

Regularized Linear Regression

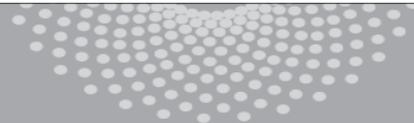
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Machine Learning 10-701

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Mean square error regression

Optimal predictor: $f^* = \arg \min_f \mathbb{E}[(f(X) - Y)^2]$

Empirical Minimizer: $\hat{f}_n = \arg \min_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n (f(X_i) - Y_i)^2$

Class of predictors

- \mathcal{F} - Class of Linear functions
- Class of Polynomial functions
- Class of nonlinear functions

Least Square solution satisfies Normal Equations

$$\underbrace{(\mathbf{A}^T \mathbf{A})}_{p \times p} \underbrace{\hat{\beta}}_{p \times 1} = \underbrace{\mathbf{A}^T \mathbf{Y}}_{p \times 1}$$

If $(\mathbf{A}^T \mathbf{A})$ is invertible,

1) If dimension p not too large, analytical solution:

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y} \quad \hat{f}_n^L(X) = X \hat{\beta}$$

2) If dimension p is large, computing inverse is expensive $O(p^3)$

Gradient descent since objective is convex ($\mathbf{A}^T \mathbf{A} \succeq 0$)

$$\begin{aligned} \beta^{t+1} &= \beta^t - \frac{\alpha}{2} \frac{\partial J(\beta)}{\partial \beta} \Big|_t \\ &= \beta^t - \alpha \mathbf{A}^T (\mathbf{A} \beta^t - \mathbf{Y}) \end{aligned}$$

Linear regression solution satisfies Normal Equations

$$\underbrace{(\mathbf{A}^T \mathbf{A})}_{p \times p} \underbrace{\hat{\boldsymbol{\beta}}}_{p \times 1} = \underbrace{\mathbf{A}^T \mathbf{Y}}_{p \times 1}$$

When is $(\mathbf{A}^T \mathbf{A})$ invertible ?

Recall: Full rank matrices are invertible. What is rank of $(\mathbf{A}^T \mathbf{A})$?

Linear regression solution satisfies Normal Equations

$$\underbrace{(\mathbf{A}^T \mathbf{A})}_{p \times p} \underbrace{\hat{\boldsymbol{\beta}}}_{p \times 1} = \underbrace{\mathbf{A}^T \mathbf{Y}}_{p \times 1}$$

When is $(\mathbf{A}^T \mathbf{A})$ invertible?

Recall: **Full rank matrices are invertible.** What is rank of $(\mathbf{A}^T \mathbf{A})$?

If $\mathbf{A} = \mathbf{U} \mathbf{S} \mathbf{V}^T$, then
 S - $r \times r$

normal equations $\underbrace{(\mathbf{S} \mathbf{V}^T)}_{r \times p} \underbrace{\hat{\boldsymbol{\beta}}}_{p \times 1} = \underbrace{(\mathbf{U}^T \mathbf{Y})}_{r \times 1}$

r equations in p unknowns. Under-determined if $r < p$, hence no unique solution.

Regularized Least Squares

What if $(\mathbf{A}^T \mathbf{A})$ is not invertible ?

r equations , p unknowns – underdetermined system of linear equations
many feasible solutions

Need to constrain solution further

e.g. bias solution to “small” values of β (small changes in input don’t translate to large changes in output)

$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \lambda \|\beta\|_2^2$$

Ridge Regression
(l2 penalty)

$$= \arg \min_{\beta} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) + \lambda \|\beta\|_2^2 \quad \lambda \geq 0$$

$$\hat{\beta}_{\text{MAP}} = (\mathbf{A}^T \mathbf{A} + \lambda \mathbf{I})^{-1} \mathbf{A}^T \mathbf{Y}$$

Is $(\mathbf{A}^T \mathbf{A} + \lambda \mathbf{I})$ invertible ?

