

Introduction to Gaussian Processes

Barnabás Póczos
University of Alberta

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Gaussian Processes for Machine Learning



Carl Edward Rasmussen and Christopher K. I. Williams

<http://www.gaussianprocess.org/>

Some of these slides in the intro are taken from D. Lizotte, R. Parr, C. Guestrin

Roadmap

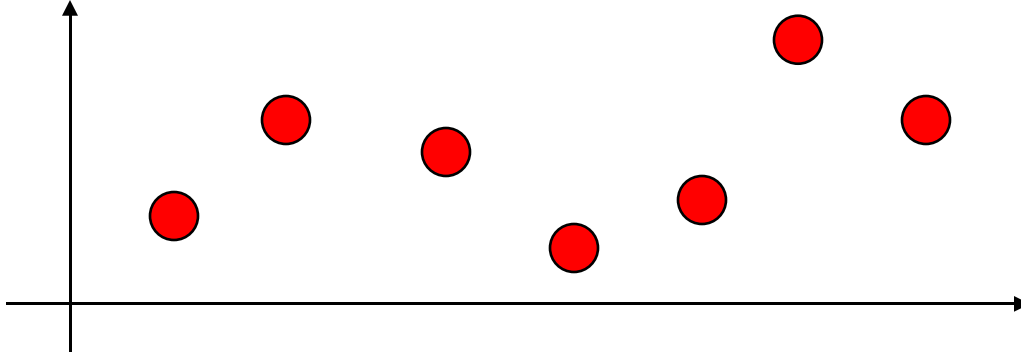
- Introduction
- Ridge Regression
- Gaussian Processes
 - Weight space view
 - Bayesian Ridge Regression + Kernel trick
 - Function space view
 - Prior distribution over functions
+ calculation posterior distributions

Roadmap

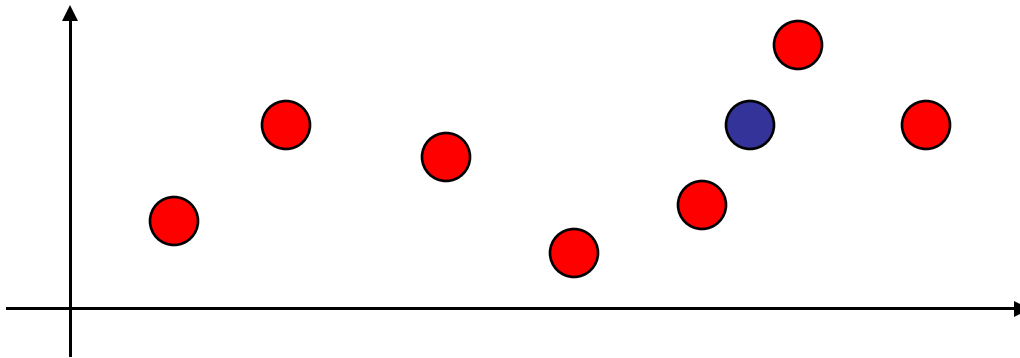
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Why GPs?

- Here are some data points! What function did they come from?



- I have *no idea*.
- Oh. Okay. Uh, you think this point is likely in the function, too?



- I still have *no idea*.

Why GPs?

- You can't get anywhere without making some assumptions
- GPs are a nice way of expressing this 'prior on functions' idea.
- Like a more 'complete' view of least-squares regression
- Can do a bunch of cool stuff
 - **Regression**
 - Classification
 - Optimization

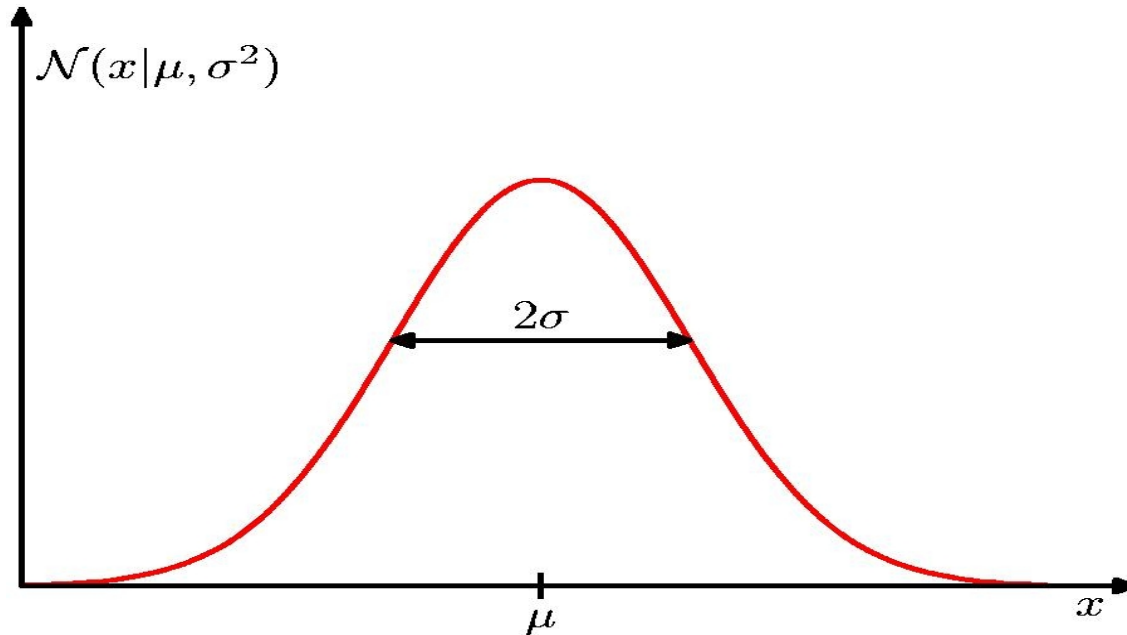
Why GPs?

Under certain assumptions GPs can answer the following questions

- Here are some data points, and here's how I rank the **likelihood of functions**.
- Here's where the function will **most likely be**.
(expected function)
- Here are some **examples** of what it might look like.
(sampling from the posterior distribution)
- Here is a prediction of what you'll see if you evaluate your function at x' , **with confidence**

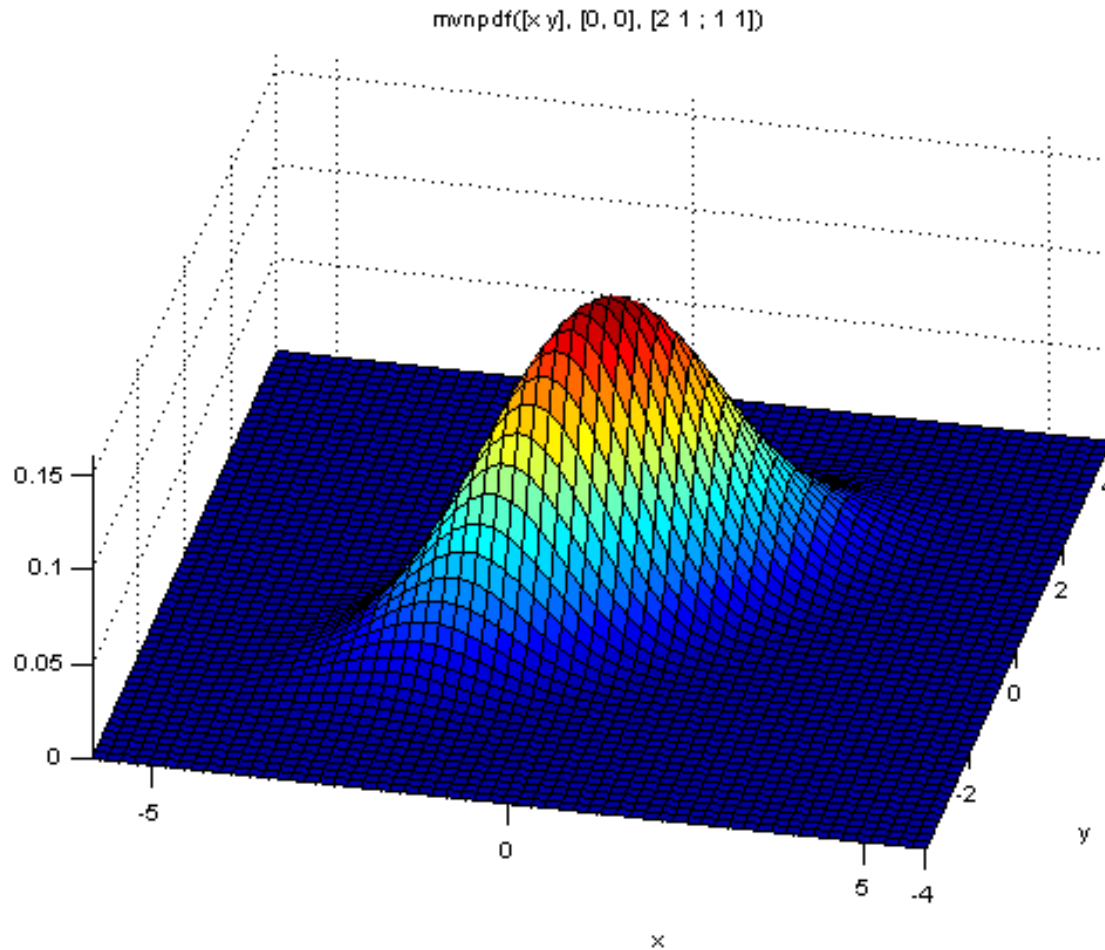
1D Gaussian Distribution

- Parameters
 - Mean, μ
 - Variance, σ^2



$$P(x \mid \mu, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{-(x-\mu)^2}{2\sigma^2}}$$

Multivariate Gaussian



$$P(\mathbf{x} \mid \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{2\pi|\boldsymbol{\Sigma}|}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right\}$$

