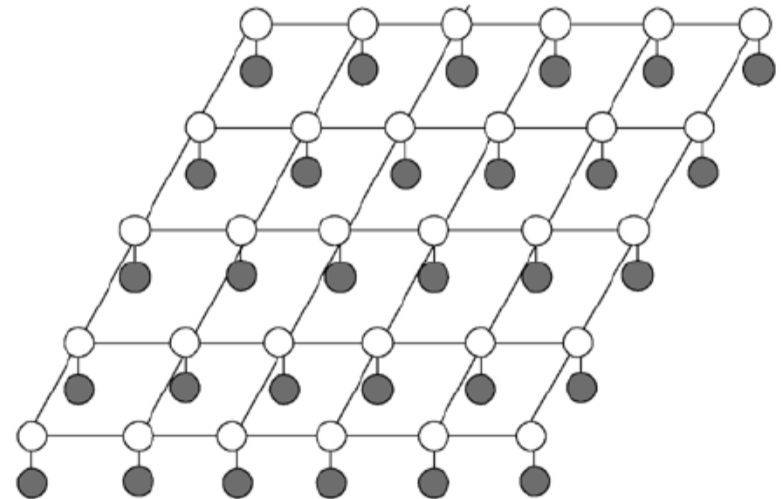
⁺Using spelling features

- Using same set of features: HMM \approx CRF > MEMM
- Using additional overlapping features: CRF⁺ > MEMM⁺ \gg HMM



Other CRFs

- So far we have discussed only 1-dimensional chain CRFs
 - Inference and learning: exact
- We could also have CRFs for arbitrary graph structure
 - E.g: Grid CRFs
 - Inference and learning no longer tractable
 - Approximate techniques used
 - MCMC Sampling
 - Variational Inference
 - Loopy Belief Propagation
 - We will discuss these techniques soon



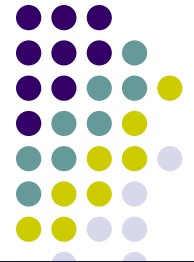


Image Segmentation

- Image segmentation (FG/BG) by modeling of interactions btw RVs
 - Images are noisy.
 - Objects occupy continuous regions in an image.

[Nowozin, Lampert 2012]



Input image



Pixel-wise separate optimal labeling

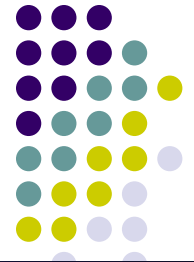


Locally-consistent joint optimal labeling

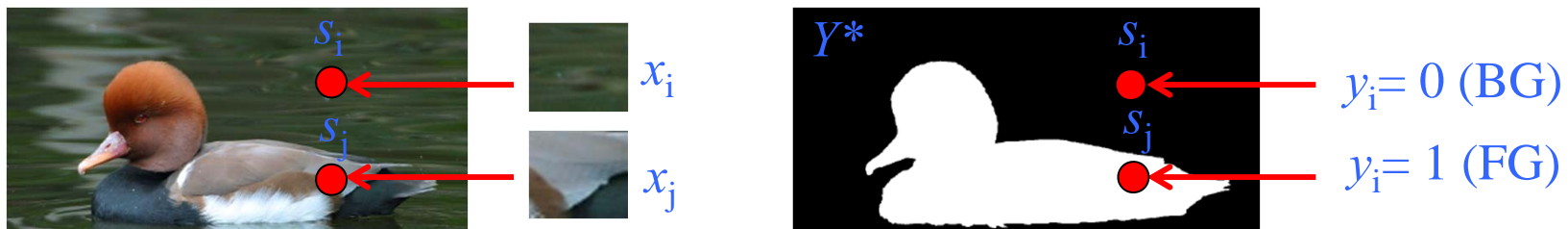
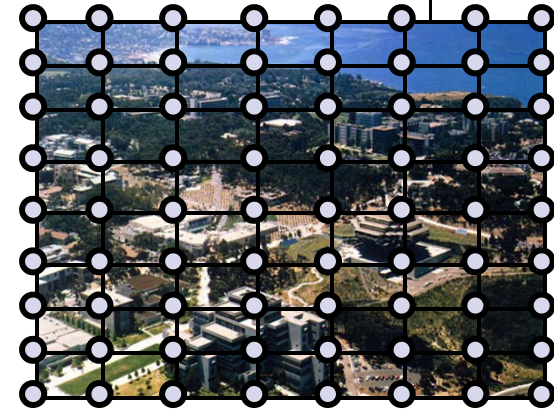
$$Y^* = \arg \max_{y \in \{0,1\}^n} \left[\overbrace{\sum_{i \in S} V_i(y_i, X)}^{\text{Unary Term}} + \overbrace{\sum_{i \in S} \sum_{j \in N_i} V_{i,j}(y_i, y_j)}^{\text{Pairwise Term}} \right].$$