Ridge Regression

$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \lambda \|\beta\|_2^2$$

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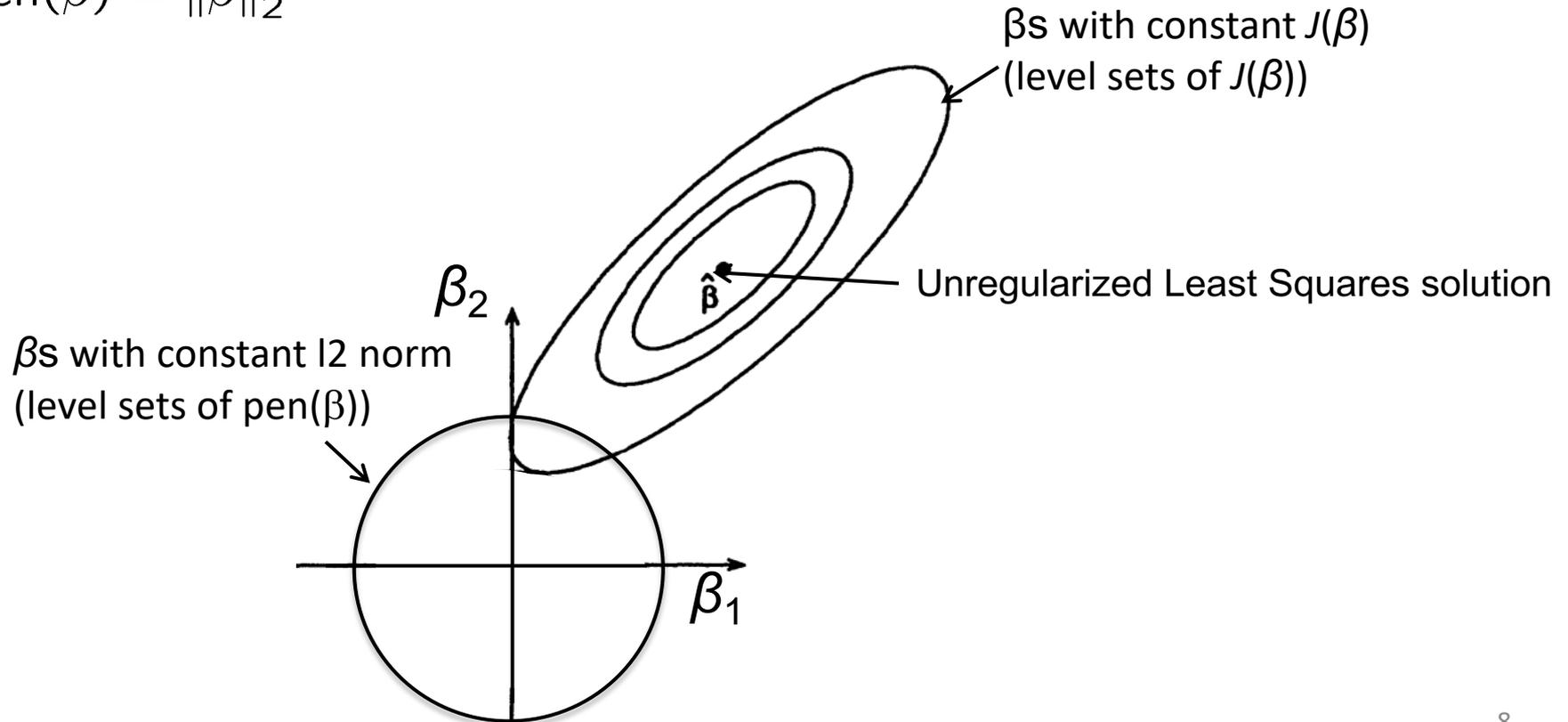
Is $(\mathbf{A}^T \mathbf{A} + \lambda \mathbf{I})$ invertible ?

Understanding regularized Least Squares

$$\min_{\beta} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) + \lambda \text{pen}(\beta) = \min_{\beta} J(\beta) + \lambda \text{pen}(\beta)$$

Ridge Regression:

$$\text{pen}(\beta) = \|\beta\|_2^2$$



Regularized Least Squares

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Ridge Regression
(l2 penalty)

$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \lambda \|\beta\|_1$$

Lasso
(l1 penalty)

$$\lambda \geq 0$$

Many β can be zero – many inputs are irrelevant to prediction in high-dimensional settings (typically intercept term not penalized)

Regularized Least Squares

What if $(\mathbf{A}^T \mathbf{A})$ is not invertible ?