The Multivariate Gaussian

- A 2-dimensional Gaussian is defined by

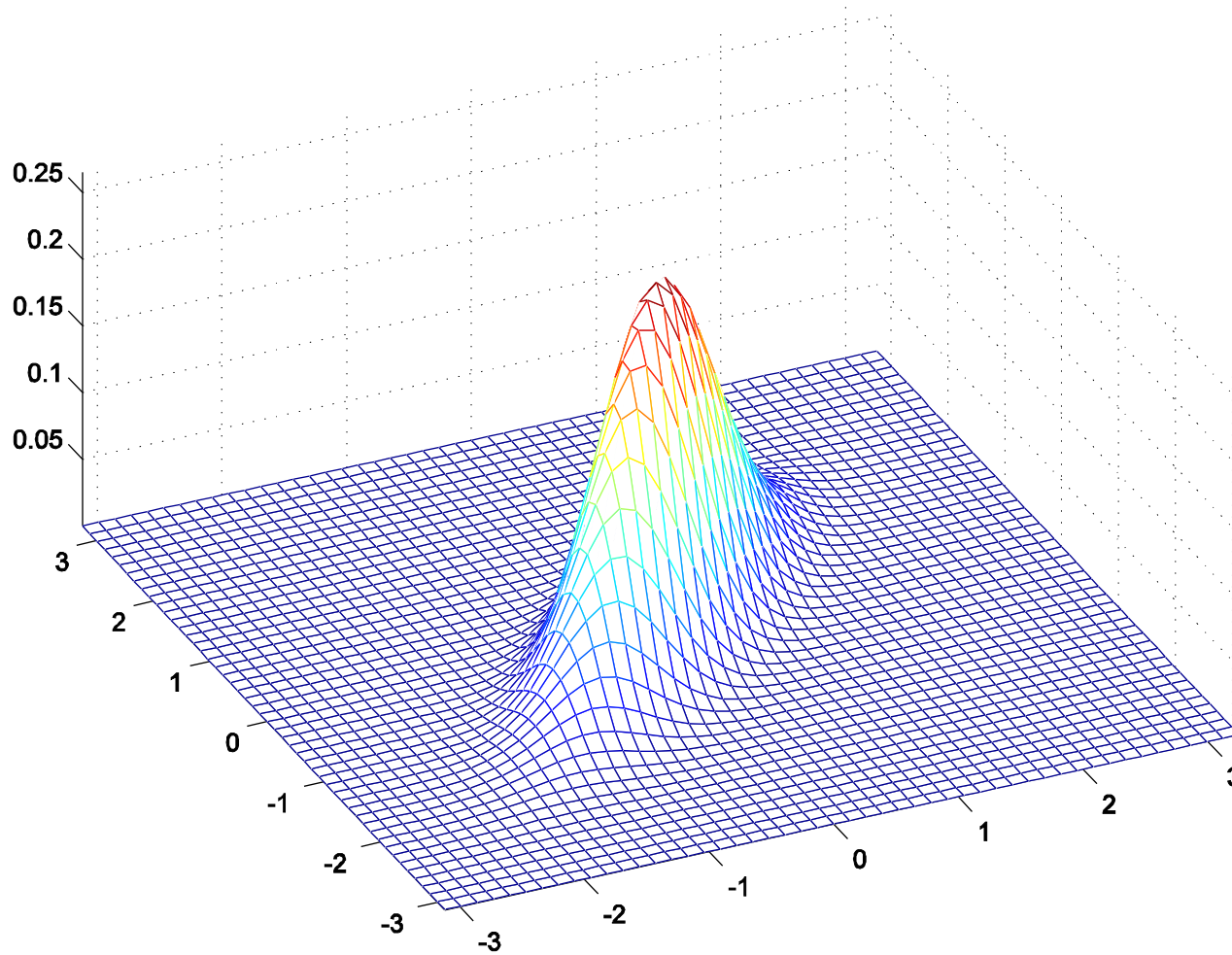
- a mean vector $\mu = [\mu_1, \mu_2]$

- a covariance matrix: $\Sigma = \begin{bmatrix} \sigma_{1,1}^2 & \sigma_{2,1}^2 \\ \sigma_{1,2}^2 & \sigma_{2,2}^2 \end{bmatrix}$

where $\sigma_{i,j}^2 = E[(x_i - \mu_i)(x_j - \mu_j)]$
is (co)variance

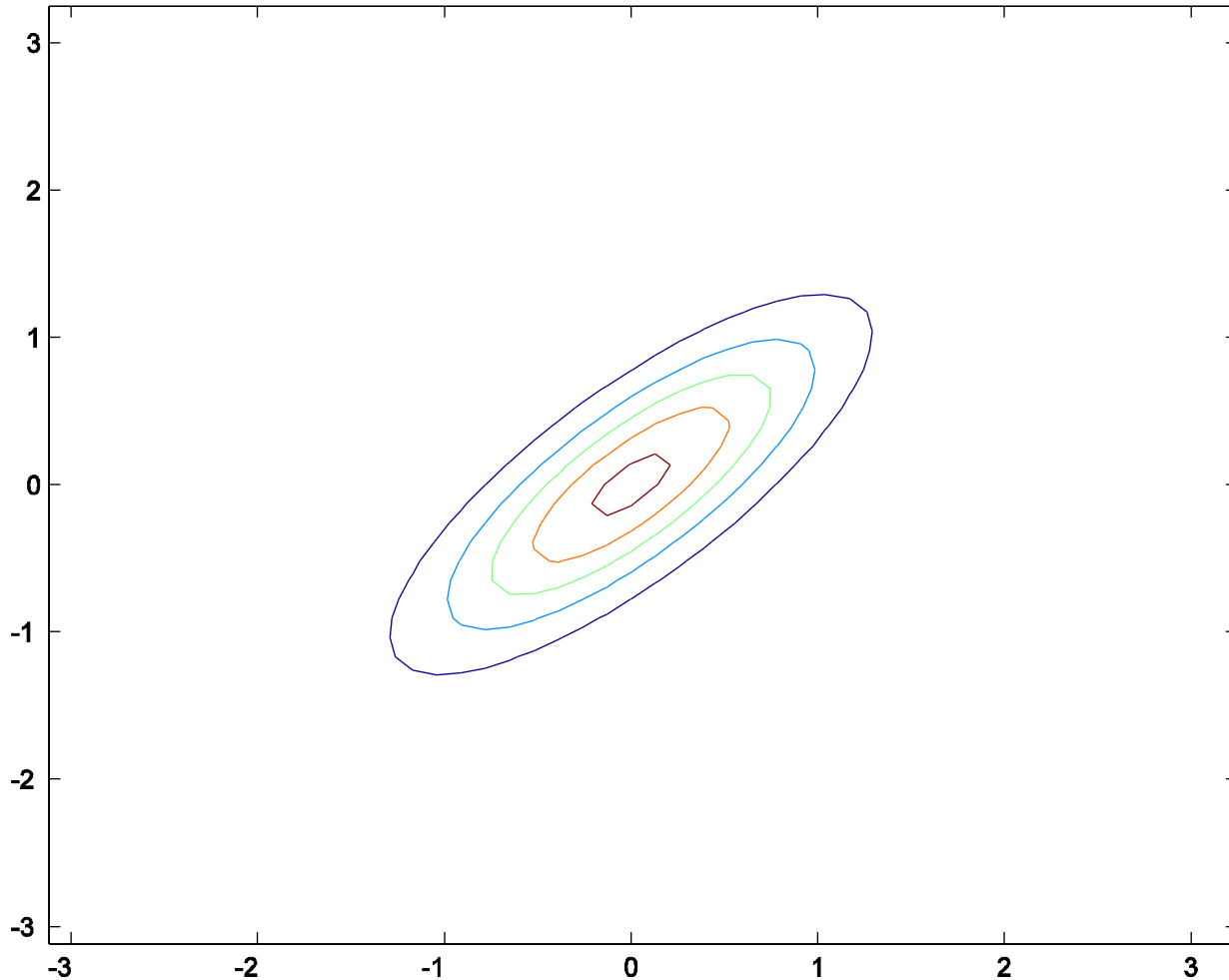
- Note: Σ is symmetric,
“positive semi-definite”: $\forall x: x^T \Sigma x \geq 0$

Multivariate Gaussian examples



$$\mu = (0,0) \quad \Sigma = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$

Multivariate Gaussian examples



$$\mu = (0,0) \quad \Sigma = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}$$

Useful Properties of Gaussians

- Marginals of Gaussians are Gaussian

- Given: $x = (x_a, x_b), \mu = (\mu_a, \mu_b)$

$$\Sigma = \begin{pmatrix} \Sigma_{aa} & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_{bb} \end{pmatrix}$$

- Marginal Distribution:

$$p(x_a) = N(x_a | \mu_a, \Sigma_{aa})$$

- (Marginalize by ignoring)

Useful Properties of Gaussians

- Conditionals of Gaussians are Gaussian
- Notation:

$$\Sigma = \begin{pmatrix} \Sigma_{aa} & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_{bb} \end{pmatrix} \quad \Lambda = \Sigma^{-1} = \begin{pmatrix} \Lambda_{aa} & \Lambda_{ab} \\ \Lambda_{ba} & \Lambda_{bb} \end{pmatrix}$$

- Conditional Distribution:

$$p(x_a | x_b) = N(x_a | \mu_{a|b}, \Lambda_{aa}^{-1})$$

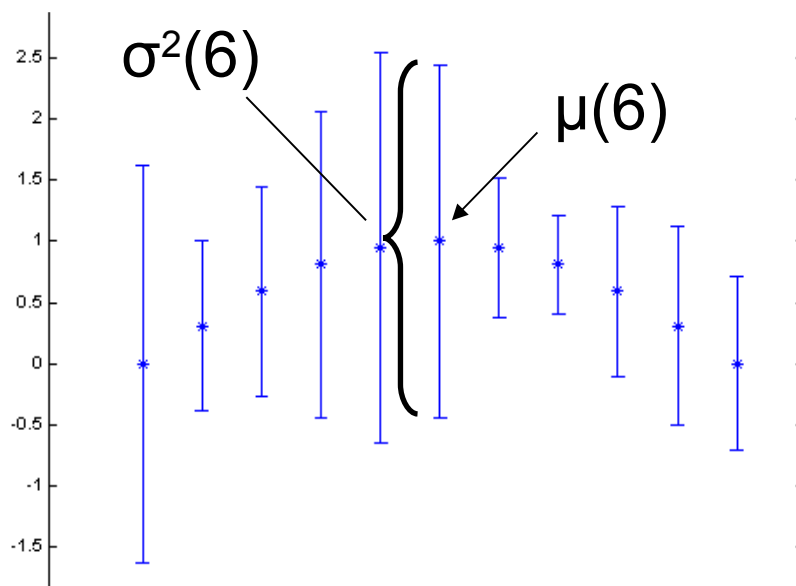
$$\mu_{a|b} = \mu_a - \Lambda_{aa}^{-1} \Lambda_{ab} (\mathbf{x}_b - \mu_b) = \mu_a - \Sigma_{ab} \Sigma_{bb}^{-1} (\mathbf{x}_b - \mu_b)$$

$$\Lambda_{aa}^{-1} = \Sigma_{aa}^{-1} - \Sigma_{ab} \Sigma_{bb}^{-1} \Sigma_{ba}$$

Higher Dimensions

- **Visualizing > 3 dimensions is... difficult**
- Means and marginals are practical
 - But then we don't see correlations
- Marginals are Gaussian, e.g., $f(6) \sim N(\mu(6), \sigma^2(6))$

Visualizing a multivariate Gaussian \mathbf{f} :