Y : labels
 X : data (features)
 S : pixels
 N_i : neighbors of pixel i

Undirected Graphical Models (with an Image Labeling Example)



- Image can be represented by 4-connected 2D grid.
- MRF / CRF with image labeling problem
 - $X = \{x_i\}_{i \in S}$: observed data of an image.
 - x_i : data at i -th site (pixel or block) of the image set S
 - $Y = \{y_i\}_{i \in S}$: (hidden) labels at i -th site. $y_i \in \{1, \dots, L\}$.
- Object: maximize the conditional probability $Y^* = \operatorname{argmax}_Y P(Y|X)$





MRF (Markov Random Field)

- Definition: $Y = \{y_i\}_{i \in S}$ is called Markov Random Field on the set S , with respect to neighborhood system N , iff for all $i \in S$,

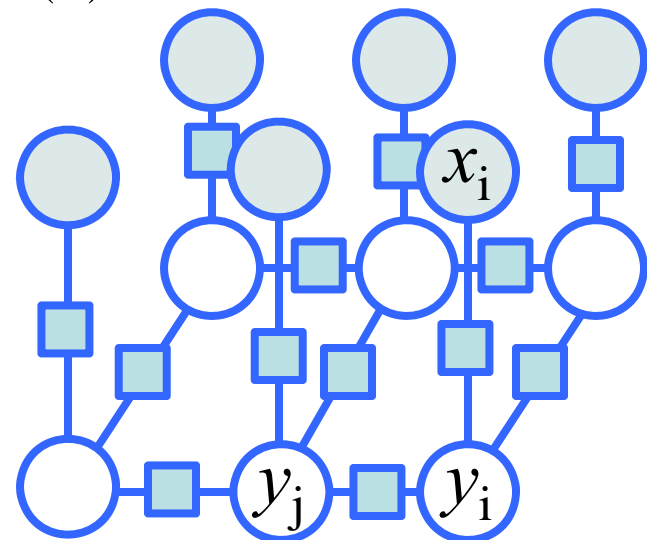
$$P(y_i | y_{S - \{i\}}) = P(y_i | y_{N_i}).$$

- The posterior probability is

$$P(Y | X) = \frac{P(X, Y)}{P(X)} \propto P(X | Y) P(Y) = \overbrace{\prod_{i \in S} P(x_i | y_i)}^{(1)} \cdot \overbrace{P(Y)}^{(2)}$$

- (1) Very strict independence assumptions for tractability: Label of each site is a function of data only at that site.
- (2) $P(Y)$ is modeled as a MRF