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Lasso
(l1 penalty)

$$\lambda \geq 0$$

No closed form solution, but can optimize using sub-gradient descent (packages available)

Ridge Regression vs Lasso

$$\min_{\beta} (\mathbf{A}\beta - \mathbf{Y})^T (\mathbf{A}\beta - \mathbf{Y}) + \lambda \text{pen}(\beta) = \min_{\beta} J(\beta) + \lambda \text{pen}(\beta)$$

Ridge Regression:

$$\text{pen}(\beta) = \|\beta\|_2^2$$

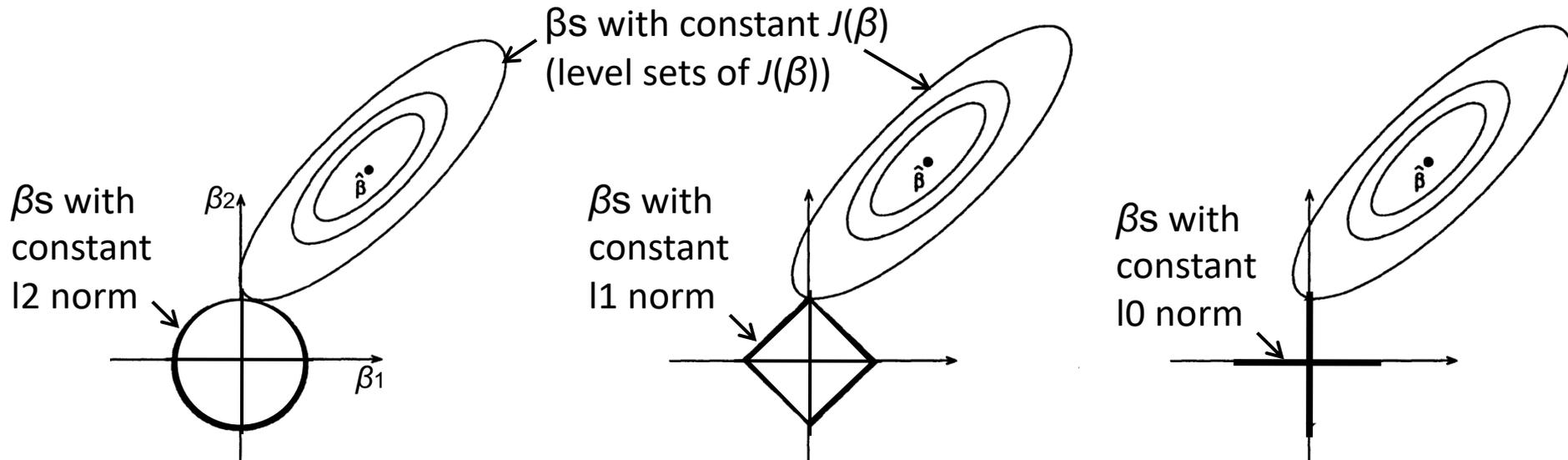
Lasso:

$$\text{pen}(\beta) = \|\beta\|_1$$

Ideally l0 penalty,

but optimization

becomes non-convex



Lasso (l1 penalty) results in sparse solutions – vector with more zero coordinates
Good for high-dimensional problems – don't have to store all coordinates, interpretable solution!

Matlab example

```
clear all  
close all
```

```
n = 80; % datapoints  
p = 100; % features  
k = 10; % non-zero features
```

```
rng(20);  
X = randn(n,p);  
weights = zeros(p,1);  
weights(1:k) = randn(k,1)+10;  
noise = randn(n,1) * 0.5;  
Y = X*weights + noise;
```

```
Xtest = randn(n,p);  
noise = randn(n,1) * 0.5;  
Ytest = Xtest*weights + noise;
```

```
lassoWeights = lasso(X,Y,'Lambda',1,  
'Alpha', 1.0);  
Ylasso = Xtest*lassoWeights;  
norm(Ytest-Ylasso)
```