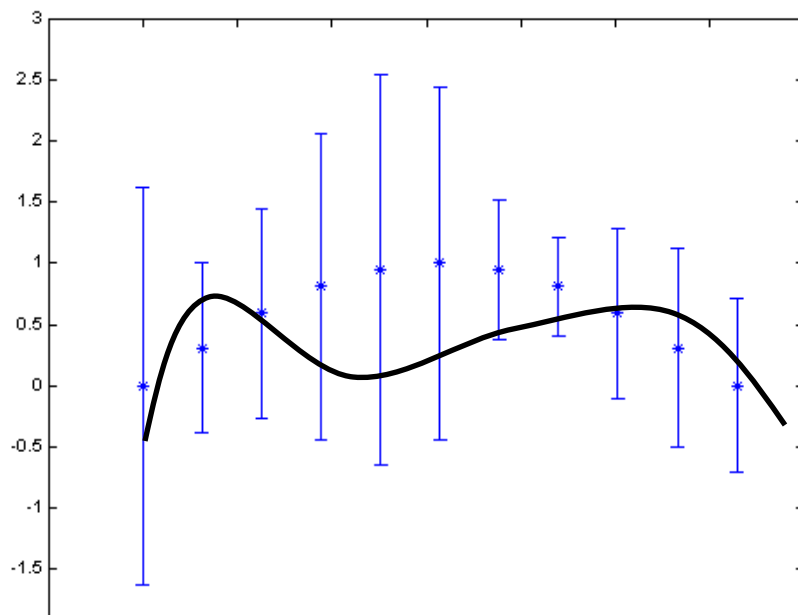


Yet Higher Dimensions

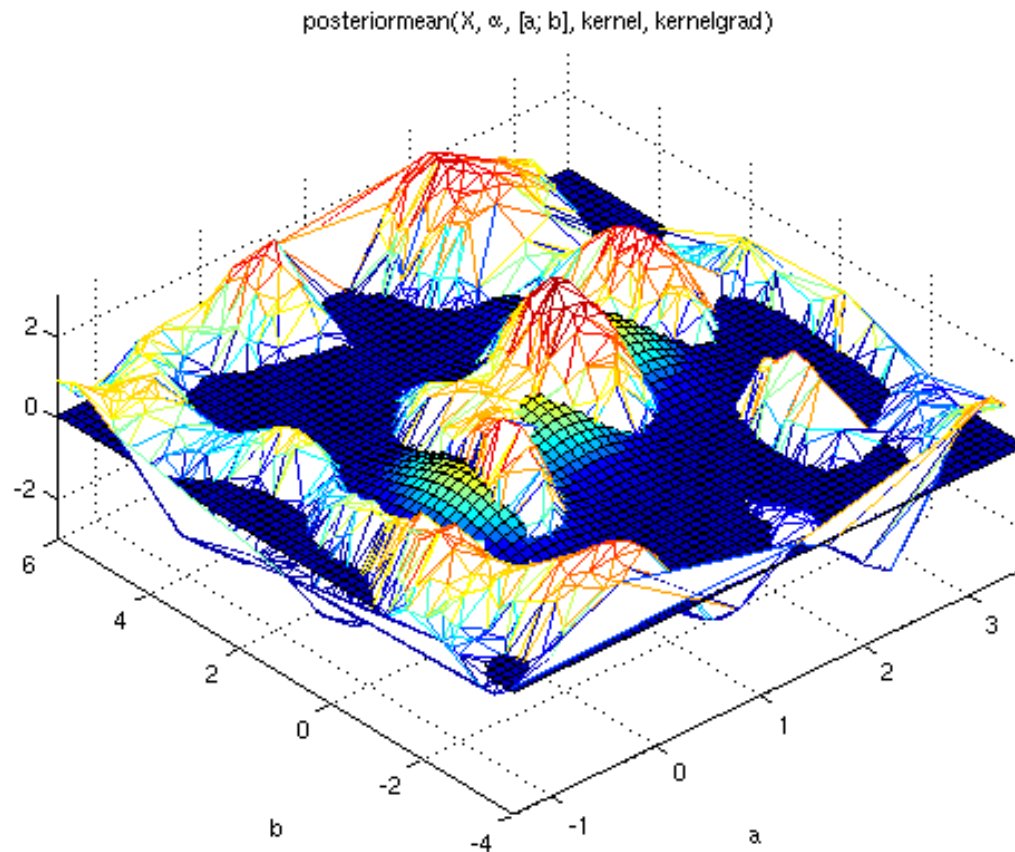
Why stop there?

- We indexed before with \mathbb{Z} , why not with \mathbb{R} ?
- Need functions $\mu(x), k(x, z), \forall x, z \in \mathbb{R}$
- x and z are indexes over the random variables
- f is now an uncountably infinite dimensional vector

Don't panic: It's just a function



Getting Ridiculous



Why stop there?

- We indexed before with \mathbb{R} , why not with \mathbb{R}^D ?
- Need functions $\mu(\mathbf{x}), k(\mathbf{x}, \mathbf{z}), \forall \mathbf{x}, \mathbf{z} \in \mathbb{R}^D$

Gaussian Process

Definition:

- Probability distribution *indexed by* an arbitrary set
- Each element gets a Gaussian distribution over the reals with mean $\mu(x)$
- These distributions are dependent/correlated as defined by $k(x,z)$
- Any finite subset of indices defines a multivariate Gaussian distribution
 - Crazy mathematical statistics and measure theory ensures this

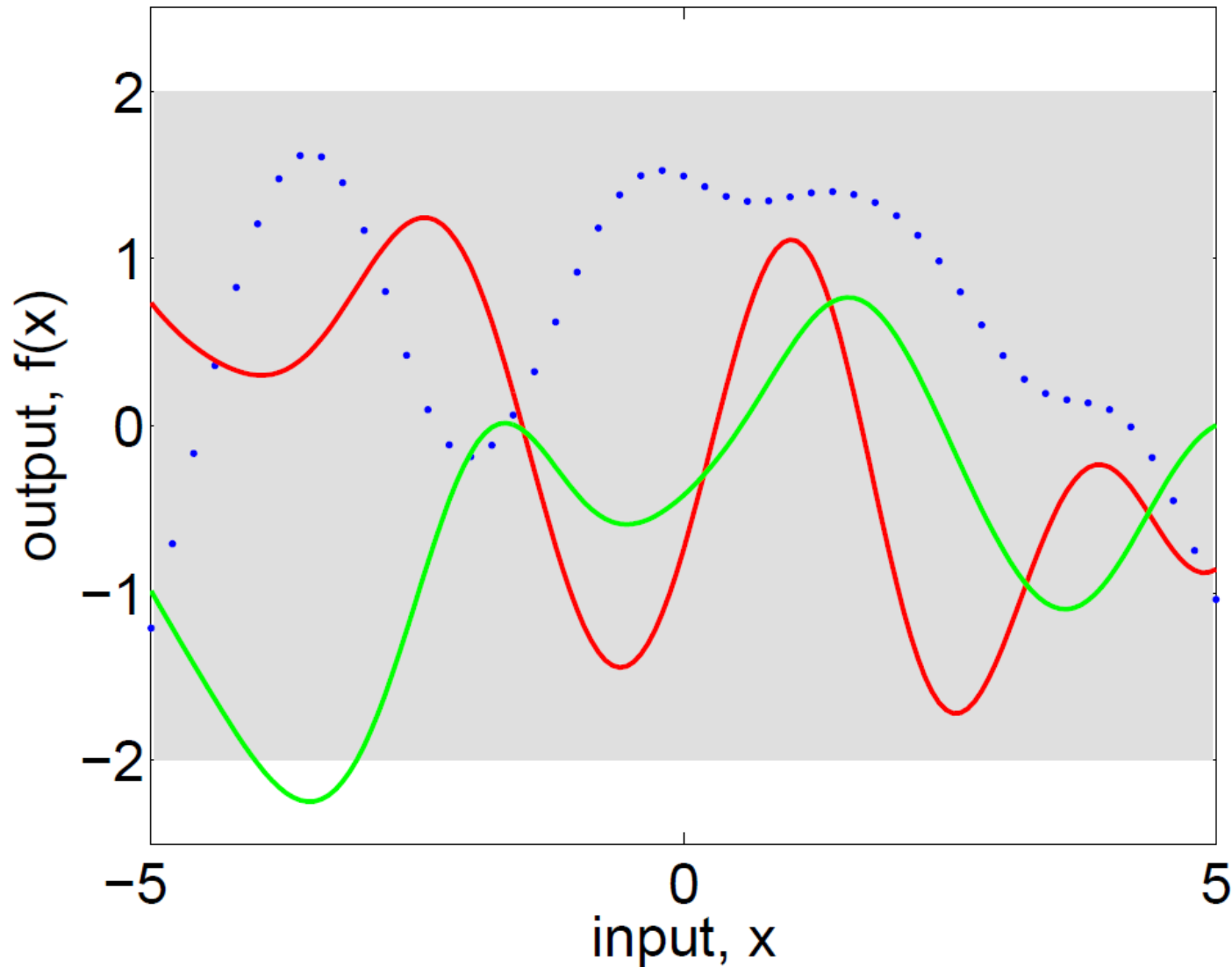
Gaussian Process

- Distribution over *functions*
- Domain of the functions (index set) can be pretty much whatever
 - Reals
 - Real vectors
 - Graphs
 - Strings
 - Sets
 - ...
- Most interesting structure is in $k(x,z)$, the 'kernel.'

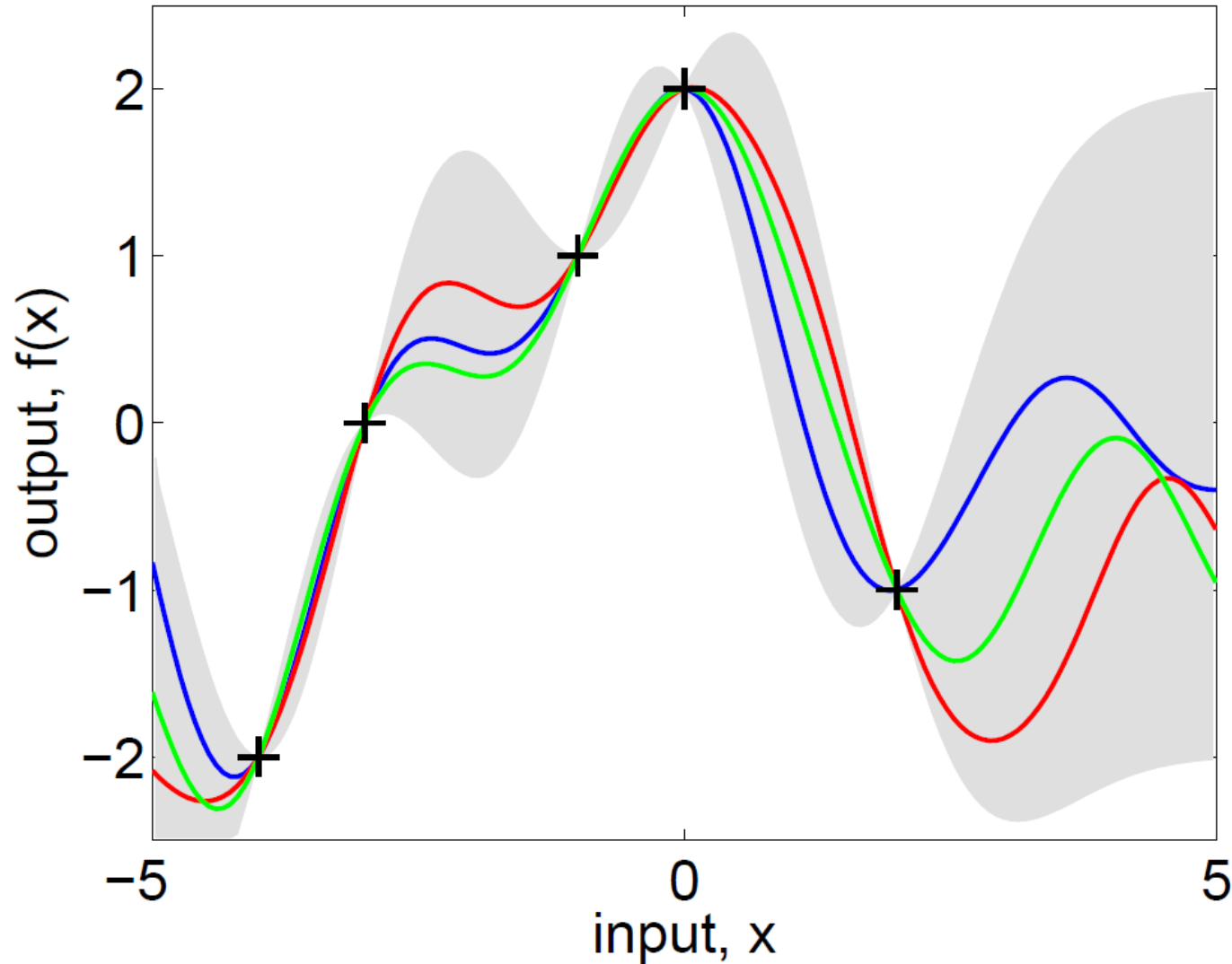
Bayesian Updates for GPs

- How do Bayesians use a Gaussian Process?
 - Start with GP prior
 - Get some data
 - Compute a posterior
- Ask interesting questions about the posterior

