# Chapter 15 Estimators for Mean & Variance

### Probability vs. Statistics



<u>This chapter</u>: Given some data, infer the mean and variance of the underlying distribution.

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# Motivation for Estimation

Call this unknown  $\theta$ 

#### **Goal:** Estimate the mean height of women

Obviously can't measure heights of all women, so sample heights of nwomen, say n = 5.



#### Use **sample average** as estimation of $\theta$ .



### Sample Mean Estimator

<u>Defn 15.2</u>: Let  $X_1, X_2, ..., X_n$  be i.i.d. samples of r.v. X with unknown mean  $\theta = E[X]$ .

The **sample mean** is a point estimator of  $\theta$ :

 $\widehat{\theta}$ 

$$X_1 = x_1, X_2 = x_2, \dots, X_n = x_n) = M_n = \overline{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$
  
Alternate name  
for sample mean Alternate name  
for sample mean

### Choices of estimators

Let  $X_1, X_2, ..., X_n$  be i.i.d. samples of r.v. X with unknown finite mean  $\theta = E[X]$ and unknown finite variance  $\sigma^2 = Var(X)$ .

Two possible estimators for  $\theta$ :

$$\hat{\theta}_A = \bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$\hat{\theta}_B = X_2$$



<u>Defn 15.3</u>: Let  $\hat{\theta}(X_1, X_2, ..., X_n)$  be a point estimator for  $\theta$ . The **bias** of  $\hat{\theta}$  is:

$$\boldsymbol{B}(\hat{\theta}) = \boldsymbol{E}[\hat{\theta}] - \theta$$

unbiased estimator is desirable property #1

Both!

If  $B(\hat{\theta}) = 0$ , we say that  $\hat{\theta}$  is and **unbiased estimator** of  $\theta$ .

Which of  $\hat{\theta}_A = \bar{X}$  and  $\hat{\theta}_B = X_2$  are unbiased estimators?

So why do we prefer  $\widehat{ heta}_A$ ?





**Lemma 15.5:** If  $\hat{\theta}(X_1, X_2, ..., X_n)$  is an unbiased estimator, then

$$MSE(\hat{\theta}) = Var(\hat{\theta})$$

**Proof:** 
$$MSE(\hat{\theta}) = E\left[\left(\hat{\theta} - \theta\right)^2\right] = E\left[\left(\hat{\theta} - E[\hat{\theta}]\right)^2\right] = Var(\hat{\theta})$$

$$MSE(\hat{\theta}) = E\left[\left(\hat{\theta} - \theta\right)^{2}\right] = Var(\hat{\theta})$$
  
if  $\hat{\theta}(X_{1}, X_{2}, ..., X_{n})$  is  
unbiased

Q: Which of 
$$\hat{\theta}_A = \overline{X}$$
 and  $\hat{\theta}_B = X_2$  have lower MSE?

Both 
$$\hat{\theta}_A = \bar{X}$$
 and  $\hat{\theta}_B = X_2$  are unbiased estimators.  
 $\Rightarrow MSE(\hat{\theta}_A) = Var(\hat{\theta}_A) = \frac{1}{n^2} \cdot n Var(X) = \frac{Var(X)}{n}$ 
  
 $\Rightarrow MSE(\hat{\theta}_B) = Var(\hat{\theta}_B) = Var(X_2) = Var(X)$ 

We also want our estimator to become more accurate as the sample size increases.

<u>Defn 15.6</u>: Let  $\hat{\theta}_1(X_1)$ ,  $\hat{\theta}_2(X_1, X_2)$ ,  $\hat{\theta}_3(X_1, X_2, X_3)$ , ... be a sequence of point estimators of  $\theta$ , where  $\hat{\theta}_n(X_1, X_2, \dots, X_n)$  is a function of n i.i.d. samples.

We say that r.v.  $\hat{\theta}_n$  is a **consistent estimator** of  $\theta$  if,  $\forall \epsilon > 0$ ,  $\lim_{n \to \infty} \mathbf{P}\{|\hat{\theta}_n - \theta| \ge \epsilon\} = 0$  consistency is desirable property #3

Seems hard to prove that our estimator is consistent ....