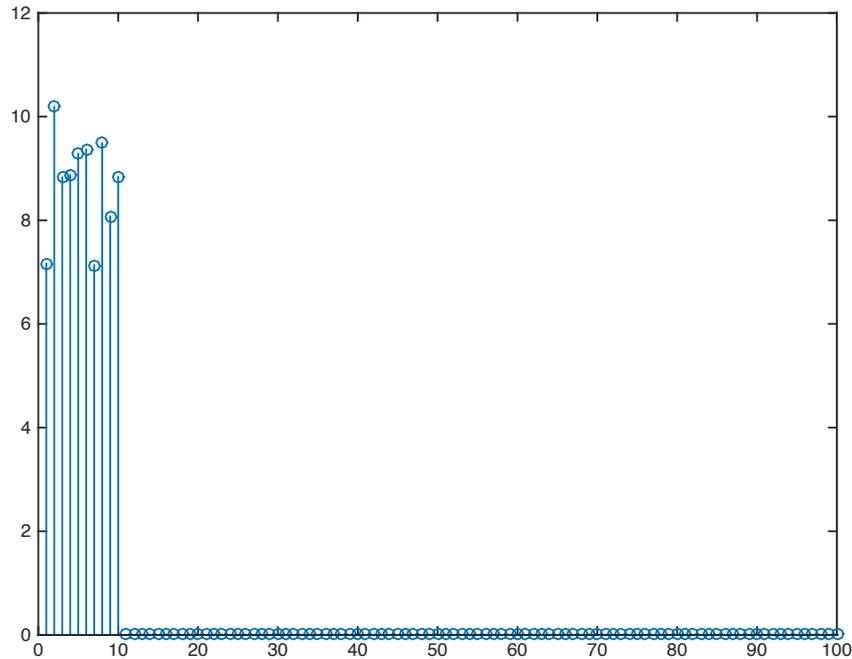
```
ridgeWeights = lasso(X,Y,'Lambda',1,  
'Alpha', 0.0001);  
Yridge = Xtest*ridgeWeights;  
norm(Ytest-Yridge)
```

```
stem(lassoWeights)  
pause  
stem(ridgeWeights)
```