Samples from the prior distribution



Samples from the posterior distribution



Picture is taken from Rasmussen and Williams²²

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Ridge Regression

Training set: $D = \{(x_i, y_i) | i = 1, \dots, n\}$

Linear regression: $f(x) = \langle \mathbf{w}, \phi(x) \rangle$

Ridge regression:

$$\hat{\mathbf{w}} = \arg \min_{\mathbf{w} \in \mathcal{K}} \sum_{i=1}^m (y_i - \underbrace{\langle \phi(x_i), \mathbf{w} \rangle}_{\mathbf{x}_i})^2 + \lambda \|\mathbf{w}\|^2$$

**The Gaussian Process is a Bayesian Generalization
of the Ridge regression**

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Weight Space View

Bayesian ridge regression in feature space + Kernel trick to carry out computations

Training set: $D = \{(\mathbf{x}_i, y_i) | i = 1, \dots, n\}$

$$\left. \begin{aligned} X &= \begin{bmatrix} \mathbf{x}_1 & \dots & \mathbf{x}_n \end{bmatrix} \in \mathbb{R}^{D \times n}, \text{ design matrix} \\ y &= \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} \in \mathbb{R}^n \end{aligned} \right\} \text{The training data}$$

Bayesian Analysis of Linear Regression with Gaussian noise

$$f(\mathbf{x}) = \mathbf{x}^T \mathbf{w} \in \mathbb{R}, \mathbf{x}, \mathbf{w} \in \mathbb{R}^D$$

$$y = f(\mathbf{x}) + \epsilon = \mathbf{x}^T \mathbf{w} + \epsilon \in \mathbb{R}$$

$$\epsilon \sim \mathcal{N}(0, \sigma^2) \in \mathbb{R}$$

(*Homoscedastic* noise, the same for all \mathbf{x})

Let us calculate the likelihood:

$$P(\mathbf{y} | X, \mathbf{w}) = \prod_{i=1}^n P(y_i | \mathbf{x}_i^T \mathbf{w})$$

and then put $\mathbf{w} \sim \mathcal{N}_{\mathbf{w}}(0, \Sigma_p)$ prior over parameters \mathbf{w} .

Bayesian Analysis of Linear Regression with Gaussian noise

The likelihood:

$$\begin{aligned} P(\mathbf{y} | X, \mathbf{w}) &= \prod_{i=1}^n P(y_i | \mathbf{x}_i^T \mathbf{w}) \\ &= \prod_{i=1}^n \mathcal{N}_{y_i}(\mathbf{x}_i^T \mathbf{w}, \sigma^2) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left[\frac{-(y_i - \mathbf{x}_i^T \mathbf{w})^2}{2\sigma^2} \right] \\ &= \frac{1}{(2\pi\sigma^2)^{n/2}} \exp \left[\frac{-1}{2\sigma^2} \|\mathbf{y} - X^T \mathbf{w}\|^2 \right] \\ &= \mathcal{N}_{\mathbf{y}}(X^T \mathbf{w}, \sigma^2 \mathbf{I}_n) \end{aligned}$$

Bayesian Analysis of Linear Regression with Gaussian noise

The prior:

$$\mathbf{w} \sim \mathcal{N}_{\mathbf{w}}(0, \Sigma_p)$$

Now, we can calculate the posterior:

$$\begin{aligned} P(\mathbf{w}|X, \mathbf{y}) &= \frac{P(\mathbf{y}|X, \mathbf{w})P(\mathbf{w})}{P(\mathbf{y}|X)} \\ &= \frac{P(\mathbf{y}|X, \mathbf{w})P(\mathbf{w})}{\int P(\mathbf{y}|X, \mathbf{w})d\mathbf{w}} \\ &= \frac{\mathcal{N}_{\mathbf{y}}(X^T \mathbf{w}, \sigma^2 \mathbf{I}_n) \mathcal{N}_{\mathbf{w}}(0, \Sigma_p)}{\int \mathcal{N}_{\mathbf{y}}(X^T \mathbf{w}, \sigma^2 \mathbf{I}_n) \mathcal{N}_{\mathbf{w}}(0, \Sigma_p) d\mathbf{w}} \\ &\sim \mathcal{N}_{\mathbf{y}}(X^T \mathbf{w}, \sigma^2 \mathbf{I}_n) \mathcal{N}_{\mathbf{w}}(0, \Sigma_p) \end{aligned}$$

Bayesian Analysis of Linear Regression with Gaussian noise

Ridge Regression

$$\begin{aligned} P(\mathbf{w}|X, \mathbf{y}) &\sim \mathcal{N}_{\mathbf{y}}(X^T \mathbf{w}, \sigma^2 \mathbf{I}_n) \mathcal{N}_{\mathbf{w}}(0, \Sigma_p) \\ &\sim \exp\left\{\frac{-1}{2\sigma^2}(\mathbf{y} - X^T \mathbf{w})^T (\mathbf{y} - X^T \mathbf{w})\right\} \exp\left\{\frac{-1}{2}\mathbf{w}^T \Sigma_p^{-1} \mathbf{w}\right\} \\ &\sim \exp\left\{\frac{-1}{2}(\mathbf{w} - \bar{\mathbf{w}})^T \underbrace{\left(\frac{1}{\sigma^2} X X^T + \Sigma_p^{-1}\right)}_A (\mathbf{w} - \bar{\mathbf{w}})\right\} \\ &\sim \boxed{N_{\mathbf{w}}(-\bar{\mathbf{w}}, A^{-1})} \quad \text{After "completing the square"}$$

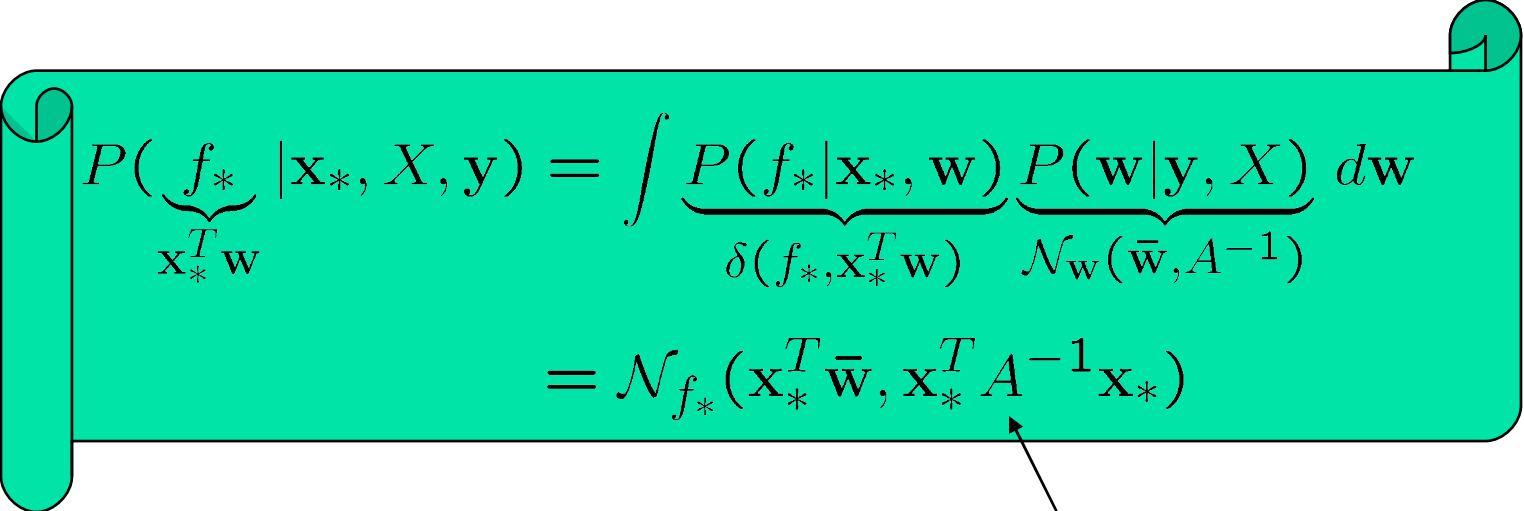
where $\bar{\mathbf{w}} \doteq \sigma^{-2} \underbrace{\left(\sigma^{-2} X X^T + \Sigma_p^{-1}\right)^{-1}}_{A^{-1} \in \mathbb{R}^{D \times D}} X \mathbf{y} \in \mathbb{R}^D$ **MAP estimation**

$$\boxed{A \doteq \left(\sigma^{-2} X X^T + \Sigma_p^{-1}\right) \in \mathbb{R}^{D \times D}}$$

Bayesian Analysis of Linear Regression with Gaussian noise

We want to use $P(\mathbf{w}|X, \mathbf{y}) = N_{\mathbf{w}}(\bar{\mathbf{w}}, A^{-1})$ posterior for predicting f in a test point \mathbf{x}_* .

$$f_* \doteq f(\mathbf{x}_*) \quad f(\mathbf{x}) = \mathbf{x}^T \mathbf{w} \in \mathbb{R}, \quad \mathbf{x}, \mathbf{w} \in \mathbb{R}^D$$
$$y = f(\mathbf{x}) + \epsilon = \mathbf{x}^T \mathbf{w} + \epsilon \in \mathbb{R}$$


$$P(\underbrace{f_*}_{\mathbf{x}_*^T \mathbf{w}} | \mathbf{x}_*, X, \mathbf{y}) = \int \underbrace{P(f_* | \mathbf{x}_*, \mathbf{w})}_{\delta(f_*, \mathbf{x}_*^T \mathbf{w})} \underbrace{P(\mathbf{w} | \mathbf{y}, X)}_{N_{\mathbf{w}}(\bar{\mathbf{w}}, A^{-1})} d\mathbf{w}$$
$$= \mathcal{N}_{f_*}(\mathbf{x}_*^T \bar{\mathbf{w}}, \mathbf{x}_*^T A^{-1} \mathbf{x}_*)$$

This posterior covariance matrix doesn't depend on the observations \mathbf{y} ,
A strange property of Gaussian Processes $\mathbf{y}^T = [y_1, \dots, y_n]$

Projections of Inputs into Feature Space

The reviewed Bayesian linear regression suffers from
limited expressiveness



To overcome the problem \Rightarrow
go to a feature space and do linear regression there

a., **explicit** features $\phi(\mathbf{x}) = [x_1, x_1x_2^2, x_1 - x_2, \dots]^T$

b., **implicit** features (kernels) $k(\vec{x}, \vec{y}) = \exp(-\|\vec{x} - \vec{y}\|^2)$

Explicit Features

$$\phi(\mathbf{x}) = [x_1, x_1x_2^2, x_1 - x_2, \dots]^T \in \mathbb{R}^N$$

$$\phi(X) = \begin{bmatrix} \phi(\mathbf{x}_1) & \phi(\mathbf{x}_2) & \dots & \phi(\mathbf{x}_n) \end{bmatrix} \in \mathbb{R}^{N \times n}$$

$$f(\mathbf{x}) = \phi(\mathbf{x})^T \mathbf{w} \in \mathbb{R}, \quad \phi(\mathbf{x}), \mathbf{w} \in \mathbb{R}^N$$

$$y = f(\mathbf{x}) + \epsilon = \phi(\mathbf{x})^T \mathbf{w} + \epsilon \in \mathbb{R}$$