$$P(Y) = \frac{1}{Z} \prod_{c \in C} \psi_c(y_c)$$





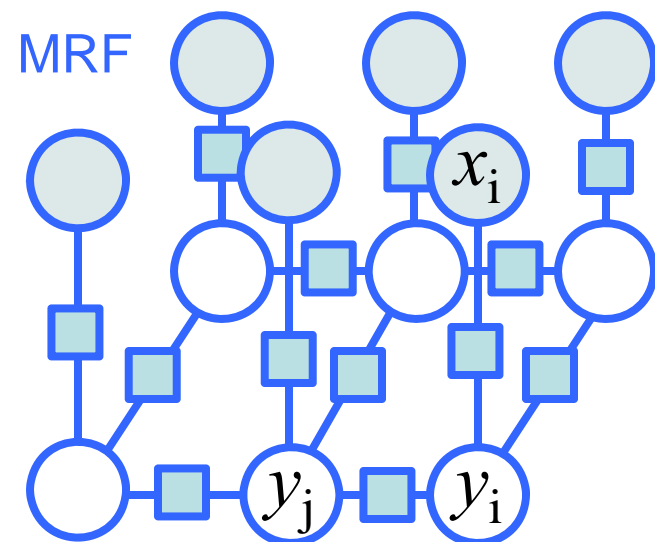
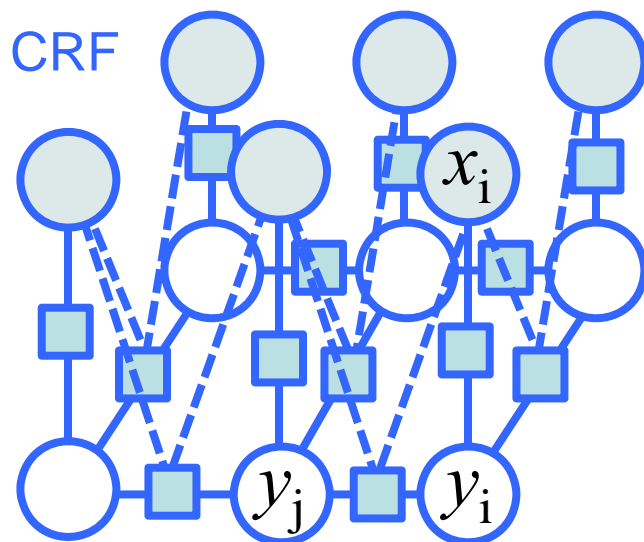
CRF

- Definition: Let $G = (S, E)$, then (X, Y) is said to be a Conditional Random Field (CRF) if, when conditioned on X , the random variables y_i obey the Markov property with respect to the graph

$$P(y_i | X, y_{S-\{i\}}) = P(y_i | X, y_{Ni})$$

$$\text{MRF: } P(y_i | y_{S-\{i\}}) = P(y_i | y_{Ni})$$

- Globally conditioned on the observation X





CRF vs MRF

- MRF: two-step generative model
 - Infer likelihood $P(X|Y)$ and prior $P(Y)$
 - Use Bayes theorem to determine posterior $P(Y|X)$

$$P(Y | X) = \frac{P(X, Y)}{P(X)} \propto P(X | Y)P(Y) = \prod_{i \in S} P(x_i | y_i) \cdot \frac{1}{Z} \prod_{c \in C} \psi_c(y_c)$$

- CRF: one-step discriminative model
 - Directly Infer posterior $P(Y|X)$

- Popular Formulation

Assumption

MRF $P(Y | X) = \frac{1}{Z} \exp\left(\sum_{i \in S} \log p(x_i | y_i) + \sum_{i \in S} \sum_{i' \in N_i} V_2(y_i, y_{i'})\right)$

Potts model for $P(Y)$ with only pairwise potential

CRF $P(Y | X) = \frac{1}{Z} \exp\left(-\sum_{i \in S} V_1(y_i | X) + \sum_{i \in S} \sum_{i' \in N_i} V_2(y_i, y_{i'} | X)\right)$

Only up to pairwise clique potentials



Example of CRF – DRF

- A special type of CRF
 - The unary and pairwise potentials are designed using local discriminative classifiers.
 - Posterior

$$P(Y | X) = \frac{1}{Z} \exp\left(\underbrace{\sum_{i \in S} A_i(y_i, X)}_{\text{Association}} + \sum_{i \in S} \sum_{j \in N_i} \underbrace{I_{ij}(y_i, y_j, X)}_{\text{Interaction}}\right)$$