Matlab example

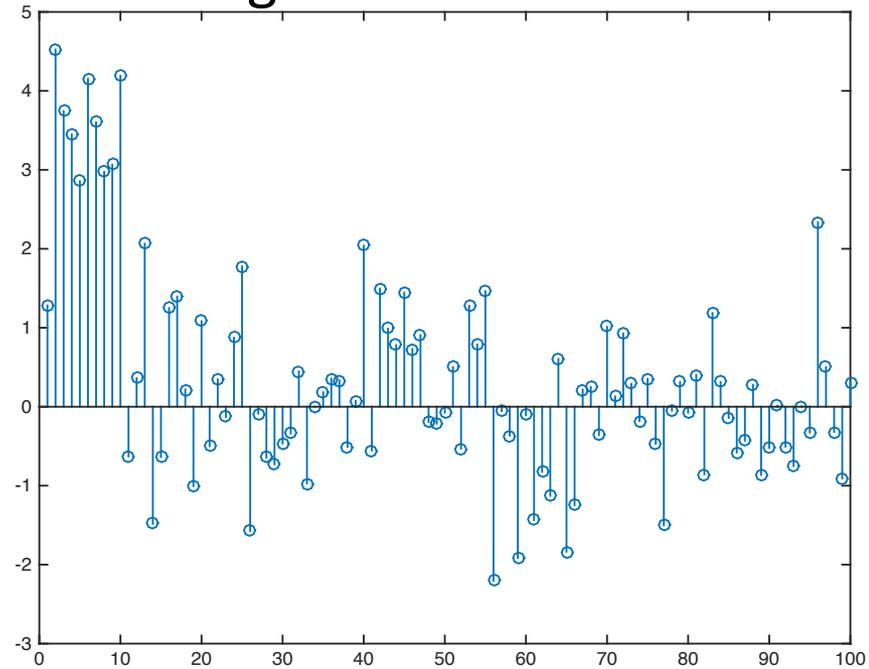
Test MSE = 33.7997

Lasso Coefficients



Test MSE = 185.9948

Ridge Coefficients



Least Squares and M(C)LE

Intuition: Signal plus (zero-mean) Noise model

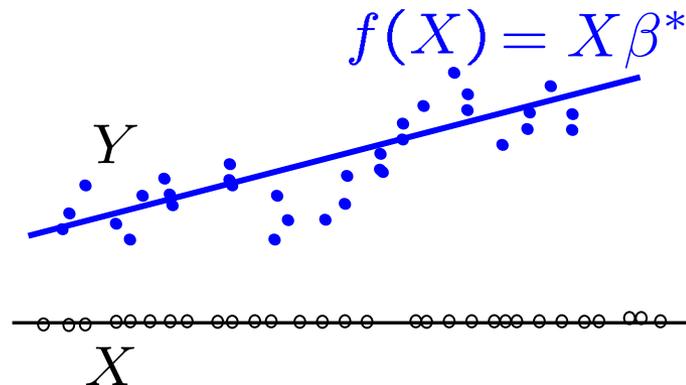
$$Y = f^*(X) + \epsilon = X\beta^* + \epsilon$$

$$\epsilon \sim \mathcal{N}(0, \sigma^2 \mathbf{I}) \quad Y \sim \mathcal{N}(X\beta^*, \sigma^2 \mathbf{I})$$

$$\hat{\beta}_{\text{MLE}} = \arg \max_{\beta} \underbrace{\log p(\{Y_i\}_{i=1}^n | \beta, \sigma^2, \{X_i\}_{i=1}^n)}_{\text{Conditional log likelihood}}$$

Conditional log likelihood

$$= \arg \min_{\beta} \sum_{i=1}^n (X_i \beta - Y_i)^2 = \hat{\beta}$$



Least Square Estimate is same as Maximum Conditional Likelihood Estimate under a Gaussian model !

Regularized Least Squares and M(C)AP

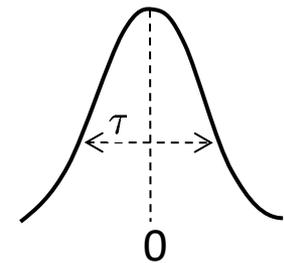
What if $(\mathbf{A}^T \mathbf{A})$ is not invertible ?

$$\hat{\beta}_{\text{MAP}} = \arg \max_{\beta} \underbrace{\log p(\{Y_i\}_{i=1}^n | \beta, \sigma^2, \{X_i\}_{i=1}^n)}_{\text{Conditional log likelihood}} + \underbrace{\log p(\beta)}_{\text{log prior}}$$

1) Gaussian Prior

$$\beta \sim \mathcal{N}(0, \tau^2 \mathbf{I})$$

$$p(\beta) \propto e^{-\beta^T \beta / 2\tau^2}$$



$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \lambda \|\beta\|_2^2$$

constant(σ^2, τ^2)

Ridge Regression

Prior belief that β is Gaussian with zero-mean biases solution to “small” β

Regularized Least Squares and M(C)AP

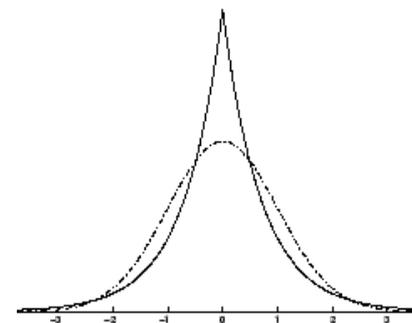
What if $(\mathbf{A}^T \mathbf{A})$ is not invertible ?

$$\hat{\beta}_{\text{MAP}} = \arg \max_{\beta} \underbrace{\log p(\{Y_i\}_{i=1}^n | \beta, \sigma^2, \{X_i\}_{i=1}^n)}_{\text{Conditional log likelihood}} + \underbrace{\log p(\beta)}_{\text{log prior}}$$

II) Laplace Prior

$$\beta_i \stackrel{iid}{\sim} \text{Laplace}(0, t)$$

$$p(\beta_i) \propto e^{-|\beta_i|/t}$$



$$\hat{\beta}_{\text{MAP}} = \arg \min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \lambda \|\beta\|_1$$