Linear regression in the feature space

Explicit Features

The predictive distribution after feature map:

$$P(\underbrace{f_*}_{\phi(\mathbf{x}_*)^T \mathbf{w}} | \mathbf{x}_*, X, \mathbf{y}) = N_{f_*} \left(\phi(x_*)^T \bar{\mathbf{w}}, \phi(x_*)^T A^{-1} \phi(x_*) \right)$$

where $\bar{\mathbf{w}} \doteq \sigma^{-2} \underbrace{\left(\sigma^{-2} \phi(X) \phi(X)^T + \Sigma_p^{-1} \right)^{-1}}_{A^{-1} \in \mathbb{R}^{N \times N}} \phi(X) \mathbf{y} \in \mathbb{R}^D$

$$A \doteq \left(\sigma^{-2} \phi(X) \phi(X)^T + \Sigma_p^{-1} \right) \in \mathbb{R}^{N \times N}$$

Explicit Features

Shorthands:

$$\begin{aligned}\phi_* &\doteq \phi(\mathbf{x}_*) \in \mathbb{R}^N & N &= \text{dim of feature space} \\ \phi &\doteq \phi(X) = \begin{bmatrix} \phi(\mathbf{x}_1) & \phi(\mathbf{x}_2) & \dots & \phi(\mathbf{x}_n) \end{bmatrix} \in \mathbb{R}^{N \times n} \\ A &\doteq \left(\sigma^{-2} \phi \phi^T + \Sigma_p^{-1} \right) \in \mathbb{R}^{N \times N} \\ \bar{\mathbf{w}} &\doteq \underbrace{\sigma^{-2} \left(\sigma^{-2} \phi \phi^T + \Sigma_p^{-1} \right)^{-1}}_{A^{-1} \in \mathbb{R}^{N \times N}} \phi \mathbf{y} \in \mathbb{R}^N\end{aligned}$$

The predictive distribution after feature map:

$$P(\underbrace{f_*}_{\phi_*^T \bar{\mathbf{w}}} | \mathbf{x}_*, X, \mathbf{y}) = \mathcal{N}_{f_*} \left(\phi_*^T \bar{\mathbf{w}}, \phi_*^T A^{-1} \phi_* \right)$$

Explicit Features

The predictive distribution after feature map:

$$P(\underbrace{f_*}_{\phi_*^T \bar{\mathbf{w}}} | \mathbf{x}_*, X, \mathbf{y}) = N_{f_*} \left(\phi_*^T \bar{\mathbf{w}}, \phi_*^T A^{-1} \phi_* \right) \quad (*)$$

$$= N_{f_*} \left(\sigma^{-2} \phi_*^T \left[\sigma^{-2} \phi \phi^T + \Sigma_p^{-1} \right]^{-1} \phi \mathbf{y}, \phi_*^T \left[\sigma^{-2} \phi \phi^T + \Sigma_p^{-1} \right]^{-1} \phi_* \right)$$

A problem with (*) is that it needs an $N \times N$ matrix inversion...

Let $K \doteq \phi^T \Sigma_p \phi \in \mathbb{R}^{n \times n}$

(*) can be rewritten: $P(f_* | \mathbf{x}_*, X, \mathbf{y}) =$

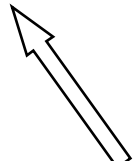
$$\underbrace{N_{f_*}}_{\mathbb{R}^{n \times n}} \left(\underbrace{(\phi_*^T \Sigma_p \phi)}_{\mathbb{R}^{n \times n}} \underbrace{(K + \sigma^2 \mathbf{I}_n)^{-1}}_{\mathbb{R}^{n \times n}} \mathbf{y}, \underbrace{(\phi_*^T \Sigma_p \phi_*)}_{\mathbb{R}^{n \times n}} - \underbrace{(\phi_*^T \Sigma_p \phi)}_{\mathbb{R}^{n \times n}} \underbrace{(K + \sigma^2 \mathbf{I}_n)^{-1}}_{\mathbb{R}^{n \times n}} \underbrace{(\phi^T \Sigma_p \phi_*)}_{\mathbb{R}^{n \times n}} \right)$$

Proofs

- **Mean expression.** We need:

$$\sigma^{-2} \phi_*^T \underbrace{[\sigma^{-2} \phi \phi^T + \Sigma_p^{-1}]^{-1}}_{A^{-1}} \phi y = (\phi_*^T \underbrace{\Sigma_p \phi}_{\sigma^{-2} A^{-1} \phi}) (K + \sigma^2 \mathbf{I}_n)^{-1} y$$

Lemma:

$$\sigma^{-2} \phi (K + \sigma^2 \mathbf{I}_n) = \sigma^{-2} \phi (\phi^T \Sigma_p \phi + \sigma^2 \mathbf{I}_n) = A \Sigma_p \phi$$


- **Variance expression.** We need:

$$\phi_*^T [\sigma^{-2} \phi \phi^T + \Sigma_p^{-1}]^{-1} \phi_* = (\phi_*^T \Sigma_p \phi_*) - (\phi_*^T \Sigma_p \phi) (K + \sigma^2 \mathbf{I}_n)^{-1} (\phi^T \Sigma_p \phi_*)$$

Matrix inversion Lemma:

$$(\underbrace{U}_{\phi} \underbrace{W}_{\sigma^{-2}} \underbrace{V^T}_{\phi^T} + \underbrace{Z}_{\Sigma_p^{-1}})^{-1} = Z^{-1} - Z^{-1} U (W^{-1} + \underbrace{V^T Z^{-1} U}_K)^{-1} V^T Z^{-1}$$

From Explicit to Implicit Features

$$P(f_* | \mathbf{x}_*, X, \mathbf{y}) =$$

$$N_{f_*} \left(\underbrace{(\phi_*^T \Sigma_p \phi)}_{\mathbb{R}^{n \times n}} \underbrace{(K + \sigma^2 \mathbf{I}_n)^{-1}}_{\mathbb{R}^{n \times n}} \mathbf{y}, \underbrace{(\phi_*^T \Sigma_p \phi_*)}_{\mathbb{R}^{n \times n}} - \underbrace{(\phi_*^T \Sigma_p \phi)}_{\mathbb{R}^{n \times n}} \underbrace{(K + \sigma^2 \mathbf{I}_n)^{-1}}_{\mathbb{R}^{n \times n}} \underbrace{(\phi^T \Sigma_p \phi_*)}_{\mathbb{R}^{n \times n}} \right)$$

We have to work only with $n \times n$ matrices, and not by $N \times N$

The feature space always enters in the form of:

$$(\phi_*^T \Sigma_p \phi_*), (\phi_*^T \Sigma_p \phi), (\phi^T \Sigma_p \phi), (\in \mathbb{R}^{n \times n} \text{ matrices})$$

$$\text{Let } k(x, \tilde{x}) \doteq \phi(x)^T \Sigma_p \phi(\tilde{x})$$

Lemma:

$k(x, \tilde{x})$ is an inner product in the feature space: $\psi(x) \doteq \Sigma_p^{1/2} \phi(x)$ 38

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Function Space View

- An alternative way to get the previous results
- Inference directly in function space

Definition: (Gaussian Processes)

GP is a collection of random variables, s.t. any finite number of which have joint Gaussian distribution

Function Space View

Notations:

$$f(\mathbf{x}) \sim GP(m(\mathbf{x}), k(\mathbf{x}, \tilde{\mathbf{x}})) \in \mathbb{R}, \mathbf{x} \in \mathbb{R}^D$$

$$m(\mathbf{x}) = \mathbb{E}[f(x)] \in \mathbb{R}, \text{ (mean function)}$$

$$k(\mathbf{x}, \tilde{\mathbf{x}}) = \mathbb{E}[(f(x) - m(\mathbf{x}))(f(\tilde{\mathbf{x}}) - m(\tilde{\mathbf{x}}))^T] \in \mathbb{R}$$

(covariance function)

GP is **completely specified** by its
mean function $m(\mathbf{x})$, and
covariance function $k(\mathbf{x}, \tilde{\mathbf{x}})$

Function Space View

Gaussian Processes:

For each $\mathbf{x} \in \mathbb{R}^D$ we associate a Gaussian variable $f(\mathbf{x})$ such that $\mathbb{R} \ni f(\mathbf{x}) \sim \mathcal{N}_{f(\mathbf{x})}(m(\mathbf{x}), k(\mathbf{x}, \mathbf{x}))$, and its correlation with other $f(\tilde{\mathbf{x}})$ variables is $k(\mathbf{x}, \tilde{\mathbf{x}})$.

$$\mathbb{R} \ni f(\mathbf{x}) \sim \mathcal{N}_{f(\mathbf{x})}(m(\mathbf{x}), k(\mathbf{x}, \mathbf{x}))$$

$$\begin{bmatrix} f(\mathbf{x}) \\ f(\tilde{\mathbf{x}}) \end{bmatrix} \sim \mathcal{N}_{\begin{bmatrix} f(\mathbf{x}) \\ f(\tilde{\mathbf{x}}) \end{bmatrix}} \left\{ \begin{bmatrix} m(\mathbf{x}) \\ m(\tilde{\mathbf{x}}) \end{bmatrix}, \begin{bmatrix} k(\mathbf{x}, \mathbf{x}) & k(\tilde{\mathbf{x}}, \mathbf{x}) \\ k(\mathbf{x}, \tilde{\mathbf{x}}) & k(\tilde{\mathbf{x}}, \tilde{\mathbf{x}}) \end{bmatrix} \right\}$$

Function Space View

The Bayesian linear regression is an example of GP

$$f(\mathbf{x}) = \phi(\mathbf{x})^T \mathbf{w} \in \mathbb{R}, \quad \phi(\mathbf{x}), \mathbf{w} \in \mathbb{R}^N \quad \mathbf{w} \sim \mathcal{N}_{\mathbf{w}}(0, \Sigma_p)$$

$\Rightarrow [f(\mathbf{x}_1), \dots, f(\mathbf{x}_k)]$ are jointly Gaussian $\forall \mathbf{x}_1, \dots, \mathbf{x}_k$
thus f is GP.

$$\mathbb{E}[f(\mathbf{x})] = \phi(\mathbf{x})^T \mathbb{E}[\mathbf{w}] = 0 \Rightarrow m(\mathbf{x}) = 0$$

$$\mathbb{E}[f(x)f(\tilde{\mathbf{x}})^T] = \phi(\mathbf{x})^T \underbrace{\mathbb{E}[\mathbf{w}\mathbf{w}^T]}_{\Sigma_p} \phi(\tilde{\mathbf{x}}) = k(\mathbf{x}, \tilde{\mathbf{x}})$$

Function Space View

Special case

$$\left. \begin{aligned} m(\mathbf{x}) &= 0 \\ k(\mathbf{x}, \tilde{\mathbf{x}}) &= \exp\left(-\frac{1}{2}\|\mathbf{x} - \tilde{\mathbf{x}}\|^2\right) \end{aligned} \right\} \Rightarrow f \text{ GP is given}$$