- Association Potential

- Local discriminative model for site i : using logistic link with GLM.

$$A_i(y_i, X) = \log P(y_i | f_i(X)) \quad P(y_i = 1 | f_i(X)) = \frac{1}{1 + \exp(-(w^T f_i(X)))} = \sigma(w^T f_i(X))$$

- Interaction Potential

- Measure of how likely site i and j have the same label given X

$$I_{ij}(y_i, y_j, X) = \underbrace{ky_i y_j}_{(1)} + \underbrace{(1-k)(2\sigma(y_i y_j \mu_{ij}(X)) - 1)}_{(2)}$$

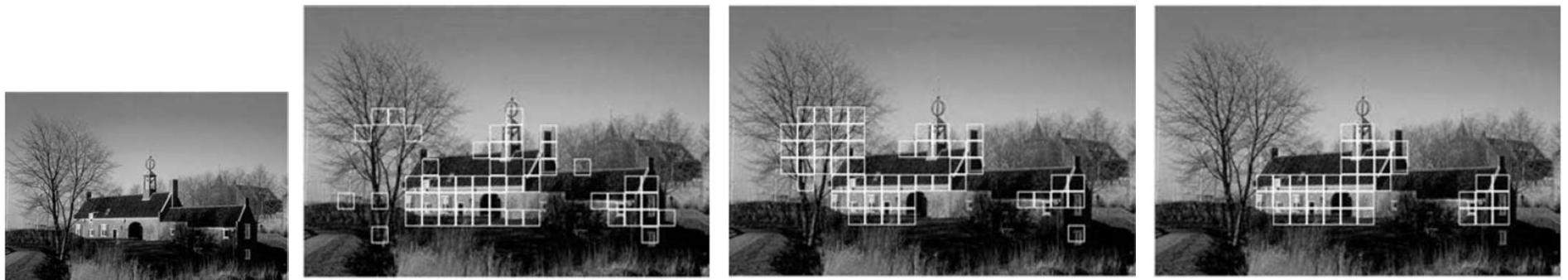
(1) Data-independent smoothing term (2) Data-dependent pairwise logistic function

S. Kumar and M. Hebert. Discriminative Random Fields. IJCV, 2006.



Example of CRF – DRF Results

- Task: Detecting man-made structure in natural scenes.
 - Each image is divided in non-overlapping 16x16 tile blocks.
- An example



Input image

Logistic

MRF

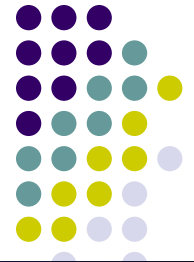
DRF

- Logistic: No smoothness in the labels
- MRF: Smoothed False positive. Lack of neighborhood interaction of the data

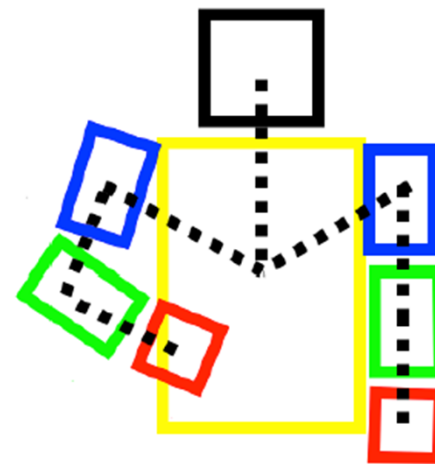
S. Kumar and M. Hebert. Discriminative Random Fields. IJCV, 2006.