\downarrow
 constant(σ^2, t)

Lasso

Prior belief that β is Laplace with zero-mean biases solution to “sparse” β

Polynomial Regression

degree m

Univariate (1-dim) $f(X) = \beta_0 + \beta_1 X + \beta_2 X^2 + \dots + \beta_m X^m = \mathbf{X}\beta$
case:

where $\mathbf{X} = [1 \ X \ X^2 \ \dots \ X^m]$, $\beta = [\beta_1 \ \dots \ \beta_m]^T$

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y}$$

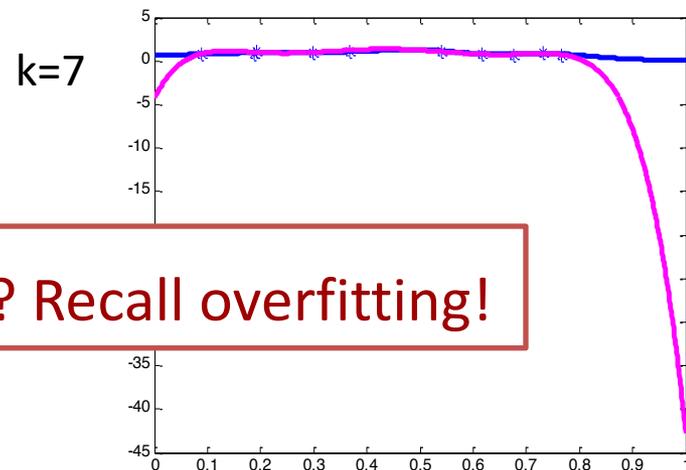
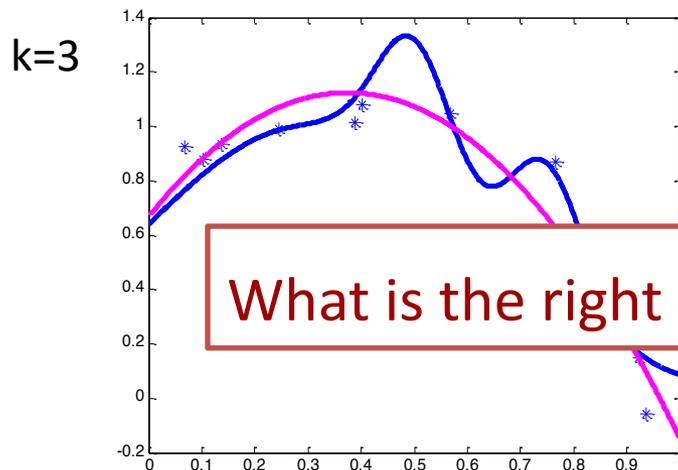
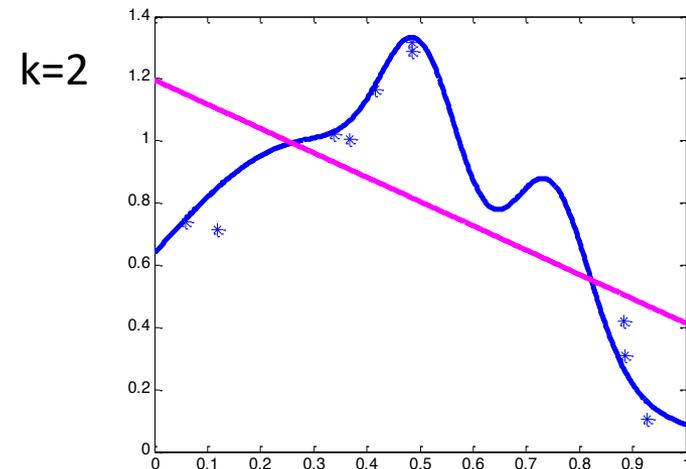
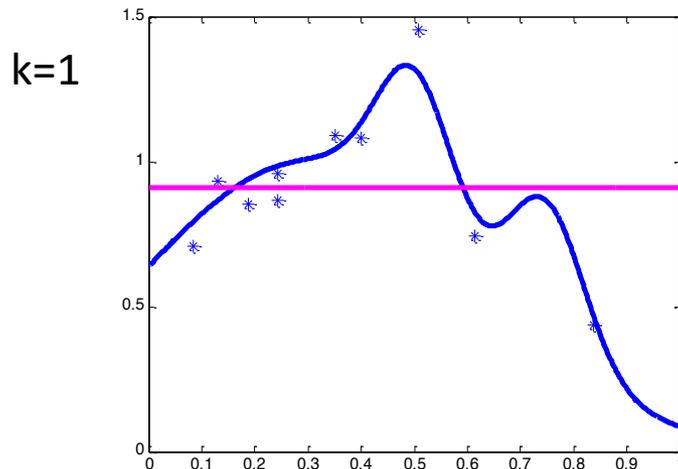
$$\hat{f}_n(X) = \mathbf{X}\hat{\beta}$$