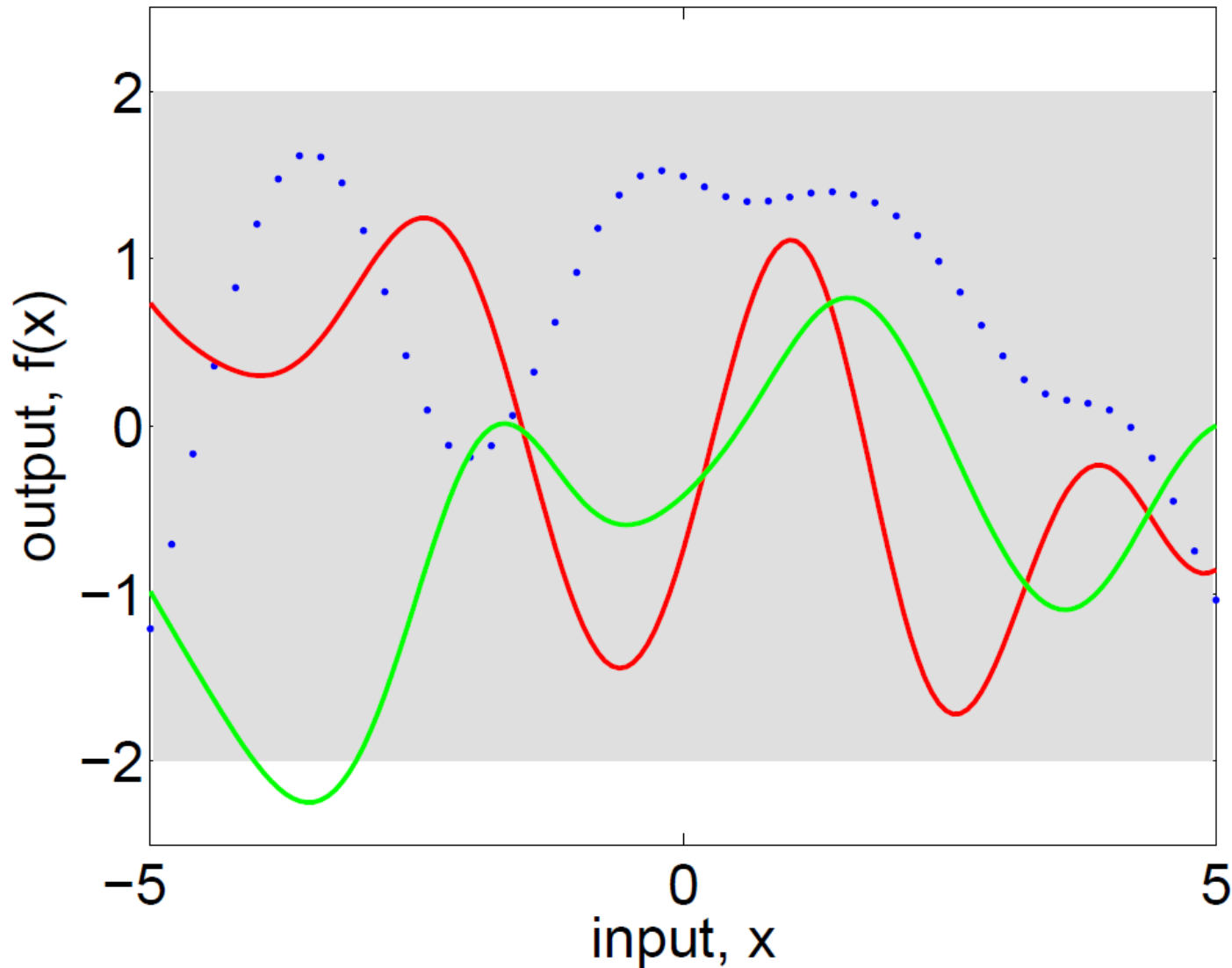
\Rightarrow implies a distribution over functions.

$$\text{Let } X_* = \begin{bmatrix} \mathbf{x}_{*1}^T \\ \vdots \\ \mathbf{x}_{*m}^T \end{bmatrix} \text{ } m \text{ input points}$$

$$\Rightarrow \mathbb{R}^m \ni f_* \sim \mathcal{N}_{f_*}(\underbrace{\mathbf{0}}_{\in \mathbb{R}^m}, \underbrace{k(X_*, X_*)}_{\in \mathbb{R}^{m \times m}})$$

At arbitray $\mathbf{x}_{*1}, \dots, \mathbf{x}_{*m}$ places, we can generate m points from f (denoted by f_*) and plot them .

Function Space View



Function Space View

Observation

The plotted $f(\mathbf{x}_{*1}), \dots, f(\mathbf{x}_{*m})$ function looks smooth.

Explanation

$$k(\mathbf{x}, \tilde{\mathbf{x}}) = \exp\left(-\frac{1}{2}\|\mathbf{x} - \tilde{\mathbf{x}}\|^2\right)$$

Thus if $\|\mathbf{x}_{*i} - \mathbf{x}_{*j}\|$ is small, then $\text{corr}(f(\mathbf{x}_{*i}), f(\mathbf{x}_{*j}))$ is high.

Prediction with noise free observations

Training set: $D = \{(\mathbf{x}_i, f_i) | i = 1, \dots, n\}$

$$X = \begin{bmatrix} \mathbf{x}_1^T \\ \vdots \\ \mathbf{x}_n^T \end{bmatrix} \in \mathbb{R}^{n \times D}, \text{ } m \text{ training inputs}$$

noise free observations



$$X_* = \begin{bmatrix} \mathbf{x}_{*1}^T \\ \vdots \\ \mathbf{x}_{*m}^T \end{bmatrix} \in \mathbb{R}^{m \times D}, \text{ } m \text{ test inputs}$$

$$f = \begin{bmatrix} f_1 \\ \vdots \\ f_n \end{bmatrix} \in \mathbb{R}^n, \text{ } n \text{ training targets}$$

$$f_* = \begin{bmatrix} f_{*1} \\ \vdots \\ f_{*m} \end{bmatrix} \in \mathbb{R}^m, \text{ } m \text{ test targets}$$

Prediction with noise free observations

$$\begin{bmatrix} f \\ f_* \end{bmatrix} \sim \mathcal{N} \begin{bmatrix} f \\ f_* \end{bmatrix} \left\{ \begin{bmatrix} \mathbf{0}_n \\ \mathbf{0}_m \end{bmatrix}, \underbrace{\begin{bmatrix} k(X, X) & k(X, X_*) \\ k(X_*, X) & k(X_*, X_*) \end{bmatrix}}_{\in \mathbb{R}^{(m+n) \times (m+n)}} \right\}$$

Goal:

We want to calculate the posterior distribution $f_* | X_*, X, f$

Methods:

- Generate f_* using the prior on \mathbf{w} , and reject those that are not compatible with the observations. (Computationally not efficient...)
- Calculate the posterior analytically.

Prediction with noise free observations

Lemma:

$$P(f_* | X_*, X, f) = N_{f_*} \left(k(X_*, X) k(X, X)^{-1} f, k(X_*, X_*) - k(X_*, X) k(X, X)^{-1} k(X, X_*) \right)$$

Proofs: a bit of calculation using the joint $(n+m)$ dim density

$$\begin{bmatrix} f \\ f_* \end{bmatrix} \sim \mathcal{N} \left[\begin{bmatrix} f \\ f_* \end{bmatrix}, \begin{bmatrix} \mathbf{0}_n & k(X, X_*) \\ \mathbf{0}_m & k(X_*, X_*) \end{bmatrix} \right]$$

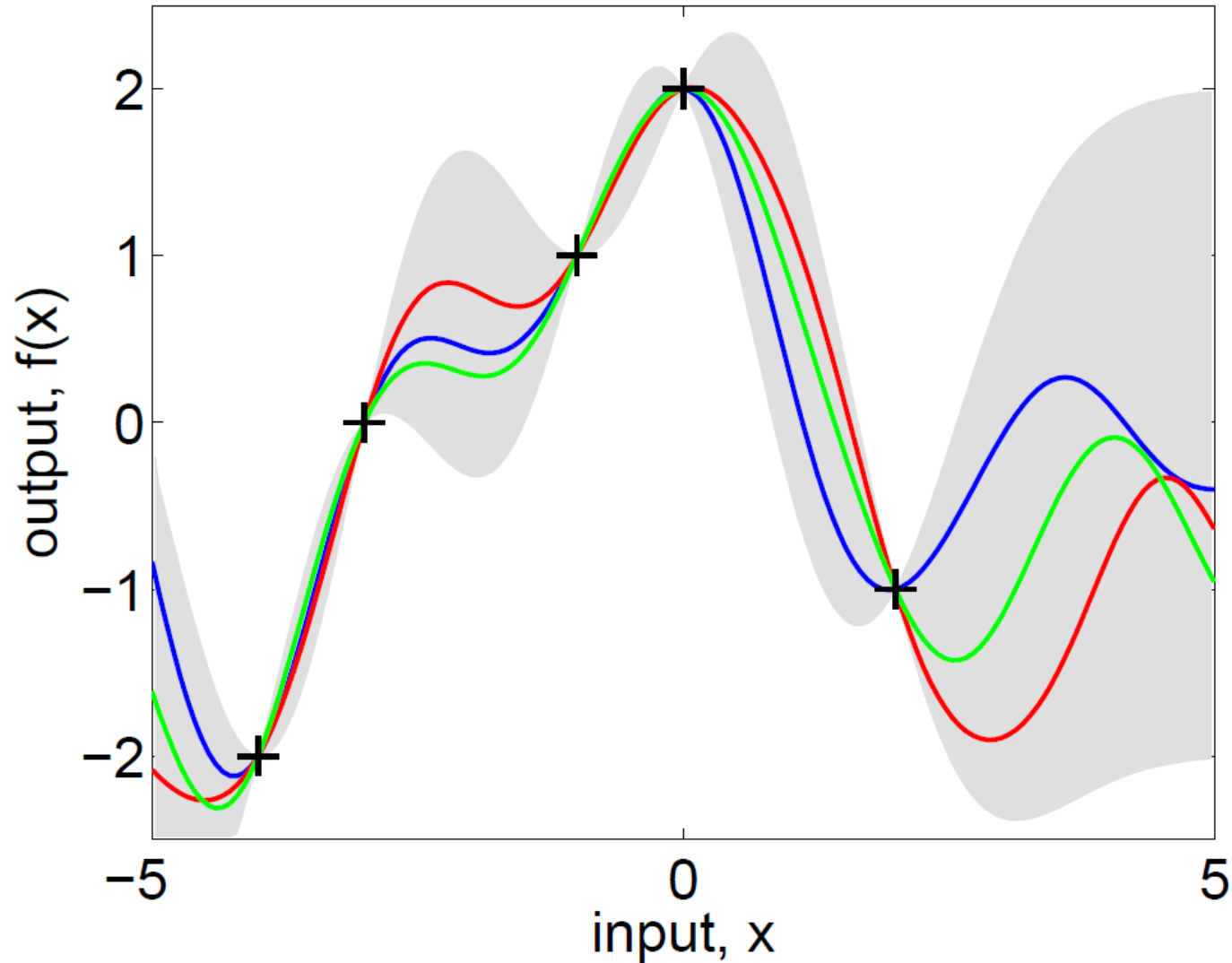
Remarks:

- If $X_* = X \Rightarrow f_* = f$ and the cov is 0. (noise free observations)
- $P(f_* | X_*, X, f)$ is isimilar to the previous results:

$$P(f_* | \mathbf{x}_*, X, \mathbf{y}) =$$

$$N_{f_*} \left((\phi_*^T \Sigma_p \phi) (K + \sigma^2 \mathbf{I}_n)^{-1} \mathbf{y}, (\phi_*^T \Sigma_p \phi_*) - (\phi_*^T \Sigma_p \phi) (K + \sigma^2 \mathbf{I}_n)^{-1} (\phi^T \Sigma_p \phi_*) \right)$$