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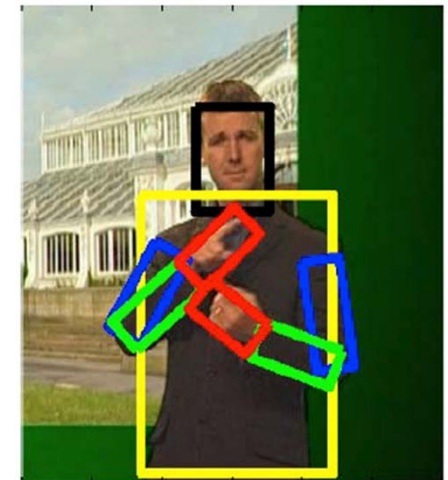
Example of CRF –Body Pose Estimation



- Task: Estimate a body pose.
 - Need to detect parts of human body
 - Appearance + Geometric configuration.
 - A large number of DOFs
- Use CRF to model a human body
 - Nodes: Parts (head, torso, upper/ lower left/right arms).
 $L=(l_1, \dots, l_6), l_i = [x_i, y_i, \theta_i]$.
 - Edges: Pairwise linkage between parts
 - Tree vs. Graph



[Zisserman 2010]



V. Ferrari et al. Progressive search space reduction for human pose estimation. CVPR 2008.
D. Ramanan. Learning to Parse Images of Articulated Bodies." NIPS 2006.

Example of CRF –Body Pose Estimation

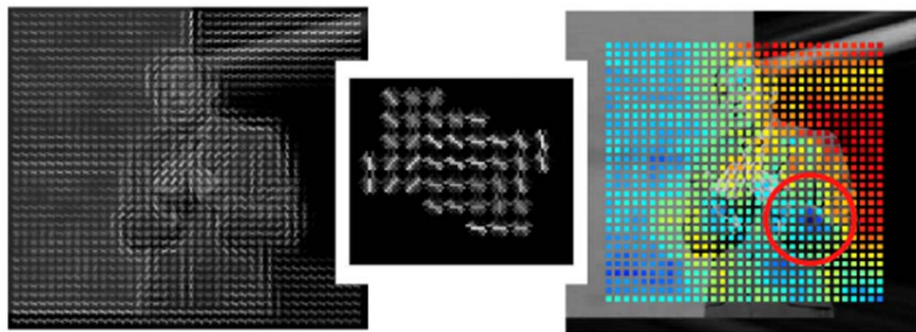


- Posterior of configuration

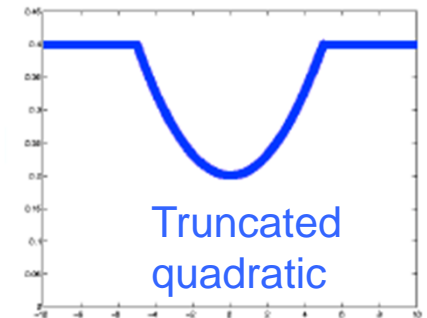
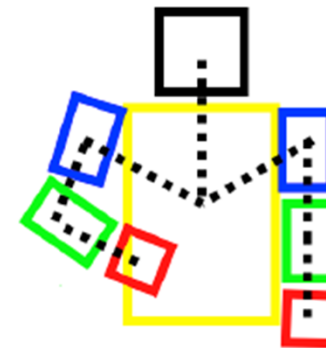
$$P(L | I) \propto \exp\left(\sum_i \Phi(l_i) + \sum_{(i,j) \in E} \Psi(l_i, l_j)\right)$$

- $\psi(l_i, l_j)$: relative position with geometric constraints
 - $\phi(l_i)$: local image evidence for a part in a particular location
 - If E is a tree, exact inference is efficiently performed by BP.
- Example of unary and pairwise terms
 - Unary term: appearance feature