**Lemma 15.7:** If  $\lim_{n\to\infty} MSE(\hat{\theta}_n) = 0$ , then  $\hat{\theta}_n$  is a consistent estimator.

Proof:  

$$P\{|\hat{\theta}_{n} - \theta| \ge \epsilon\} = P\{|\hat{\theta}_{n} - \theta|^{2} \ge \epsilon^{2}\} \le \frac{E[|\hat{\theta}_{n} - \theta|^{2}]}{\epsilon^{2}}$$

$$\le \frac{E[(\hat{\theta}_{n} - \theta)^{2}]}{\epsilon^{2}} = \frac{MSE(\hat{\theta}_{n})}{\epsilon^{2}}$$

$$\implies \lim_{n \to \infty} P\{|\hat{\theta}_{n} - \theta| \ge \epsilon\} = \lim_{n \to \infty} \frac{MSE(\hat{\theta}_{n})}{\epsilon^{2}} = 0$$

**Lemma 15.7:** If  $\lim_{n\to\infty} MSE(\hat{\theta}_n) = 0$ , then  $\hat{\theta}_n$  is a consistent estimator.

Q: Is 
$$\hat{\theta}_A = \bar{X} = M_n$$
 a consistent estimator of  $\boldsymbol{E}[X]$ ?

Yes!
$$MSE(\hat{\theta}_A) = Var(\hat{\theta}_A) = \frac{Var(X)}{n}$$
Assuming  
 $Var(X)$   
finite $\Longrightarrow$  $\lim_{n \to \infty} MSE(\hat{\theta}_A) = \lim_{n \to \infty} Var(\hat{\theta}_A) = \lim_{n \to \infty} \frac{Var(X)}{n} = 0$  $\widehat{\theta}_A$  is a consistent estimator

### An estimator for Variance

Call this unknown  $\theta$ 

#### Goal: Estimate the variance of height of women

Obviously can't measure heights of all women, so sample heights of nwomen, say n = 5.



Q: Can we use sample variance as an estimation of θ?Q: And how do we define sample variance ?

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### An estimator for Variance

#### Two cases:

- The case where know E[X] and want to estimate Var(X)
- 2. The case where *don't know E*[X] and want to estimate *Var*(X)



### Estimating Variance when Mean is Known

Assume  $\mu = \mathbf{E}[X]$  is known. Let  $\hat{\theta}$  be an estimator of  $\theta = Var(X)$ :  $\hat{\theta}(X_1, X_2, \dots, X_n) = \overline{S^2} \equiv \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$ mean squared distance from mean sample variance "Introduction to Probability for Computing", Harchol-Balter '24

### Estimating Variance when Mean is Known

Assume  $\mu = \mathbf{E}[X]$  is known. Let  $\hat{\theta}$  be an estimator of  $\theta = \mathbf{Var}(X)$ :

$$\hat{\theta}(X_1, X_2, \dots, X_n) = \overline{S^2} \equiv \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$$



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# Estimating Variance when Mean is UNknown

How about we Assume  $\mu = \mathbf{E}[X]$  is **not** known. replace  $\mu$  by  $\overline{X}$ ? Let  $\hat{\theta}$  be an estimator of  $\theta = Var(X)$ :  $\hat{\theta}(X_1, X_2, \dots, X_n) = \overline{S^2} \equiv \frac{1}{n} \sum_{i=1}^{n} (X_i - \mu)^2$  $\hat{\theta}(X_1, X_2, \dots, X_n) = \overline{S^2} \equiv \frac{1}{n} \sum_{i=1}^n (X_i - \overline{X})^2$ 

$$\overline{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

# Estimating Variance when Mean is UNknown

Assume  $\mu = \mathbf{E}[X]$  is **not** known. Let  $\hat{\theta}$  be an estimator of  $\theta = \mathbf{Var}(X)$ :

$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$\hat{\theta}(X_1, X_2, \dots, X_n) = \overline{S^2} \equiv \frac{1}{n} \sum_{i=1}^n (X_i - \overline{X})^2$$

