

1 Parameterizations for Gaussians

There are two common parameterizations for Gaussians, the moment parameterization and the natural parameterization.

The Moment Parameterization has the form

$$
\mathcal{N}(\mu, \Sigma) = p(\theta) = \frac{1}{z} \exp\left(-\frac{1}{2} (\theta - \mu) \Sigma^{-1} (\theta - \mu)\right)
$$
 (1)

The Natural Parameterization is

$$
\tilde{\mathcal{N}}(J, P) = \tilde{p}(\theta) = \frac{1}{z} \exp\left(J^T \theta - \frac{1}{2} \theta^T P \theta\right)
$$
\n(2)

The matrix P of the natural parameterization has a graphical model interpretation. If there is a non-zero entry for (z_1, z_2) , then there is a correspondence.

$$
P = \begin{bmatrix} X & X & 0 \\ X & X & X \\ 0 & X & X \\ \end{bmatrix} \longrightarrow \begin{bmatrix} z_1 \\ z_2 \end{bmatrix}
$$

Following the graphical model interpretation, P is in many cases highly structured. Consider for example the graphical model of a markov chain

$$
\overset{x_1}{\bigcirc} \overset{x_2}{\underbrace{\qquad \qquad } } \overset{x_3}{\underbrace{\qquad \qquad } } \overset{x_4}{\underbrace{\qquad \qquad } } \overset{x_5}{\underbrace{\qquad \qquad } } \overset{x_6}{\underbrace{\qquad \qquad } }
$$

This corresponds to a band structure in P:

$$
P = \begin{pmatrix} X & X & 0 & 0 & 0 \\ X & X & X & 0 & 0 \\ 0 & X & X & X & 0 \\ 0 & 0 & X & X & X \\ 0 & 0 & 0 & X & X \end{pmatrix}
$$
 (3)

Note: P^{-1} is, in general, not sparse! (this makes intuitve sense since $P^{-1} = \Sigma$ the covariance matrix, and the covariance of two states along the markov chain are not independent.)

2 Bayes Linear Regression Update

Scalar version of the likelihood field:

$$
p(y|x,\theta) = \mathcal{N}(\theta^T x_t, \sigma_t^2) = \frac{1}{z} \exp\left(\frac{-(\theta^T x - y)(\theta^T x - y)}{2\sigma^2}\right)
$$
(4)

(Don't worry about the weird notation of $\mathcal N$ as a function of σ^2 . This is an arbitrary definition)

2.1 Deriving the update rules

Apply Bayes' Rule to the probability of a weight vector θ given a datapoint D.

$$
p(\theta|D) = \frac{p(D|\theta)p(\theta)}{z} \tag{5}
$$

This results in the multiplication of two exponential functions. Adding the exponent of the prior to that of the likelihood yields

$$
-\frac{1}{2\sigma^2} \left(\theta^T x - y\right)^2 + J^T \theta - \frac{1}{2} \theta^T P \theta \tag{6}
$$

collecting terms to find updates J'_{θ} and P'_{θ} :

$$
= -\frac{1}{2\sigma^2} \left(\theta^T x x^T \theta - 2\theta^T x y + y^2 \right) + J^T \theta - \frac{1}{2} \theta^T P \theta \tag{7}
$$

$$
= \left(\frac{x^T y}{\sigma^2} + J^T\right)\theta - \frac{1}{2}\theta^T \left(\frac{x x^T}{\sigma^2} + P\right)\theta - \frac{y^2}{2\sigma^2} \tag{8}
$$

Since this all happens in the exponent of an exponential function, the constat y^2 -term can be shifted into the regularizing z. Thus, the update rules for J'_{θ} and P'_{θ} are

$$
J'_{\theta} = \frac{xy}{\sigma^2} + J \tag{9}
$$

$$
P'_{\theta} = \frac{xx^T}{\sigma^2} + P \tag{10}
$$

- 1. in a gaussian model, a new datapoint always lowers the variance this downgrading of the variance does not always make sense
- 2. if you believe there are outliers, this model won't work for you
- 3. the variance is not a function of y. The precision if only affected by input not output. This is a consequence of having the same σ (observation error) everywhere in space.