- Pairwise term: kinematic layout

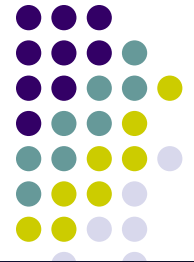


HOG of image HOG of lower arm template (learned) L2 Distance

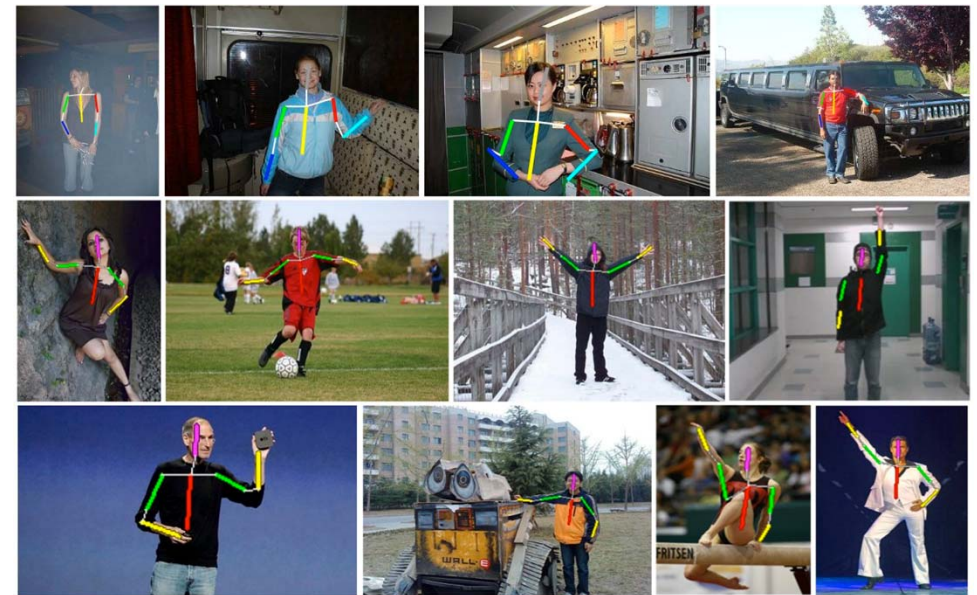
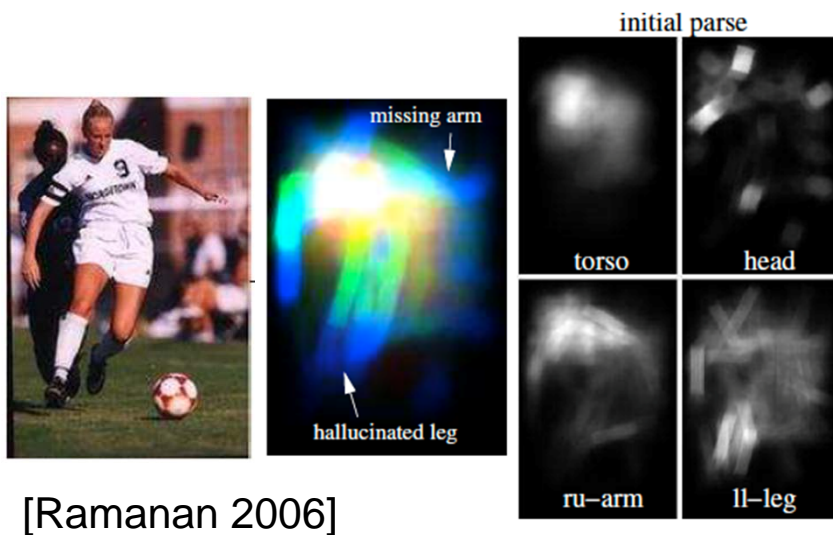


[Zisserman 2010]

Example of CRF – Results of Body Pose Estimation



- Examples of results



[Ferrari et al. 2008]

- Datasets and codes are available.
 - <http://www.ics.uci.edu/~dramanan/papers/parse/>
 - http://www.robots.ox.ac.uk/~vgg/research/pose_estimation/



Summary

- Conditional Random Fields are partially directed discriminative models
- They overcome the label bias problem of MEMMs by using a global normalizer
- Inference for 1-D chain CRFs is exact
 - Same as Max-product or Viterbi decoding
- Learning also is exact
 - globally optimum parameters can be learned
 - Requires using sum-product or forward-backward algorithm
- CRFs involving arbitrary graph structure are intractable in general
 - E.g.: Grid CRFs
 - Inference and learning require approximation techniques
 - MCMC sampling
 - Variational methods
 - Loopy BP