**Q:** Is  $\overline{S^2}$  an unbiased estimator?

A: Not quite. In Exercise 15.4, you will prove that:

$$\boldsymbol{E}[\overline{S^2}] = \frac{n-1}{n} \cdot \boldsymbol{Var}(X)$$

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So need to multiply  $\overline{S^2}$  by  $\frac{n}{n-1}$  to get unbiased estimator

# Estimating Variance when Mean is UNknown

<u>Defn 15.8</u>: Let  $X_1, X_2, X_3, ...$  be i.i.d. samples of r.v X with **unknown** mean and variance. The **sample variance** is a point estimator of  $\theta = Var(X)$ . It is denoted by  $S^2$  and defined by:

$$\hat{\theta}(X_1, X_2, \dots, X_n) = \frac{n}{n-1}\overline{S^2} = \frac{1}{n-1}\sum_{i=1}^n (X_i - \overline{X})^2 \equiv S^2$$
sample variance

### More complex estimators

The sample mean estimator is often a component of more complex estimators. Consider the following example:

**Goal:** Estimate # German tanks in WW II.

Each tank has serial number. Below are the tanks we've seen so far ...



**Mathematically**, we're trying to estimate a maximum, call it  $\theta$ , based on seeing *n* samples,  $X_1, X_2, \dots, X_n$ , each randomly picked without replacement, from the integers 1, 2, 3, ...,  $\theta$ .



**Mathematically**, we're trying to estimate a maximum, call it  $\theta$ , based on seeing *n* samples,  $X_1, X_2, \dots, X_n$ , each randomly picked without replacement, from the integers 1, 2, 3, ...,  $\theta$ .

We know how to estimate the mean. Can we write the max in terms of the sample mean?



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$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

$$\boldsymbol{E}[\overline{X}] = \frac{1}{n} \left( \boldsymbol{E}[X_1] + \boldsymbol{E}[X_2] + \dots + \boldsymbol{E}[X_n] \right) = \frac{\theta + 1}{2}$$
$$\boldsymbol{E}[X_i] = 1 \cdot \frac{1}{\theta} + 2 \cdot \frac{1}{\theta} + 3 \cdot \frac{1}{\theta} + \dots + \theta \cdot \frac{1}{\theta} = \frac{\theta + 1}{2}$$
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**Mathematically**, we're trying to estimate a maximum, call it  $\theta$ , based on seeing *n* samples,  $X_1, X_2, ..., X_n$ , each randomly picked without replacement, from the integers 1, 2, 3, ...,  $\theta$ .

$$E[\bar{X}] = \frac{1}{n} (E[X_1] + E[X_2] + \dots + E[X_n]) = \frac{\theta + 1}{2}$$

Hence a reasonable estimator for  $\theta$  is:

$$\widehat{\theta}(X_1, X_2, \dots, X_n) = 2\overline{X} - 1$$

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$$\boldsymbol{E}[\bar{X}] = \frac{1}{n} (\boldsymbol{E}[X_1] + \boldsymbol{E}[X_2] + \dots + \boldsymbol{E}[X_n]) = \frac{\theta + 1}{2}$$
  
Is  $\hat{\theta}$  unbiased?

Yes!

= 0

 $\boldsymbol{E}[\hat{\theta}] = 2\boldsymbol{E}[\bar{X}] - 1$ 

Hence a reasonable estimator for  $\theta$  is:

$$\widehat{\theta}(X_1, X_2, \dots, X_n) = 2\overline{X} - 1$$

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$$E[\overline{X}] = \frac{1}{n} (E[X_1] + E[X_2] + \dots + E[X_n]) = \frac{\theta + 1}{2}$$
  
Is  $\hat{\theta}$  good?  
Not necessarily.

 $\hat{\theta} = 2\bar{X} - 1$ 

could be smaller

than some samples

Hence a reasonable estimator for  $\theta$  is:

$$\hat{\theta}(X_1, X_2, \dots, X_n) = 2\bar{X} - 1$$