2.2 Transfer to moment parameterization

The update rules for the natural parameterization at timestep t are

$$
J + = \frac{y_t x_t}{\sigma^2} \tag{11}
$$

$$
P + = \frac{1}{\sigma^2} x x^T. \tag{12}
$$

Having no prior knowledge about the data, we choose standard initial conditions

$$
J_0 = 0 \tag{13}
$$

$$
P_0 = \mathbb{I},\tag{14}
$$

I being the identity matrix. Given the transfer rules to the moment parameterization

$$
\Sigma = P^{-1} \tag{15}
$$

$$
\mu = P^{-1}J \tag{16}
$$

the moment parameterization after N timesteps is then

$$
\Sigma_{\theta} = \left[\sum_{i=1}^{N} \frac{x_i x_i^T}{\sigma^2} + \mathbb{I} \right]^{-1} \tag{17}
$$

$$
\mu_{\theta} = \left(\sum_{i=1}^{N} \frac{x_i x_i^T}{\sigma^2} + \mathbb{I}\right) \sum_{t=1}^{N} \frac{y_t x_t}{\sigma^2} \tag{18}
$$

- 1. $\sum_{t=1}^{N} \frac{y_t x_t}{\sigma^2}$ is the gradient of Bayes online linear regression
- 2. this looks just like Newton's method
- 3. Computation time: $o(d^2)$ for update, $o(d^3)$ for mean (that can be reduced to $o(d^2)$ with tricks)

2.3 Making predictions

Given all data D up to timestep t and x_{t+1} , the probability of an observation \tilde{y}_{t+1} is

$$
p(\tilde{y}_{t+1}|x_{t+1}, D) = \int p(\tilde{y}_{t+1}|x_{t+1}, D, \theta) \cdot p(\theta|D) d\theta \qquad (19)
$$

$$
= \int p(\tilde{y}_{t+1}|x_{t+1}, \theta) \cdot p(\theta, D) d\theta \tag{20}
$$

To know $p(\tilde{y}_{t+1}|x_{t+1}, D)$, we only need y and σ^2 , because these parameters determine the gaussian.

2.4 Marginal and Conditional Distributions in different parameters

These computations are crucial for gaussian processes and Kalman filters.

2.4.1 Moment parameterization

Given:

$$
\mathcal{N}\left(\left[\begin{array}{c}\mu_1\\ \mu_2\end{array}\right], \left[\begin{array}{cc}\Sigma_{11} & \Sigma_{12}\\ \Sigma_{21} & \Sigma_{22}\end{array}\right]\right) \tag{21}
$$

Marginal: computing $p(x_2)$

$$
\mu_2^{\text{marg}} = \mu_2 \tag{22}
$$

$$
\Sigma_2^{\text{marg}} = \Sigma_{22} \tag{23}
$$

Conditional: computing $p(x_1|x_2)$

$$
\mu_{1|2} = \mu_1 + \Sigma_{12} \Sigma_{22}^{-1} (x_2 - \mu_2) \tag{24}
$$

$$
\Sigma_{1|2} = \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21}
$$
\n(25)

2.4.2 Natural parameterization

Given:

$$
\mathcal{N}\left(\left[\begin{array}{c}J_1\\J_2\end{array}\right],\left[\begin{array}{cc}P_{11}&P_{12}\\P_{21}&P_{22}\end{array}\right]\right) \tag{26}
$$

Marginal: computing $p(x_2)$

$$
J_2^{\text{marg}} = J_2 - P_{21} P_{11}^{-1} J_1 \tag{27}
$$

$$
P_1^{\text{marg}} = P_{12} - P_{21} P_{11}^{-1} P_{12} \tag{28}
$$

Conditional: computing $p(x_1|x_2)$

$$
J_{1|2} = J_1 - P_{12}x_2 \tag{29}
$$

$$
P_{1|2} = P_{11} \tag{30}
$$