Prediction with noise free observations



Picture is taken from Rasmussen and Williams

Prediction using noisy observations

$$y = f(\mathbf{x}) + \epsilon \in \mathbb{R}$$

$$\epsilon \sim \mathcal{N}(0, \sigma^2) \in \mathbb{R}$$

(Homoscedastic noise, the same for all \mathbf{x})

$$\begin{bmatrix} f \\ f_* \end{bmatrix} \sim \mathcal{N} \begin{bmatrix} f \\ f_* \end{bmatrix} \left\{ \begin{bmatrix} \mathbf{0}_n \\ \mathbf{0}_m \end{bmatrix}, \begin{bmatrix} k(X, X) & k(X, X_*) \\ k(X_*, X) & k(X_*, X_*) \end{bmatrix} \right\}$$

$$\Rightarrow \text{cov}(y_p, y_q) = k(\mathbf{x}_p, \mathbf{x}_q) + \sigma^2 \delta_{p,q}$$

$$\Rightarrow \text{cov}([y_1, \dots, y_n]) = k(X, X) + \sigma^2 \mathbf{I}_n \in \mathbb{R}^{n \times n}$$

The joint distribution:

$$\Rightarrow \begin{bmatrix} y \\ f_* \end{bmatrix} \sim \mathcal{N} \begin{bmatrix} y \\ f_* \end{bmatrix} \left\{ \begin{bmatrix} \mathbf{0}_n \\ \mathbf{0}_m \end{bmatrix}, \begin{bmatrix} k(X, X) + \sigma^2 \mathbf{I}_n & k(X, X_*) \\ k(X_*, X) & k(X_*, X_*) \end{bmatrix} \right\}$$

Prediction using noisy observations

The posterior for the noisy observations:

$$P(f_*|X, \mathbf{y}, X_*) = \mathcal{N}_{f_*}(\bar{f}_*, \text{cov}(f_*))$$

where

$$\bar{f}_* = \mathbb{E}[f_*|X, \mathbf{y}, X_*] = k(X_*, X)[k(X, X) + \sigma^2 I_n]^{-1} \mathbf{y} \in \mathbb{R}^m$$

$$\text{cov}(f_*) = k(X_*, X_*) - k(X_*, X)[k(X, X) + \sigma^2 I_n]^{-1} K(X, X_*) \in \mathbb{R}^{m \times m}$$

In the weight space view we had:

$$\bar{f}_* = (\phi_*^T \Sigma_p \phi)(\phi^T \Sigma_p \phi + \sigma^2 \mathbf{I}_n)^{-1} \mathbf{y}$$

$$\text{cov}(f_*) = (\phi_*^T \Sigma_p \phi_*) - (\phi_*^T \Sigma_p \phi)(\phi^T \Sigma_p \phi + \sigma^2 \mathbf{I}_n)^{-1} (\phi^T \Sigma_p \phi_*)$$

If $k(\mathbf{x}, \tilde{\mathbf{x}}) = \phi(x)^T \Sigma_p \phi(\tilde{x})$, then they are the same.

Prediction using noisy observations

Short notations:

$$K = k(X, X) \in \mathbb{R}^{n \times n}$$

$$K_* = k(X, X_*) \in \mathbb{R}^{n \times m}$$

$$k(\mathbf{x}_*) = k_* = k(X, \mathbf{x}_*) = \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_*) \\ \vdots \\ k(\mathbf{x}_n, \mathbf{x}_*) \end{bmatrix} \in \mathbb{R}^n$$

\Rightarrow for a single test point \mathbf{x}_* :

$$\bar{f}_* = \underbrace{k_*^T}_{\mathbb{R}^{1 \times n}} \underbrace{[K + \sigma^2 I_n]^{-1}}_{\mathbb{R}^{n \times n}} \underbrace{\mathbf{y}}_{\mathbb{R}^n} \in \mathbb{R}$$

$$\text{cov}(f_*) = \underbrace{k(\mathbf{x}_*, \mathbf{x}_*)}_{\mathbb{R}} - \underbrace{k_*^T}_{\mathbb{R}^{1 \times n}} \underbrace{[K + \sigma^2 I_n]^{-1}}_{\mathbb{R}^{n \times n}} \underbrace{k_*}_{\mathbb{R}^n} \in \mathbb{R}$$

Prediction using noisy observations

$$\bar{f}_* = \underbrace{k_*^T}_{\mathbb{R}^{1 \times n}} \underbrace{[K + \sigma^2 I_n]^{-1}}_{\mathbb{R}^{n \times n}} \underbrace{\mathbf{y}}_{\mathbb{R}^n} \in \mathbb{R}$$

Two ways to look at it:

- **Linear predictor**

$$\bar{f}_* = \beta^T \mathbf{y} = \beta_1 y_1 + \dots + \beta_n y_n$$

$$\text{where } \beta^T = k_*^T [K + \sigma^2 I_n]^{-1} \in \mathbb{R}^{1 \times n}$$

- **Manifestation of the Representer Theorem**

$$\bar{f}_* = \alpha^T k_* = \alpha_1 k(\mathbf{x}_1, \mathbf{x}_*) + \dots + \alpha_n k(\mathbf{x}_n, \mathbf{x}_*)$$

$$\text{where } \alpha = [K + \sigma^2 I_n]^{-1} \mathbf{y}$$

\bar{f}_* is a linear combination of n kernel values.

Prediction using noisy observations

$$\bar{f}_* = \underbrace{k_*^T}_{\mathbb{R}^{1 \times n}} \underbrace{[K + \sigma^2 I_n]^{-1}}_{\mathbb{R}^{n \times n}} \underbrace{\mathbf{y}}_{\mathbb{R}^n} \in \mathbb{R}$$

Remarks:

- While the GP in general is quite complex, for the prediction of $\bar{f}_* = f(\mathbf{x}_*)$ we need only the $(n+1)$ dimensional joint Gaussian distribution of $[y_1, \dots, y_n, f(\mathbf{x}_*)]$

- The posterior covariance of

$$\text{cov}(f_* | X, \mathbf{y}, X_*) = k(X_*, X_*) - k(X_*, X)[k(X, X) + \sigma^2 I_n]^{-1} K(X, X_*)$$

does not depend on the observed targets \mathbf{y} .

This is a peculiarity of GP.

GP pseudo code I

Inputs:

$$X = \begin{bmatrix} \mathbf{x}_1^T \\ \vdots \\ \mathbf{x}_n^T \end{bmatrix} \in \mathbb{R}^{n \times D}, \text{ } m \text{ training inputs}$$

$$\mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \in \mathbb{R}^n, \text{ } n \text{ training targets}$$

$k(\cdot, \cdot) : \mathbb{R}^{D \times D} \rightarrow \mathbb{R}$ covariance function (kernel)

\mathbf{x}_* test input

σ^2 noise level on the observations

$$[y(\mathbf{x}) = f(\mathbf{x}) + \epsilon, \epsilon \sim \mathcal{N}(0, \sigma^2)]$$

GP pseudo code II

1., $K \in \mathbb{R}^{n \times n}$ Gram matrix. $K_{ij} = k(\mathbf{x}_i, \mathbf{x}_j)$

$$k(\mathbf{x}_*) = k_* = k(X, \mathbf{x}_*) = \begin{bmatrix} k(\mathbf{x}_1, \mathbf{x}_*) \\ \vdots \\ k(\mathbf{x}_n, \mathbf{x}_*) \end{bmatrix} \in \mathbb{R}^n$$

$$2., \alpha = (K + \sigma^2 \mathbf{I}_n)^{-1} \mathbf{y}$$

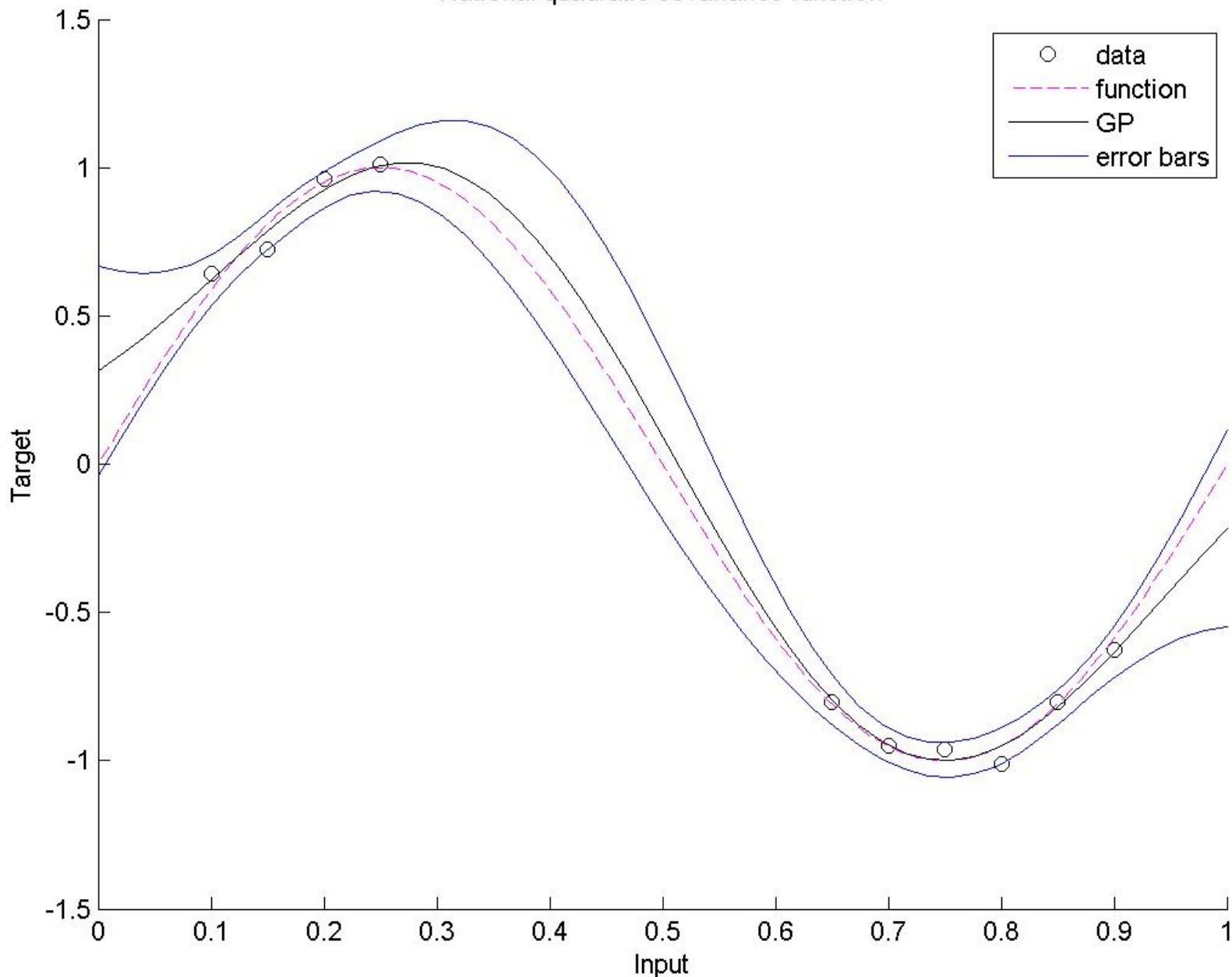
$$3., \bar{f}_* = k_*^T \alpha \in \mathbb{R}$$

$$4., \text{cov}(f_*) = \underbrace{k(\mathbf{x}_*, \mathbf{x}_*)}_{\mathbb{R}} - \underbrace{k_*^T}_{\mathbb{R}^{1 \times n}} \underbrace{[K + \sigma^2 I_n]^{-1}}_{\mathbb{R}^{n \times n}} \underbrace{k_*}_{\mathbb{R}^n} \in \mathbb{R}$$