$$\text{where } \mathbf{A} = \begin{bmatrix} 1 & X_1 & X_1^2 & \dots & X_1^m \\ \vdots & & & \ddots & \vdots \\ 1 & X_n & X_n^2 & \dots & X_n^m \end{bmatrix}$$

Multivariate (p-dim) $f(X) = \beta_0 + \beta_1 X^{(1)} + \beta_2 X^{(2)} + \dots + \beta_p X^{(p)}$
case:
 $+ \sum_{i=1}^p \sum_{j=1}^p \beta_{ij} X^{(i)} X^{(j)} + \sum_{i=1}^p \sum_{j=1}^p \sum_{k=1}^p X^{(i)} X^{(j)} X^{(k)}$
 $+ \dots$ terms up to degree m

Polynomial Regression

Polynomial of order k , equivalently of degree up to $k-1$



What is the right order? Recall overfitting!

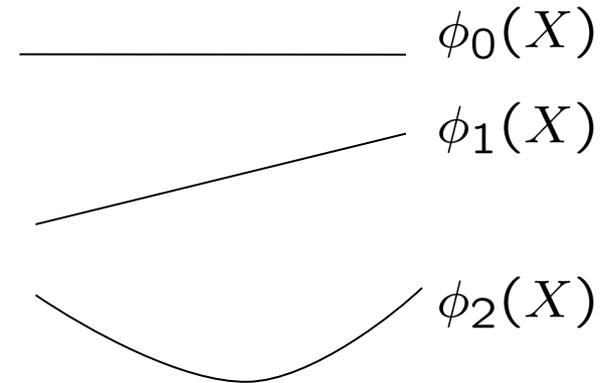
Regression with nonlinear features

$$f(X) = \sum_{j=0}^m \beta_j X^j = \sum_{j=0}^m \beta_j \phi_j(X)$$

Weight of
each feature



Nonlinear
features



In general, use any nonlinear features

e.g. e^X , $\log X$, $1/X$, $\sin(X)$, ...

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A})^{-1} \mathbf{A}^T \mathbf{Y}$$

$$\mathbf{A} = \begin{bmatrix} \phi_0(X_1) & \phi_1(X_1) & \dots & \phi_m(X_1) \\ \vdots & & \ddots & \vdots \\ \phi_0(X_n) & \phi_1(X_n) & \dots & \phi_m(X_n) \end{bmatrix}$$

$$\hat{f}_n(X) = \mathbf{X} \hat{\beta}$$

$$\mathbf{X} = [\phi_0(X) \ \phi_1(X) \ \dots \ \phi_m(X)]$$

Poll

- The maximum likelihood estimate of model parameter α for the random variable $y \sim N(\alpha x_1 x_2^3, \sigma^2)$, where x_1 and x_2 are random variables, can be learned using linear regression on n iid samples of (x_1, x_2, y)
 - True
 - False

Can we kernelize linear regression?