Outputs: $\bar{f}_*, \text{cov}(f_*)$

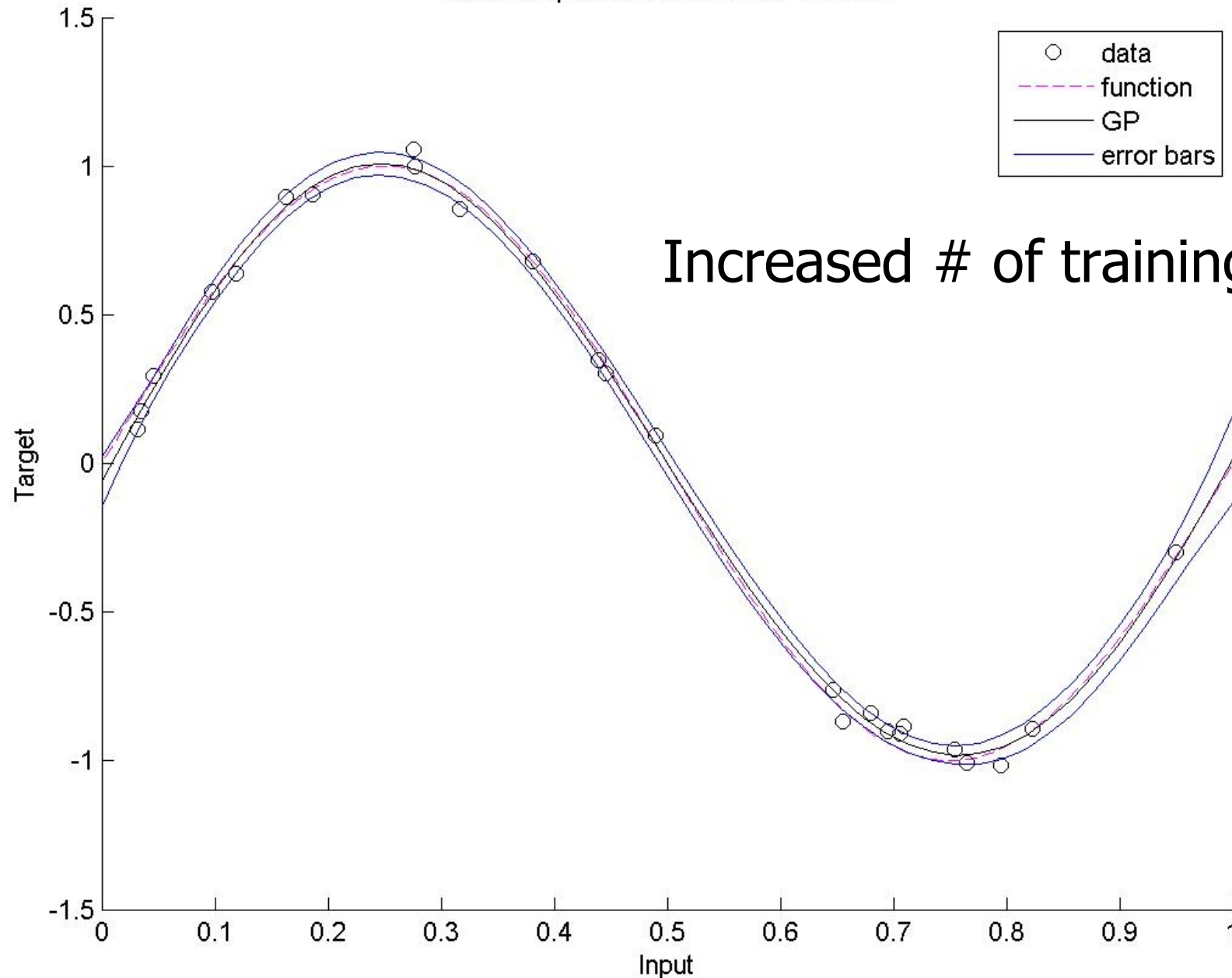
Results using Netlab , Sin function

Rational quadratic covariance function

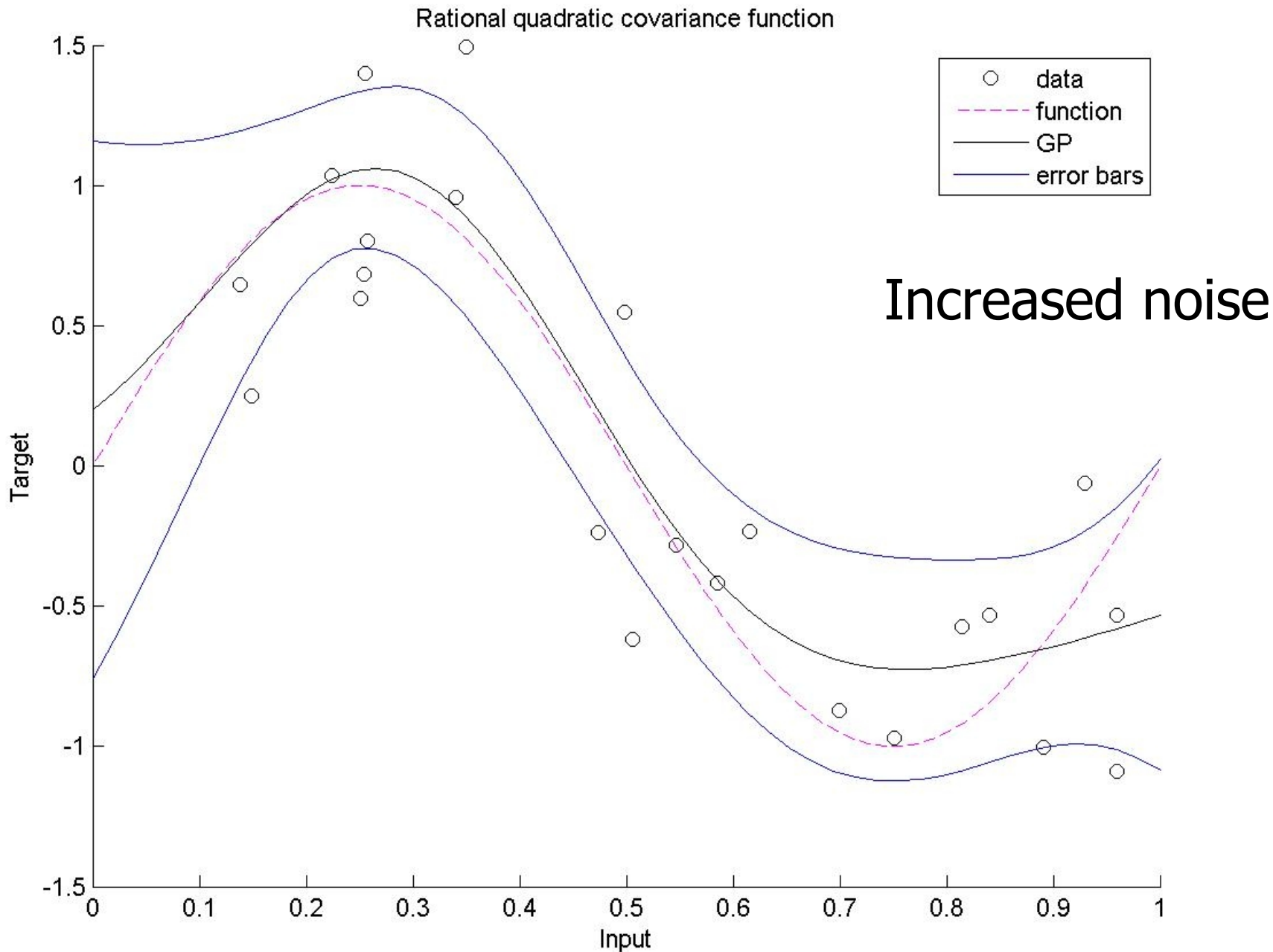


Results using Netlab, Sin function

Rational quadratic covariance function

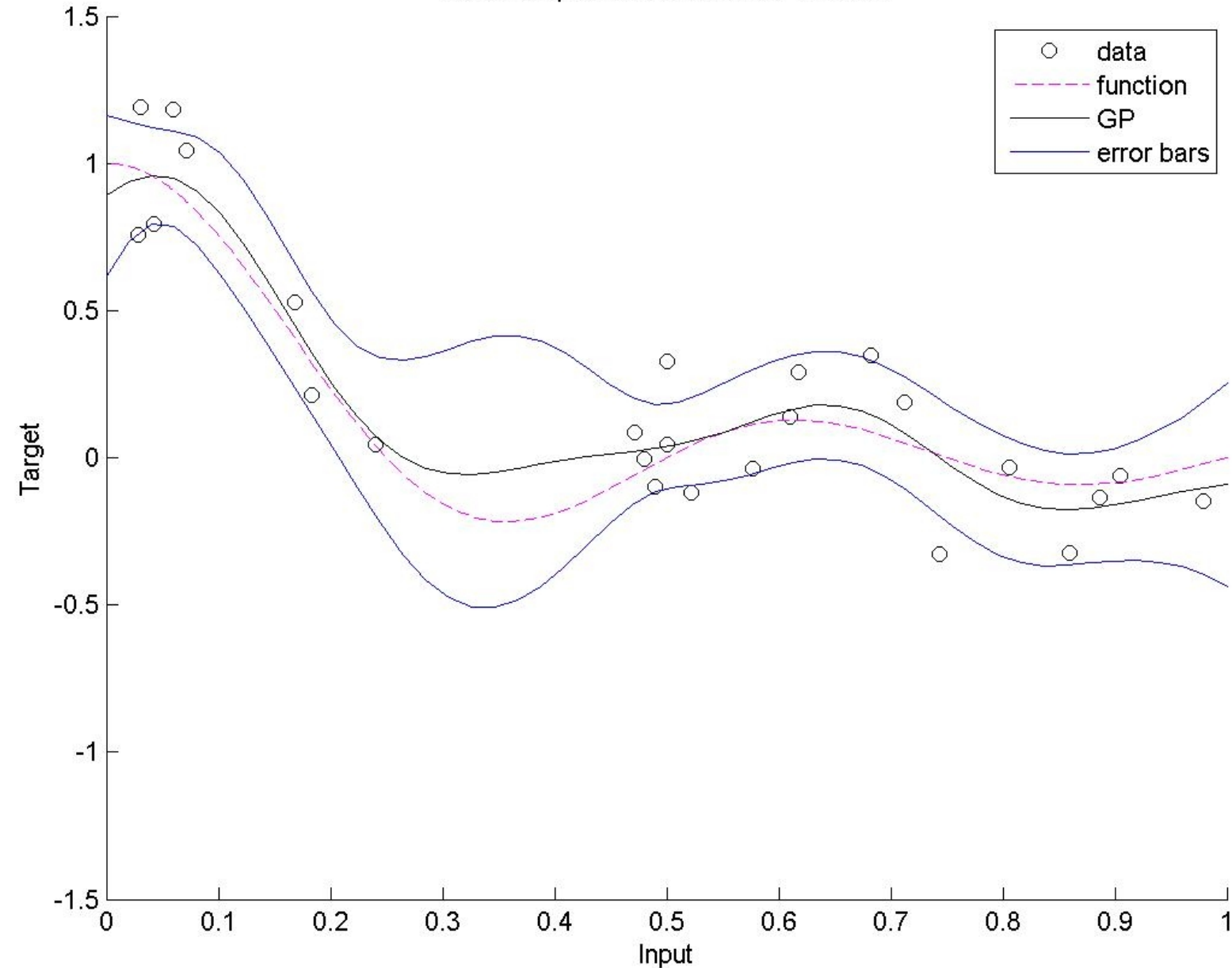


Results using Netlab, Sin function



Results using Netlab, Sinc function

Rational quadratic covariance function



Thanks for the Attention! 😊