Linear (Ridge) regression

$$\min_{\beta} \sum_{i=1}^n (Y_i - X_i\beta)^2 + \lambda \|\beta\|_2^2 \quad \hat{\beta} = (\mathbf{A}^T \mathbf{A} + \lambda \mathbf{I})^{-1} \mathbf{A}^T \mathbf{Y}$$

Recall

$$\mathbf{A} = \begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix} = \begin{bmatrix} X_1^{(1)} & \dots & X_1^{(p)} \\ \vdots & \ddots & \vdots \\ X_n^{(1)} & \dots & X_n^{(p)} \end{bmatrix}$$

Hence $\mathbf{A}^T \mathbf{A}$ is a $p \times p$ matrix whose entries denote the (sample) correlation between the features

NOT inner products between the data points – the inner product matrix would be $\mathbf{A} \mathbf{A}^T$ which is $n \times n$ (also known as Gram matrix)

Using dual formulation, we can write the solution in terms of $\mathbf{A} \mathbf{A}^T$

Ridge regression

$$\min_{\beta} \sum_{i=1}^n (Y_i - X_i \beta)^2 + \lambda \|\beta\|_2^2$$

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A} + \lambda \mathbf{I})^{-1} \mathbf{A}^T \mathbf{Y}$$

Similarity with SVMs

Primal problem:

$$\begin{aligned} \min_{\beta, z_i} \quad & \sum_{i=1}^n z_i^2 + \lambda \|\beta\|_2^2 \\ \text{s.t.} \quad & z_i = Y_i - X_i \beta \end{aligned}$$

SVM Primal problem:

$$\begin{aligned} \min_{w, \xi_i} \quad & C \sum_{i=1}^n \xi_i + \frac{1}{2} \|w\|_2^2 \\ \text{s.t.} \quad & \xi_i = \max(1 - Y_i X_i w, 0) \end{aligned}$$

Lagrangian:

$$\sum_{i=1}^n z_i^2 + \lambda \|\beta\|_2^2 + \sum_{i=1}^n \alpha_i (z_i - Y_i + X_i \beta)$$

α_i – Lagrange parameter, one per training point

Kernelized ridge regression

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A} + \lambda \mathbf{I})^{-1} \mathbf{A}^T \mathbf{Y}$$

Using dual, can re-write solution as:

$$\hat{\beta} = \mathbf{A}^T (\mathbf{A} \mathbf{A}^T + \lambda \mathbf{I})^{-1} \mathbf{Y}$$

How does this help?

- Only need to invert $n \times n$ matrix (instead of $p \times p$ or $m \times m$)
- More importantly, kernel trick!

$\mathbf{A} \mathbf{A}^T$ involves only inner products between the training points
BUT still have an extra \mathbf{A}^T

$$\begin{aligned} \text{Recall the predicted label is } \hat{f}_n(X) &= \mathbf{X} \hat{\beta} \\ &= \mathbf{X} \mathbf{A}^T (\mathbf{A} \mathbf{A}^T + \lambda \mathbf{I})^{-1} \mathbf{Y} \end{aligned}$$

$\mathbf{X} \mathbf{A}^T$ contains inner products between test point \mathbf{X} and training points!

Kernelized ridge regression

$$\hat{\beta} = (\mathbf{A}^T \mathbf{A} + \lambda \mathbf{I})^{-1} \mathbf{A}^T \mathbf{Y}$$

$$\hat{f}_n(X) = \mathbf{X} \hat{\beta}$$

Using dual, can re-write solution as:

$$\hat{\beta} = \mathbf{A}^T (\mathbf{A} \mathbf{A}^T + \lambda \mathbf{I})^{-1} \mathbf{Y}$$

How does this help?

- Only need to invert $n \times n$ matrix (instead of $p \times p$ or $m \times m$)
- More importantly, kernel trick!

$$\hat{f}_n(X) = \mathbf{K}_X (\mathbf{K} + \lambda \mathbf{I})^{-1} \mathbf{Y} \quad \text{where} \quad \begin{aligned} \mathbf{K}_X(i) &= \phi(X) \cdot \phi(X_i) \\ \mathbf{K}(i, j) &= \phi(X_i) \cdot \phi(X_j) \end{aligned}$$

Work with kernels, never need to write out the high-dim vectors

Ridge Regression with (implicit) nonlinear features $\phi(X)$! $f(X) = \phi(X) \beta$